Globally Sparse Probabilistic PCA - Supplementary Material

1. Detailed Inference for the relaxed model

We denote $\theta = (\mathbf{u}, \alpha, \sigma)$ the vector of parameters. In order to maximize the evidence $p(\mathbf{x_1}, ... \mathbf{x_n} | \theta)$, we adopt a variational approach. We view $\mathbf{y_1}, ... \mathbf{y_n}$ and \mathbf{W} as latent variables.

Given a (variational) distribution q over the space of latent variables, the variational free energy is given by

$$\mathcal{F}_q(\mathbf{x_1}, ... \mathbf{x_n} | \boldsymbol{\theta}) = -\mathbb{E}_q[\ln p(\mathbf{X}, \mathbf{Y}, \mathbf{W} | \boldsymbol{\theta})] - H(q)$$
(1)

where H denotes the differential entropy, and is an upper bound to the negative log-evidence:

$$-\ln p(\mathbf{X}|\boldsymbol{\theta}) = \mathcal{F}_q(\mathbf{X}|\boldsymbol{\theta}) - \mathrm{KL}(q||p(\cdot|\boldsymbol{\theta})) \le \mathcal{F}_q(\mathbf{X}|\boldsymbol{\theta}).$$

To be able to minimize $\mathcal{F}_q(\mathbf{X}|\boldsymbol{\theta})$, we make the following mean-field approximation on the variational distribution:

$$q(\mathbf{Y}, \mathbf{W}) = q(\mathbf{Y})q(\mathbf{W}) \tag{2}$$

We can now compute the variational posterior distribution q^* which minimizes the free energy. Note that two factorizations arise naturally. This will conveniently keep the size of the covariance matrices lower than d.

Proposition 1 The variational posterior distribution of the latent variables which minimizes the free energy is given by

$$q^*(\mathbf{Y}) = \prod_{i=1}^n \mathcal{N}(\mathbf{y}_i | \boldsymbol{\mu}_i, \boldsymbol{\Sigma})$$
(3)

and

$$q^*(\mathbf{W}) = \prod_{k=1}^p \mathcal{N}(\mathbf{w}_k | \mathbf{m}_k, \mathbf{S}_k)$$
 (4)

where

$$\boldsymbol{\mu}_i = \frac{1}{\sigma} \boldsymbol{\Sigma} \mathbf{M}^T \mathbf{U} \mathbf{x}_i, \ \boldsymbol{\Sigma} = \left(\mathbf{I}_d + \frac{1}{\sigma^2} \mathbf{M}^T \mathbf{U}^2 \mathbf{M} + \frac{1}{\sigma^2} \sum_{k=1}^p u_k^2 \mathbf{S}_k \right)^{-1}, \ \mathbf{M} = (\mathbf{m}_1, ... \mathbf{m}_n)$$

and

$$\mathbf{m}_k = \frac{u_k}{\sigma} \mathbf{S}_k \sum_{i=1}^n x_{i,k} \boldsymbol{\mu}_i, \ \mathbf{S}_k = \left(\frac{1}{\alpha} \mathbf{I}_d + \frac{n u_k^2}{\sigma^2} \boldsymbol{\Sigma} + \frac{u_k^2}{\sigma^2} \boldsymbol{\mathcal{M}}^T \boldsymbol{\mathcal{M}}\right)^{-1}, \ \boldsymbol{\mathcal{M}} = (\boldsymbol{\mu}_1, ... \boldsymbol{\mu}_p).$$

Proof Variational distribution of the latent vectors. Using a standard result in variational mean-field approximations (Bishop, 2006, chap. 10), we can write

$$\ln q^*(\mathbf{y}) = \mathbb{E}_{q(\mathbf{W})}[\ln p(\mathbf{X}, \mathbf{Y}, \mathbf{W}|\boldsymbol{\theta})]$$

which leads to the factorization $q^*(\mathbf{y}) = \prod_{i \leq n} q^*(\mathbf{y}_i)$. Then, for each $i \leq n$, we can write

$$\ln q^*(\mathbf{y}_i) = \mathbb{E}_{q(\mathbf{W})}[\ln p(\mathbf{X}, \mathbf{Y}, \mathbf{W}|\boldsymbol{\theta})] = \mathbb{E}_{q(\mathbf{W})}[\frac{-1}{2\sigma^2}||\mathbf{x}_i - \mathbf{U}\mathbf{W}\mathbf{y}_i||_2^2] - \frac{1}{2}||\mathbf{y}_i||_2^2$$

therefore

$$\ln q^*(\mathbf{y}_i) = \frac{-1}{2\sigma^2} \mathbf{y_i}^T \mathbb{E}_{q(\mathbf{W})} [\mathbf{W}^T \mathbf{U}^2 \mathbf{W}] \mathbf{y}_i + \frac{1}{\sigma^2} \mathbf{y}_i^T \mathbf{W}^T \mathbf{U} \mathbf{x}_i - \frac{1}{2} ||\mathbf{y}_i||_2^2$$

which lead to the desired form.

Variational distribution of the loading matrix. Similarly,

$$\ln q^*(\mathbf{W}) = \frac{-1}{2\sigma^2} \sum_{i=1}^n \mathbb{E}_{q(\mathbf{y}_i)}[||\mathbf{x}_i - \mathbf{U}\mathbf{W}\mathbf{y}_i||_2^2] - \frac{1}{2\alpha} \sum_{i=1}^n ||\mathbf{w}_i||_2^2$$

$$\ln q^*(\mathbf{W}) = \sum_{i=1}^n \left(\frac{-1}{2\sigma^2} \sum_{j=1}^p v_j^2 \mathbf{w}_j^T \mathbf{E}_{q(\mathbf{y}_i)}[||\mathbf{y}_i||_2^2] \mathbf{w}_j + \frac{1}{\sigma^2} \sum_{j=1}^p \mathbf{w}_j^T x_{i,j} v_j \mathbb{E}_{q(\mathbf{y}_i)}(\mathbf{y}_i) - \frac{1}{\alpha} ||\mathbf{w}_j||_2^2 \right)$$

which leads to the factorization $q^*(\mathbf{W}) = \prod_{j \leq p} q^*(\mathbf{w}_i)$ and to the desired expression.

We can now compute the value of the free energy.

Proposition 2 Up to unnecessary additive constants, the entropy of the variational distribution is given by

$$H(q) = \frac{n}{2} \ln |\mathbf{\Sigma}| + \frac{1}{2} \sum_{k=1}^{p} \ln |\mathbf{S}_k|.$$
 (5)

Proof We have, using the factorizations of the former proposition

$$H(q) = -\mathbb{E}_q[\ln q(\mathbf{Y}, \mathbf{W})] = -\sum_{i=1}^n \mathbb{E}_{q(\mathbf{y}_i)}[\ln q(\mathbf{y}_i)] - \sum_{j=1}^p \mathbb{E}_{q(\mathbf{W})}[\ln q(\mathbf{W})]$$

which allows us to conclude.

Proposition 3 Up to unnecessary additive constants, the negative free energy is given by

$$-\mathcal{F}_{q}(\mathbf{x}_{1},...\mathbf{x}_{n}|\boldsymbol{\theta}) = -np\ln\sigma - \frac{dp}{2}\ln\alpha - \frac{1}{2\sigma^{2}}\mathrm{Tr}(\mathbf{X}^{T}\mathbf{X}) - \frac{1}{2\sigma^{2}}\sum_{i=1}^{n}\sum_{k=1}^{p}u_{k}^{2}\mathrm{Tr}[(\boldsymbol{\Sigma}+\boldsymbol{\mu}_{i}\boldsymbol{\mu}_{i}^{T})(\boldsymbol{S}_{k}+\boldsymbol{m}_{i}\boldsymbol{m}_{i}^{T})]$$

$$+\frac{1}{\sigma^{2}}\sum_{i=1}^{n}\mathbf{x}_{i}^{T}\mathbf{U}\mathbf{M}\boldsymbol{\mu}_{i} + \sum_{k=1}^{p}-\frac{1}{2\alpha}\mathrm{Tr}(\mathbf{S}_{k}+\mathbf{m}_{k}\mathbf{m}_{k}^{T}) - \sum_{i=1}^{n}\mathrm{Tr}(\boldsymbol{\Sigma}+\boldsymbol{\mu}_{i}\boldsymbol{\mu}_{i}^{T})$$

$$+\frac{n}{2}\ln|\boldsymbol{\Sigma}| + \frac{1}{2}\sum_{k=1}^{p}\ln|\mathbf{S}_{k}|. \quad (6)$$

Proof By definition, we have

$$-\mathcal{F}_q(\mathbf{x_1},...\mathbf{x_n}|\boldsymbol{\theta}) = -\mathbb{E}_q[\ln p(\mathbf{X}, \mathbf{Y}, \mathbf{W}|\boldsymbol{\theta})] - H(q)$$

therefore

$$\mathcal{F}_{q}(\mathbf{x_{1}},...\mathbf{x_{n}}|\boldsymbol{\theta}) = -np\ln\sigma - \frac{1}{2\sigma^{2}}\mathrm{Tr}(\mathbf{X}^{T}\mathbf{X}) - \frac{1}{2\sigma^{2}}\sum_{i=1}^{n}\mathbb{E}_{q}[\mathbf{y}_{i}\mathbf{W}^{T}\mathbf{U}^{2}\mathbf{W}\mathbf{y}_{i}] + \frac{1}{\sigma^{2}}\sum_{i=1}^{n}\mathbf{x}_{i}^{T}\mathbf{U}\mathbf{M}\boldsymbol{\mu}_{i}$$

$$+ \sum_{k=1}^{p}\left(-\frac{d}{2}\ln\alpha - \frac{1}{2\alpha}\mathbb{E}_{q}[\mathbf{w}_{k}^{T}\mathbf{w}_{k}]\right) - \sum_{i=1}^{n}\mathbb{E}_{q}[\mathbf{y}_{i}^{t}\mathbf{y}_{i}]$$

$$+ \frac{n}{2}\ln|\boldsymbol{\Sigma}| + \frac{1}{2}\sum_{i=1}^{p}\ln|\mathbf{S}_{k}|.$$

and computing the expectations leads to the desired expression.

This expression allows us to find the following M-step updates:

$$\alpha^* = \frac{1}{dp} \sum_{k=1}^p \text{Tr}(\mathbf{S}_k + \mathbf{m}_k \mathbf{m}_k^T), \tag{7}$$

$$\sigma^* = \frac{\operatorname{Tr}(\mathbf{X}^T \mathbf{X})}{np} + \frac{1}{np} \sum_{i=1}^n \left(\sum_{k=1}^p u_k^2 \operatorname{Tr}[(\mathbf{\Sigma} + \boldsymbol{\mu}_i \boldsymbol{\mu}_i^T)(\boldsymbol{S}_k + \boldsymbol{m}_i \boldsymbol{m}_i^T)] - 2\mathbf{x}_i^T \mathbf{U} \mathbf{M} \boldsymbol{\mu}_i \right), \quad (8)$$

and, for $k \in \{1, ..., p\}$,

$$u_k^* = \operatorname{argmin}_{u \in [0,1]} \frac{u^2}{2\sigma^2} \sum_{i=1}^n \operatorname{Tr}[(\boldsymbol{\Sigma} + \boldsymbol{\mu}_i \boldsymbol{\mu}_i^T)(\boldsymbol{S}_k + \boldsymbol{m}_i \boldsymbol{m}_i^T)] - u \sum_{i=1}^n x_{i,k} \boldsymbol{m}_k^T \boldsymbol{\mu}_i$$
(9)

Note that the objective function of the optimization problem (9) is simply a univariate polynomial.

References

C. M. Bishop. Pattern recognition and machine learning. Springer, 2006.