Constant Step Size Stochastic Gradient Descent for Probabilistic Modeling

Supplementary material

In this supplementary material we provide explicit expressions for the asymptotic expansions from the main paper. All assumptions from [2] are reused, namely, beyond the usual sampling assumptions, smoothness of the cost functions.

1 Explicit form of $B, \bar{B}, \bar{\bar{B}}^w$ and $\bar{\bar{B}}^m$

We have, even for mis-specified models:

$$\mathcal{F}(\eta) - \mathcal{F}(\eta_{*}) = \mathcal{G}(\mu) - \mathcal{G}(\mu_{*}) =$$

$$= \mathbb{E}\left[-\mathbb{E}_{p(y_{n}|x_{n})}y_{n}\eta(x_{n}) + a(\eta(x_{n})) + \mathbb{E}_{p(y_{n}|x_{n})}y_{n}\eta_{*}(x_{n}) - a(\eta_{*}(x_{n}))\right] =$$

$$= \mathbb{E}\left[a(\eta(x_{n})) - a(\eta_{*}(x_{n})) - a'(\eta_{*}(x_{n}))(\eta(x_{n}) - \eta_{*}(x_{n}))\right] + \mathbb{E}\left[\left(a'(\eta_{*}(x_{n})) - \mathbb{E}_{p(y_{n}|x_{n})}y_{n}\right) \cdot \left(\eta(x_{n}) - \eta_{*}(x_{n})\right)\right] =$$

$$\mathbb{E}\left[D_{a}(\eta(x_{n})|\eta_{*}(x_{n}))\right] + \mathbb{E}\left[\left(\mu_{*}(x) - \mu_{**}(x)\right) \cdot \left(\eta(x_{n}) - \eta_{*}(x_{n})\right)\right] =$$

$$\mathbb{E}\left[D_{a^{*}}\left(\mu(x_{n})|\mu_{*}(x_{n})\right)\right] + \mathbb{E}\left[\left(\mu_{*}(x) - \mu_{**}(x)\right) \cdot \eta(x_{n})\right]$$

$$(1)$$

for D_a the Bregman divergence associated to a, and D_{a^*} the one associated to a^* . We also use the optimality condition for the predictor $\eta_*(x)$: $\mathbb{E}\eta_*(x)[a'(\eta_*(x)) - \mathbb{E}_{p(x|y)}y] = 0$ in the last step.

When the model is well-specified, we have $a'(\eta_*(x_n)) = \mathbb{E}(y_n|x_n)$ and thus $F(\eta) - F(\eta_*) = \mathbb{E}[D_{a^*}(\mu_*(x_n)||\mu(x_n))]$. If η is linear in $\Phi(x)$, and even if the model is mis-specified, then we also have $F(\eta) - F(\eta_*) = \mathbb{E}[D_{a^*}(\mu_*(x_n)||\mu(x_n))]$.

Using asymptotic expansions of moments of the averaged SGD iterate with zero-mean statistically independent noise $f_n(\theta) = F(\theta) + \varepsilon_n(\theta)$ from [1], Theorem 2 one obtains:

$$\overline{\theta}_{\gamma} = \mathbb{E}_{\pi_{\gamma}}(\theta) = \theta_* + \gamma \Delta + O(\gamma^2), \tag{2}$$

$$\mathbb{E}_{\pi_{\gamma}}(\theta - \theta_*)(\theta - \theta_*)^{\top} = \gamma C + O(\gamma^2), \tag{3}$$

where

$$C = \left[F''(\theta_*) \otimes I + I \otimes F''(\theta_*) \right]^{-1} \Sigma.$$

and
$$\Sigma = \int_{\mathbb{R}^d} \varepsilon(\theta)^{\otimes 2} \pi_{\gamma}(d\theta) \in \mathbb{R}^{d \times d}$$
.

The "drift" $\bar{\theta}_{\gamma} - \theta_*$ is linear in γ and can be interpreted as an additional error due to the function is not being quadratic and step sizes are not decaying to zero.

Connection between Δ and C can be easily obtained using $\theta_n = \theta_{n-1} - \gamma [F'(\theta_{n-1}) + \varepsilon_n]$. Taking expectation of both parts and using Tailor expansion up to the second order:

$$F''(\theta_*)(\overline{\theta}_{\gamma} - \theta_*) = -\frac{1}{2}F'''(\theta_*)\mathbb{E}_{\pi_{\gamma}}(\theta - \theta_*)^{\otimes 2} \Rightarrow$$

$$F''(\theta_*)\Delta = -\frac{1}{2}F'''(\theta_*)C. \tag{4}$$

1.1 Estimation without averaging

We start with the simplest estimator of the prediction function: $\mu_0(x) = a'(\Phi^{\top}\theta_n)$, where we do not use any averaging:

$$\mathcal{G}(\mu_n) - \mathcal{G}(\mu_*) = f(\theta_n) - f(\theta_*) = f'(\theta_*)(\theta_n - \theta_*) + \frac{1}{2}f''(\theta_*)(\theta_n - \theta_*)^{\otimes 2} + \frac{1}{6}f'''(\theta_*)(\theta_n - \theta_*)^{\otimes 3} + O(\gamma^{3/2})$$

Taking expectation of both sides, when $n \to \infty$ and using Eq. (3) one obtains:

$$A(\gamma) = \mathbb{E}_{\pi_{\gamma}} f(\theta_n) - f(\theta_*) = \frac{1}{2} \operatorname{tr} f''(\theta_*) \gamma C + O(\gamma^{3/2}).$$

So, we have linear dependence of γ and $B = \frac{1}{2} \text{tr} f''(\theta_*) C$.

1.2 Estimation with averaging parameters

Now, let us estimate $\overline{A}(\gamma)$:

$$\Im(\bar{\mu}_n) - \Im(\mu_*) = f(\bar{\theta}_n) - f(\theta_*) = f'(\theta_*)(\bar{\theta}_n - \theta_*) + \frac{1}{2}(\bar{\theta}_n - \theta_*)f''(\theta_*)(\bar{\theta}_n - \theta_*) + O(\gamma^3).$$

Taking expectation of both sides, when $n \to \infty$:

$$\Im(\bar{\mu}_{\gamma}) - \Im(\mu_{*}) = f(\bar{\theta}_{\gamma}) - f(\theta_{*}) = \frac{1}{2} \operatorname{tr} f''(\theta_{*}) (\bar{\theta}_{\gamma} - \theta_{*})^{\otimes 2} + O(\gamma^{3}) = \frac{1}{2} \operatorname{tr} f''(\theta_{*}) \gamma^{2} \Delta^{\otimes 2} + O(\gamma^{3}).$$

Finally we have a quadratic dependence of γ :

$$\bar{A}(\gamma) = \frac{1}{2} \operatorname{tr} f''(\theta_*) \gamma^2 \Delta^{\otimes 2} + O(\gamma^3).$$

And the coefficient $\bar{B} = \frac{1}{2} \text{tr} f''(\theta_*) \Delta^{\otimes 2}$.

1.3 Estimation with averaging predictions

Recall, that by definition, $\bar{A}(\gamma) = \mathcal{G}(\bar{\mu}_{\gamma}) - \mathcal{G}(\mu_{*})$, where $\bar{\mu}_{\gamma}(x) = \mathbb{E}_{\pi_{\gamma}} a'(\theta^{\top} \Phi(x))$. We again use Tailor expansion for $a'(\theta^{\top} \Phi(x))$ at θ^{*} :

$$a'\left(\theta^{\top}\Phi(x)\right) = a'\left(\theta_{*}^{\top}\Phi(x)\right) + a''\left(\theta_{*}^{\top}\Phi(x)\right)\left(\theta - \theta_{*}\right)^{\top}\Phi(x) + \frac{1}{2}a'''\left(\theta_{*}^{\top}\Phi(x)\right) \cdot \left(\left(\theta - \theta_{*}\right)^{\top}\Phi(x)\right)^{2} + O(\gamma^{3/2}).$$

Taking expectation of both parts:

$$\bar{\mu}_{\gamma}(x) = \mu_{*}(x) + a'' \left(\theta_{*}^{\top} \Phi(x)\right) (\bar{\theta}_{\gamma} - \theta_{*})^{\top} \Phi(x) + \frac{1}{2} a''' \left(\theta_{*}^{\top} \Phi(x)\right) \operatorname{tr} \left[\Phi(x) \Phi(x)^{\top} \mathbb{E}(\theta - \theta_{*})^{\otimes 2}\right] + O(\gamma^{3/2}) =$$

$$= \mu_{*}(x) + a'' \left(\theta_{*}^{\top} \Phi(x)\right) \gamma \Delta^{\top} \Phi(x) + \frac{1}{2} a''' \left(\theta_{*}^{\top} \Phi(x)\right) \operatorname{tr} \left[\Phi(x) \Phi(x)^{\top} \gamma C\right] + O(\gamma^{3/2}).$$

Finally, we showed, that:

$$\bar{\bar{\mu}}_{\gamma}(x) - \mu_{*}(x) = O(\gamma^{3/2}) + \gamma \left[a'' \left(\eta_{*}(x) \right) \Delta^{\top} \Phi(x) + \frac{1}{2} a''' \left(\eta_{*}(x) \right) \operatorname{tr} \left[\Phi(x)^{\otimes 2} C \right] \right]$$

Now we use Bregram divergence notation Eq. (1):

$$\bar{\bar{A}}(\gamma) = \mathfrak{G}(\bar{\bar{\mu}}_{\gamma}) - \mathfrak{G}(\mu_*) = \mathfrak{G}_1 + \mathfrak{G}_2,$$

As mentioned above, the term \mathcal{G}_2 vanishes if model is well-specified or η is linear in $\Phi(x)$. Note, that for the case $\bar{A}(\gamma)$ indeed linear in $\Phi(x)$.

1.3.1 Estimation of \mathfrak{G}_1 .

By definition $D_{a^*}(\mu_*(x)||\mu(x)) = \frac{1}{2}(\mu_*(x) - \mu(x))(a^*)''(\mu_*(x))(\mu_*(x) - \mu(x))$ and

$$\mathcal{G}_1 = \frac{1}{2} \mathbb{E} \frac{\left(\mu_*(x) - \bar{\mu}_{\gamma}(x)\right)^2}{a''(\theta_*^\top \Phi(x))} = \frac{\gamma^2}{2} \mathbb{E} \left[\frac{\left(a''(\eta_*(x))\Delta^\top \Phi(x) + \frac{1}{2}a'''(\eta_*(x))\operatorname{tr}\left[\Phi(x)^{\otimes 2}C\right]\right)^2}{a''(\theta_*^\top \Phi(x))} \right].$$

Since

$$\mathbb{E}_x \Big[a^{\prime\prime} (\left(\theta_*^\top \Phi(x) \right) (\Delta^\top \Phi(x))^2 \Big] = \Delta^\top f^{\prime\prime} (\theta_*) \Delta$$

and

$$\mathbb{E}_x \left[\Delta^\top a''' \left(\theta_*^\top \Phi(x) \right) \Phi(x)^{\otimes 3} C \right] = \Delta^\top f'''(\theta_*) C = -2\Delta^\top f''(\theta_*) \Delta,$$

$$\mathcal{G}_1 = \gamma^2 \left[-\frac{1}{2} \Delta^\top f''(\theta_*) \Delta + \frac{1}{8} \mathbb{E} \left[\frac{a'''(\eta_*(x))^2}{a''(\eta_*(x))} \cdot \left(\text{tr} \left[\Phi(x)^{\otimes 2} C \right] \right)^2 \right] \right] + O(\gamma^3)$$

And the coefficient $\bar{\bar{B}}^w = -\frac{1}{2}\Delta^{\top}f''(\theta_*)\Delta + \frac{1}{8}\mathbb{E}\left[\frac{a'''(\eta_*(x))^2}{a''(\eta_*(x))}\cdot\left(\operatorname{tr}\left[\Phi(x)^{\otimes 2}C\right]\right)^2\right].$

1.3.2 Estimation of \mathfrak{G}_2 .

$$\mathfrak{G}_2 = \mathbb{E}\Big[\left(\mu_*(x) - \mu_{**}(x) \right) \cdot \left(\bar{\bar{\eta}}(x_n) - \eta_*(x_n) \right) \Big],$$

using properties of conjugated functions,

$$\mathcal{G}_{2} = \mathbb{E}\Big[\big((a^{*})'(\bar{\mu}(x)) - (a^{*})'(\mu_{*}(x)) \big) \cdot \big(\mu_{*}(x) - \mu_{**}(x) \big) \Big] = \\
\mathbb{E}\Big[(a^{*})''(\bar{\mu}_{*}(x))(\bar{\mu}(x) - \mu_{*}(x)) \cdot \big(\mu_{*}(x) - \mu_{**}(x) \big) + O(\gamma^{2}) = \\
= \mathbb{E}\frac{\bar{\mu}(x) - \mu_{*}(x)}{a''(\eta_{*}(x))} \cdot \big(\mu_{*}(x) - \mu_{**}(x) \big) + O(\gamma^{2}) = \\
= \gamma \cdot \mathbb{E}\Big[\Big(\Delta^{\top} \Phi(x) + \frac{a'''(\eta_{*}(x))}{2a''(\eta_{*}(x))} \text{tr} \big[\Phi(x)^{\otimes 2} C \big] \Big) \cdot \Big(\mu_{*}(x) - \mu_{**}(x) \Big) \Big] + O(\gamma^{2}).$$

And the coefficient $\bar{\bar{B}}^m = \mathbb{E}\Big(\Delta^{\top}\Phi(x) + \frac{a'''(\eta_*(x))}{2a''(\eta_*(x))} \mathrm{tr}\big[\Phi(x)^{\otimes 2}C\big]\Big) \cdot \Big(\mu_*(x) - \mu_{**}(x)\Big).$

References

- [1] A. Dieuleveut and F. Bach. Nonparametric stochastic approximation with large step-sizes. *Ann. Statist.*, 44(4):1363–1399, 08 2016.
- [2] A. Dieuleveut, A. Durmus, and F. Bach. Bridging the gap between constant step size stochastic gradient descent and markov chains. Technical Report 1707.06386, arXiv, 2017.