

Instance-Dependent Generalization Bounds via Optimal Transport

Songyan Hou*

SONGYAN.HOU@ETHZ.CH

Department of Mathematics, ETH Zurich

Parnian Kassraie*†

PKASSRAIE@ETHZ.CH

Department of Computer Science, ETH Zurich

Anastasis Kratsios*

KRATSIOA@MCMASTER.CA

Department of Mathematics, McMaster University and the Vector Institute

Andreas Krause

KRAUSEA@ETHZ.CH

Department of Computer Science, ETH Zurich

Jonas Rothfuss*

JONAS.ROTHFUSS@INF.ETHZ.CH

Department of Computer Science, ETH Zurich

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Abstract

Existing generalization bounds fail to explain crucial factors that drive the generalization of modern neural networks. Since such bounds often hold uniformly over all parameters, they suffer from over-parametrization and fail to account for the strong inductive bias of initialization and stochastic gradient descent. As an alternative, we propose a novel *optimal transport* interpretation of the generalization problem. This allows us to derive *instance-dependent* generalization bounds that depend on the *local Lipschitz regularity* of the *learned prediction function* in the data space. Therefore, our bounds are agnostic to the parametrization of the model and work well when the number of training samples is much smaller than the number of parameters. With small modifications, our approach yields accelerated rates for data on *low-dimensional manifolds* and guarantees under *distribution shifts*. We empirically analyze our generalization bounds for neural networks, showing that the bound values are meaningful and capture the effect of popular regularization methods during training.

Keywords: Generalization Bound, Instance-Dependent, Optimal Transport, Local Lipschitz Regularity, Distributional Robustness

1. Introduction

A core challenge in machine learning is to generalize well beyond the training data. We want to choose a hypothesis $f \in \mathcal{F}$, from a hypothesis class \mathcal{F} , that not only gives a small training error but also yields good predictions for previously unseen data points. Accordingly, statistical learning theory aims to provide generalization guarantees and understand the factors that drive it. Generalization is typically described through the discrepancy between two key quantities: The empirical risk $\hat{\mathfrak{R}}(f)$, i.e., the prediction error of f on the training data and the expected risk $\mathfrak{R}(f)$, i.e., the expected error under the unknown data-distribution.

*. Equal contribution, all authors are listed in alphabetic order.

†. Corresponding author.

A common type of guarantees are *uniform bounds* which control the generalization gap $\mathfrak{R}(f) - \hat{\mathfrak{R}}(f)$ with high probability, simultaneously for *all hypotheses* $f \in \mathcal{F}$ (e.g., [Vapnik and Chervonenkis, 1971](#); [Bartlett and Mendelson, 2002](#)). Such bounds include terms that quantify the complexity of the hypothesis f or hypothesis space \mathcal{F} . For neural networks (NNs), this complexity term grows rapidly with the number of parameters (e.g., [Bartlett et al., 2017](#); [Neyshabur et al., 2015](#); [Harvey et al., 2017](#)). While the parameter space of NNs is vast, regular networks that are used in practice only seem to populate a small subset of the parameter space. This subset seemingly generalizes well and depends on model structure, initialization scheme, and optimization method in a complex manner. In addition, there are many NN parameter configurations that correspond to the same neural network mapping, artificially inflating the complexity of the parametric hypothesis space. Thus, such uniform bounds in the parameter space fail to explain the empirical generalization behavior of neural networks in the over-parameterized setting where the number of training examples is much smaller than the number of parameters ([Belkin et al., 2019](#)).

Addressing this issue, we base our analysis on the geometric properties of the learned prediction function (i.e., hypothesis f) in the data domain. In particular, we partition the input domain into smaller neighborhoods and locally characterize f via its *local Lipschitz constant* when the domain is restricted to each neighborhood. Using principles from *optimal transport*, we obtain a bound that depends on the instance f through its local Lipschitz constants. In particular, we view the generalization gap as the worst-case impact on the loss when probability mass is transported from the empirical measure to the true data distribution. The magnitude of this impact depends on the local regularity of f multiplied by the local transport cost, which decreases w.h.p. with the number of samples.

Classical uniform bounds depend on the complexity of the hypothesis space or the global regularity of f which is typically determined by the single most irregular part of f in the domain. However, neural network functions often vary significantly in their (local) regularity across the domain. This typically leads to extremely loose or vacuous bounds. In contrast, our bounds can adapt to the local regularity properties of f and, thus, minimize the negative impact of irregular parts of the learned function on the tightness of the bound.

Overall, the presented generalization bound (Theorem 4) has the following properties: 1) It is instance-dependent, i.e., it adapts to the trained function f and thus can capture the combined effect of initialization, training method, and model structure. 2) It characterizes f geometrically via its local Lipschitz regularity; therefore, in contrast to parametric bounds, it does not suffer from over-parametrization. 3) It is tighter than bounds based on the global Lipschitz properties of f due to the fine-grained local analysis, which takes into account changes in the regularity of f throughout the domain. While our bounds generically hold for any machine learning model, we focus our exposition on neural network generalization and empirically verify the mentioned properties through experiments. When applied to fully-connected ReLU networks, trained on simple regression and classification tasks, we observe that our result provides meaningful bound values in the same order of magnitude as the empirical risk, even for small sample sizes. We empirically show that, unlike the majority of prior works, the bound does not explode as the number of network parameters increases. Moreover, the value of the bound reflects the effect of regularization techniques applied *during* training, e.g., weight-decay, early-stopping, and adversarial training.

Due to its transport-based derivation, our framework can be seamlessly adapted to obtain generalization certificates under distribution shifts or adversarial perturbations. The results mentioned above are corollaries of our core theorem, which is an optimal-transport-based concentration inequality for data-dependent locally regular functions. This theorem may be of independent interest and considers a spectrum of functions with different degrees of regularity, from non-smooth α -Hölder functions to smooth and s -time differentiable instances.

Outline The paper is structured as follows.

- Section 3 formalizes the problem setting and presents our main generalization bound (Theorem 4) together with an extension to when the data is known to be concentrated on a low-dimensional manifold (Proposition 5).
- Section 4 discusses the key properties of our generalization bound which is data-dependent (Section 4.1), non-parametric (Section 4.2), and localized (Section 4.3). Every section also presents corresponding experiments on neural networks.
- Section 5 considers generalization under distribution shifts, which is a natural corollary of our approach (Corollary 8).
- Section 6 focuses on our core result (Theorem 10). Section 6.1 highlights the key technical tools used for this theorem, and Section 6.2 outlines the proof methodology.

2. Related Work

Our work provides generalization bounds for learned prediction functions, contributing to the rich literature on generalization. A classic approach to explaining generalization are uniform bounds, which provide *uniform* guarantees over a class of estimators, also referred to as the hypothesis space. Uniform bounds often depend on the combinatorial complexity of the hypothesis space, e.g., expressed in the form of the VC-dimension (Vapnik and Chervonenkis, 1971) or the Rademacher complexity (Koltchinskii, 2001; Bartlett and Mendelson, 2002). For neural networks, however, the hypothesis space is large and combinatorially explodes in size with the neural network width and depth, making the corresponding bounds loose (cf. Bartlett et al., 1998; Harvey et al., 2017; Bartlett et al., 2019; Sun et al., 2016). Uniform bounds that utilize the parametric characterization of the network rapidly with the size of the neural network (e.g., Neyshabur et al., 2015). Overall, these approaches hardly explain the empirical generalization behavior of neural networks in the over-parameterized setting, where the number of samples is much smaller than the number of parameters (Belkin et al., 2019). In fact, measures of neural network complexity based on the VC-dimension or parameter norm were found to be negatively correlated with the expected risk of convolutional neural networks (Jiang et al., 2019b; Kuhn et al., 2021). In contrast, we present results that use the geometric properties, i.e., local regularity, of the learned prediction function f . This allows us to avoid the dependence on the combinatorial complexity of function classes as well as direct dependency on the parametrization of f .

An alternative to guarantees that hold uniformly over a hypothesis space are instance-dependent bounds, where the value of the bound changes based on properties of the learned hypothesis, which generally depends on the training data. In this spirit, PAC-Bayesian

learning theory provides generalization bounds which depend on the chosen posterior distribution, e.g., an instance of the random (Gibbs) learner (McAllester, 1998; Shawe-Taylor and Williamson, 1997; Catoni, 2007; Alquier et al., 2016; Mhammedi et al., 2019). Since PAC-Bayesian bounds do not trivially explode with the number of parameters of the model, they have gained increasing popularity in the context of neural networks (Langford and Caruana, 2001; Dziugaite and Roy, 2017, 2018; Neyshabur et al., 2018; Zhou et al., 2018; Golowich et al., 2020). For instance, they have been related to the sharpness of minima, i.e., the robustness to perturbations in the weight space (Keskar et al., 2017; Neyshabur et al., 2017; Dziugaite and Roy, 2017), or the compressibility of a neural network (Zhou et al., 2018; Arora et al., 2018; Kuhn et al., 2021). Alternatively, in the case of neural networks, there also exist instance-dependent bounds that directly depend on the norm of the weights (Bartlett et al., 2017; Golowich et al., 2020). Nonetheless, due to their inherent focus on a model’s parameters, the above results all suffer from the standard pitfalls of the over-parameterized setting so that the bounds become very loose once employed for larger networks. We argue that the generalization capability of a learner is directly influenced by the geometrical properties of the learned model in the data domain rather than the number or values of its constructing parameters. Following this idea, a body of work uses the properties of the classification margin (Antos et al., 2002; Sokolic et al., 2017; Jiang et al., 2019a; Soudry et al., 2018; Gunasekar et al., 2018) to quantify generalization. A common theme in such works is that the generalization ability of neural networks relies crucially on the optimization procedure and can not be solely described by the hypothesis class. Following this logic, Dziugaite and Roy (2017, 2018) adjust the training procedure so that it minimizes the bounds and, thereby, attain non-vacuous PAC-Bayesian bounds. While the bound of Dziugaite and Roy (2018) depends on a data-dependent prior, it considers generalization error with respect to a posterior distribution over neural network parameters. In contrast, we focus on the generalization properties of a single learning hypothesis (e.g., a single neural network) which is the result of training.

Our work also relates to approaches that quantify the local regularity of the learned prediction function. Examples of this are counting the number of linear regions of trained neural networks (Montufar et al., 2014), calculating the local Lipschitz constant of neural networks (Jordan and Dimakis, 2020; Herrera et al., 2020) or the local Rademacher complexity (Bartlett et al., 2005).

We also contribute to the literature of distributionally robust optimization, (cf. Sinha et al., 2018), since our bounds can be easily extended into a distributional robustness certificate (see Section 5). Our bound suggests that local Lipschitz estimators are more robust to distribution shifts, confirming recent results which control the global or local Lipschitz constants in order to achieve adversarially robust neural networks (Cisse et al., 2017; Salman et al., 2019; Cohen et al., 2019; Gouk et al., 2021; Anil et al., 2019; Muthukumar and Sulam, 2023). Similar to recent work on distributional robustness (Gao and Kleywegt, 2022; Kuhn et al., 2019; Cranko et al., 2021), we rely on a transport-based change of measure inequality. The majority of prior work only bounds the difference between the empirical validation error and the expected risk under distribution shift. In contrast, we present a much stronger result, which bounds the gap to the training error. Mohajerin Esfahani and Kuhn (2018); Staib and Jegelka (2019) also consider the gap to the training error, but only for a particular minimax estimator and,

in the latter case, only under much more restrictive smoothness assumptions. We provide a more in-depth comparison in Section 5, once the notation is formally set.

3. A Localized Bound on the Generalization Error

We consider datapoints (x, y) where $x \in \mathcal{X}$ are observed input features and $y \in \mathcal{Y}$ are target values/labels. To formulate the learning problem, we assume that the data is generated via an *unknown* probability measure $\mu \in \mathcal{P}(\mathcal{X} \times \mathcal{Y})$, where \mathcal{P} denotes the space of probability measures defined over $\mathcal{X} \times \mathcal{Y}$. Given a dataset $\mathcal{D}^N = \{(x_i, y_i)\}_{i=1}^N$ of i.i.d. draws from μ , the goal of supervised learning is to find a function \hat{f}^N which can accurately predict the targets. The quality of an estimator \hat{f}^N is measured through a loss function $\ell : \mathcal{Y} \times \mathcal{Y} \rightarrow \mathbb{R}$. Accordingly, we seek to attain a small *expected risk*, i.e., the expected loss under the data generating distribution

$$\mathfrak{R}(\hat{f}^N; \mu) := \mathbb{E}_{(x,y) \sim \mu} \left[\ell(\hat{f}^N(x), y) \right].$$

Since μ is unknown, it is not possible to directly evaluate $\mathfrak{R}(\hat{f}^N; \mu)$ given the training data \mathcal{D}^N . However, based on the dataset, we can compute the *empirical risk*

$$\hat{\mathfrak{R}}(\hat{f}^N) := \mathfrak{R}(\hat{f}^N; \mu^N) = \frac{1}{N} \sum_{n=1}^N \ell(\hat{f}^N(x_n), y_n).$$

which corresponds to the expected loss under an empirical measure $\mu^N := \frac{1}{N} \sum_{i=1}^N \delta_{(x_i, y_i)}$ for \mathcal{D}^N . Often $\hat{\mathfrak{R}}(\hat{f}^N)$ is also referred to as training error. In this work, we aim to bound the *generalization gap* $\mathfrak{R}(\hat{f}^N; \mu) - \hat{\mathfrak{R}}(\hat{f}^N)$. Importantly, as \hat{f}^N already depends on the data \mathcal{D}^N , $\hat{\mathfrak{R}}(\hat{f}^N)$ is a biased estimator of the expected risk $\mathfrak{R}(\hat{f}^N; \mu)$. Thus, standard results for the concentration of averages do not apply. Instead, to bound the generalization gap, we also need to take into consideration the learning hypothesis \hat{f}^N and quantify how well it generalizes from the training data \mathcal{D}^N to the general data distribution μ .

In the following, we introduce the basic assumptions and tools which form the foundation of our generalization bounds:

Assumption 1 *The domain \mathcal{X} is a compact subset of \mathbb{R}^d , the d -dimensional Euclidean space, and \mathcal{Y} is a compact subset of \mathbb{R} .*

The assumption that \mathcal{X} and \mathcal{Y} are compact is very common in statistical learning theory and implies that the bounded target values y are observed with a bounded noise. For instance, they are commonly used for uniform generalization bounds (e.g., Alon et al., 1997; Bartlett and Mendelson, 2002), PAC-Bayesian Bounds (e.g., McAllester, 1998; Catoni, 2007), and the more recent instance-dependent generalization bounds (e.g., Dziugaite and Roy, 2017; Neyshabur et al., 2018; Golowich et al., 2020).

In addition, we require geometric regularity assumptions on both the estimator and the loss function. For this purpose, we define the *local* Lipschitz constant of a function $g : \mathcal{X} \rightarrow \mathcal{Y}$ when restricted to $P \subset \mathcal{X}$ as,

$$\text{Lip}(g|P) := \sup_{\substack{x_1, x_2 \in P \\ x_1 \neq x_2}} \frac{|g(x_1) - g(x_2)|}{\|x_1 - x_2\|_2}.$$

For $0 \leq L < \infty$, we say that a function g is L -Lipschitz if the global Lipschitz constant $\text{Lip}(g) := \text{Lip}(g|\mathcal{X})$ is bounded by L . We assume that the learned function \hat{f} is Lipschitz continuous with Lipschitz constant $L_{\hat{f}}$:

Assumption 2 *There exists a constant $L_{\hat{f}} \geq 0$ such that the estimator \hat{f}^N almost surely satisfies $\text{Lip}(\hat{f}^N) \leq L_{\hat{f}}$.*

Lipschitz estimators are perhaps the most common class of estimators and include, Gaussian processes with non-smooth kernels and neural networks with popular activation functions such as ReLUs, ELUs and tanh functions. In Section 6, we extend our result to α -Hölder and smooth estimators. We also require a Lipschitz loss function:

Assumption 3 *The loss function $\ell : \mathcal{Y} \times \mathcal{Y} \rightarrow \mathbb{R}$ is L_ℓ -Lipschitz.*

Examples of Lipschitz continuous loss functions for classification are logit, hinge, or ramp loss (Hajek and Raginsky, 2019). A Lipschitz loss for regression is the Huber loss, which satisfies $L_\ell = 1$. This loss is commonly used for training neural networks (e.g., Morales, 2020; Meyer, 2021), since compared to the squared error loss, it is more robust to outliers and large gradients that destabilize training.

We take a localized approach, and instead of bounding the generalization error directly on the entire $\mathcal{X} \times \mathcal{Y}$ space, we first partition the space and then compare the empirical and expected risk separately on each element of this partitioning. A partitioning \mathbf{P} of size k is a collection $\{P_1, \dots, P_i, \dots, P_k\}$ subsets of $\mathcal{X} \times \mathcal{Y}$, where $P_i \cap P_j = \emptyset$ and $\cup_{i=1}^k P_i = \mathcal{X} \times \mathcal{Y}$, for every $1 \leq i < j \leq k$. Consequently, our analysis relies on two key localized notions: $\text{Lip}(\hat{f}^N|P)$, the local Lipschitz constant of the estimator restricted to a part $P \in \mathbf{P}$, and $\mu|_P$, the data generating distribution restricted to P , defined via $\mu|_P(\cdot) := \mu(\cdot \cap P)/\mu(P)$. The localized empirical distribution can be similarly defined as $\mu^N|_P(\cdot) := \mu^N(\cdot \cap P)/\mu^N(P)$. We note that $\mu^N(P) = N_P/N$ where $N_P := |\{\mathcal{D}^N \cap P\}|$ counts the number of samples which fall into the set P . We are now ready to present our instance-dependent bound on the generalization error. This theorem is a corollary of our main result of Theorem 10, and Appendix A.1 presents its proof.

Theorem 4 (Generalization error of Lipschitz estimators) *Let \hat{f}^N be a learned function which may depend on the dataset \mathcal{D}^N . Suppose Assumptions 1, 2, and 3 hold with some $L_\ell, L_{\hat{f}} > 0$. For any $0 < \delta \leq 1$ and any data-independent partitioning \mathbf{P} of $\mathcal{X} \times \mathcal{Y}$, we have*

$$\mathfrak{R}(\hat{f}^N; \mu) - \hat{\mathfrak{R}}(\hat{f}^N) \leq \text{cost}_{\text{transport}}(\mathbf{P}) + \text{err}_{\text{transport}}(\mathbf{P}) + \text{cost}_{\text{partition}}(\mathbf{P})$$

with probability greater than $1 - \delta$, where

$$\begin{aligned} \text{cost}_{\text{transport}}(\mathbf{P}) &:= \frac{C_{d+1,1}L_\ell}{N} \sum_{P \in \mathbf{P}} N_P^{\frac{d}{d+1}} \max\{1, \text{Lip}(\hat{f}^N|P_{\mathcal{X}})\} \text{diam}(P), \\ \text{err}_{\text{transport}}(\mathbf{P}) &:= \sqrt{\frac{\ln(4/\delta)}{N}} L_\ell \max\{1, L_{\hat{f}}\} \max_{P \in \mathbf{P}} \text{diam}(P), \\ \text{cost}_{\text{partition}}(\mathbf{P}) &:= \begin{cases} \|\ell\|_\infty \max\left\{\sqrt{\frac{2\ln(4/\delta)}{N}}, \sqrt{\frac{|P|}{N}}\right\} & |\mathbf{P}| > 1 \\ 0 & |\mathbf{P}| = 1 \end{cases} \end{aligned}$$

Here $P_{\mathcal{X}}$ denotes the projection of P onto \mathcal{X} , and the constant $C_{d+1,1}$ is recorded in Table 2.

The generalization gap is the discrepancy in calculating the expectation of the loss calculated with respect to the two distributions μ and μ^N . Intuitively, our bound is based on the cost of transporting probability mass from μ^N to μ . In particular, this cost accounts for how far we have to transport probability mass on average and how much the loss can change in the process. We perform this transport-based analysis locally by partitioning \mathcal{X} and bounding the cost of changing the measure from $\mu|_P$ to $\mu^N|_P$ for every $P \in \mathbf{P}$. Since the dataset \mathcal{D}^N is drawn at random, this cost is a random variable. The term $\text{cost}_{\text{transport}}$ upper bounds the expected value of this cost, and the term $\text{err}_{\text{transport}}$ controls the deviation from the expected cost. The last term, $\text{cost}_{\text{partition}}$, denotes the cost we pay for partitioning, and it is equal to zero if $\mathbf{P} = \{\mathcal{X} \times \mathcal{Y}\}$. The previous two terms account for transporting probability mass within parts of the domain. However, if $\mu(P) \neq \mu^N(P)$, mass also needs to be transported across parts. $\text{cost}_{\text{partition}}$ upper bounds the potential change in the risk due to this global transport of mass. Naturally, the more parts we have in our partitioning, the higher the $\text{cost}_{\text{partition}}$.

The error bound of Theorem 4 converges with $\mathcal{O}(N^{-1/(d+1)})$ which, for higher dimensional domains, implies relatively slow convergence. However, this rate is already an improvement upon Rademacher generalization bounds for Lipschitz estimators (see Section 4.1). We do not impose any constraints on μ other than having compact support. Thus, our bound holds for any $\mu \in \mathcal{P}(\mathcal{X} \times \mathcal{Y})$, also unfavorable edge-cases such as a uniform distribution over the domain. Without further assumptions, the optimal transport cost (i.e., Wasserstein distance) of μ^N to μ inherently has an exponential dependence on the dimensionality of the domain. For sufficiently smooth functions, we can remove the exponential dependence on d and obtain $\mathcal{O}(N^{-1/2})$ convergence rates (see Theorem 10 and Table 2).

In many applications with high-dimensional data domains, it has been postulated that the data lies on some low-dimensional manifold (Narayanan and Mitter, 2010; Fefferman et al., 2016). For instance, Pope et al. (2021) empirically demonstrate the validity of this assumption on popular image datasets. Under the assumption that the data lies on a \tilde{d} -dimensional manifold where $\tilde{d} \ll d$, we can improve the convergence rate. Proposition 5 shows that the generalization error would then *only* depend on the intrinsic dimension \tilde{d} . The proof is presented in Appendix A.2.

Proposition 5 (Fast rates for structured data) *Consider the setting and assumptions of Theorem 4. In addition, suppose μ is such that the data lies almost surely on a \tilde{d} -dimensional C^1 -Riemannian manifold. Then for any $0 < \delta \leq 1$ and any data-independent partitioning \mathbf{P} on $\mathcal{X} \times \mathcal{Y}$, there exists $C(\tilde{d})$ for which*

$$\begin{aligned} \mathfrak{R}(\hat{f}^N; \mu) - \hat{\mathfrak{R}}(\hat{f}^N) &\leq \frac{C(\tilde{d})L_\ell}{N} \sum_{P \in \mathbf{P}} N_P^{1-1/\tilde{d}} \max \left\{ 1, \text{Lip}(\hat{f}^N|_{P_{\mathcal{X}}}) \right\} \text{diam}(P) \\ &\quad + \text{err}_{\text{transport}}(\mathbf{P}) + \text{cost}_{\text{partition}}(\mathbf{P}) \end{aligned}$$

with probability $1 - \delta$ where the terms $\text{err}_{\text{transport}}$ and $\text{cost}_{\text{partition}}$ are as defined in Theorem 4.

Bounds of Theorem 4 and Proposition 5 can be made tighter by directly considering $\text{Lip}(\ell \circ \hat{f}^N|_P)$ the local Lipschitz regularity of the *composition* of the loss and the prediction function. In fact, a direct instantiation of Theorem 10 would result in a bound that depends on this quantity. However, for a clearer exposition, we split the Lipschitz constants of $\ell \circ \hat{f}^N$ using $\text{Lip}(\ell \circ \hat{f}^N|_P) \leq L_\ell \cdot \text{Lip}(\hat{f}^N|_P)$, and present the bounds in terms of the global Lipschitz constant of the loss. In our numerical experiments, we use the tighter variant based on $\text{Lip}(\ell \circ \hat{f}^N|_P)$.

4. Key Properties of the Generalization Bound: NN Perspective

Theorem 4 presents a generalization bound that captures the post-training properties of the learned prediction function \hat{f}^N . In this section, we elaborate on these properties with a focus on neural networks. As an empirical running example, we consider two simple supervised learning tasks. We generate synthetic random datasets for 1D regression and 2D binary classification (Fig. 7) and train over-parametrized fully-connected ReLU networks on them with stochastic gradient descent (SGD). We then evaluate the bound of Theorem 4 for the resulting estimator.¹ Details of the experiments are reserved for Appendix D. To calculate the local Lipschitz constant $\text{Lip}(\hat{f}|P)$, we simply consider a fine grid of the domain and evaluate the gradient of the network over this mesh. This only requires light computations since our toy examples are two-dimensional at most. For higher dimensional domains, Jordan and Dimakis (2020) and Fazlyab et al. (2019) propose scalable algorithms that approximate the local Lipschitz constant of a neural network. For the regression task, we use the Huber loss (Equation D.2). Since the Huber loss has a Lipschitz constant of $L_\ell = 1$, Theorem 4 applies directly to the regression case. For the binary classification, we use the labels $\mathcal{Y} = \{-1, 1\}$ and aim to bound the expected classification error $\mathbb{P}(\hat{f}^N(X) \neq Y)$. The 0-1 classification error $\mathbf{1}(\hat{f}^N(x) \cdot y < 0)$ is not Lipschitz. However, following Hajek and Raginsky (2019), we use the ramp loss

$$\ell_\gamma(y_1, y_2) := \min \left\{ 1, \left(1 - \frac{y_1 y_2}{\gamma} \right)_+ \right\}, \text{ with } \gamma > 0, \quad (1)$$

as a Lipschitz proxy and upper bound of the 0-1 loss. This allows us to obtain a corollary of Theorem 4 which upper bounds the classification error:

Corollary 6 (Classification error bound) *Consider a compact input domain and labels in $\mathcal{Y} = \{-1, 1\}$. Assume that the observation noise is i.i.d. and may only flip the label. Let $\gamma > 0$, \mathbf{P} be any partitioning of size k on \mathcal{X} , independent of the data \mathcal{D}^N . Then under Assumption 3, with probability greater than $1 - \delta$,*

$$\begin{aligned} \mathbb{P}(\hat{f}^N(X) \neq Y) &\leq \frac{1}{N} \sum_{i=1}^N \ell_\gamma(\hat{f}^N(X_i), Y_i) + \frac{2^{1/d} C_{d,1}}{\gamma} \sum_{P \in \mathbf{P}} \frac{N_P^{1-1/d}}{N} \text{Lip}(\hat{f}^N|P) \text{diam}(P) \\ &\quad + \sqrt{\frac{\ln(4/\delta)}{N}} \frac{L_{\hat{f}}}{\gamma} \max_{P \in \mathbf{P}} \text{diam}(P) + \sqrt{\frac{2}{N}} \max \left\{ \sqrt{\ln(4/\delta)}, \sqrt{k} \right\} \end{aligned}$$

here $N_P = |\{(X, Y) \in \mathcal{D}^N \text{ s.t. } X \in P\}|$ counts the number of samples that lie in partition P .

The proof of Corollary 6 is given in Appendix A.3. Since for classification \mathcal{Y} is only a finite set, we can marginally reduce the dimension dependence of the generic bound from $\mathcal{O}(N^{1-1/(d+1)})$ to $\mathcal{O}(N^{1-1/d})$.

1. More precisely, we visualize the tighter variant of Theorem 4 which directly depends on $\text{Lip}(\ell \circ \hat{f}^N|P)$, since splitting the constant as $L_\ell \cdot \text{Lip}(\hat{f}^N|P)$ may loosen the bound.

4.1 Instance-Dependent vs. Fixed Hypothesis Classes

Understanding the generalization of over-parametrized neural networks requires analyzing the *combination* of model architecture, initialization method, and training procedure. A trained network \hat{f}^N inherits the joint effect of the three elements. Therefore, instance-dependent bounds, which are calculated for \hat{f}^N post-hoc are more informative in describing the generalization behavior of over-parametrized networks. In particular, Theorem 4 can be understood as adapting the sub-hypothesis space in each partition P to the local Lipschitz regularity $L(\hat{f}^N|P)$. The adaptive hypothesis space that arises from the intersection of the sub-hypothesis spaces across partitions is tailored to the trained function \hat{f}^N and is typically much smaller than the hypothesis space considered a-priori. On the other hand, bounds that only reflect the properties of the function/hypothesis class prior to the training are known to be vacuous for large hypothesis classes such as neural networks (Bartlett and Long, 2021; Golowich et al., 2020).

We empirically evaluate our generalization bound when applied to neural networks trained on regression and classification tasks. Figures 2 and 5 show that contrary to the majority of prior works, which yield vacuous bounds for over-parametrized neural networks, our bound assumes values in the same order of magnitude as the expected error and becomes non-vacuous for around $N > 10000$ classification examples.

Furthermore, our instance-dependent bound reflects the positive effect of common regularization techniques on the generalization performance. It is known that certain training techniques, such as adversarial training, weight decay, and early stopping, can lead stochastic gradient descent to solutions that generalize better. Fig. 1 illustrates the effect of these methods on our bound when applied to neural networks. As we can observe in Fig. 1, the bound improves once the aforementioned regularization techniques are employed during training. In particular, for smaller sample sizes, the change in the value of the bound suggests that all three methods of adversarial training, weight decay, and early stopping produce networks that tend to generalize better. This observation matches the prior works of Xing et al. (2021), Krogh and Hertz (1991) and Li et al. (2020) respectively. This empirically supports our core idea that, since Theorem 4 directly depends on the learned neural network instance \hat{f}^N , it is able to capture the joint effect of model structure, initialization, and training.

4.2 Geometric vs. Parametric Characterization of the Estimator

We characterize the estimator via its local Lipschitz regularity. A key idea in our work is that this local geometry has an immediate effect on the generalization ability of the network compared to the network size or architecture. An alternative approach are parametric bounds which consider the network structure. Such bounds are data-dependent and are often a function of the Frobenius norm of the network’s weights, i.e., $\|\mathbf{W}_j\|_F$ where $1 \leq j \leq l$ indexes the layer number. Examples are the Rademacher-type bound of Neyshabur et al. (2015), Bartlett et al. (2017), which follows a covering number argument, and Neyshabur et al. (2018), which takes a PAC-Bayesian approach. These norm-based bounds roughly grow with $\mathcal{O}(\text{diam}(\mathcal{X})\text{Poly}(d, h, l) \prod_{j=1}^l \|\mathbf{W}_j\|_F \sqrt{1/N})$, where h denotes the width of the network.² Golowich et al. (2020) improve prior results and present a bounds of the rate

2. Not all the bounds have this polynomial dependency, e.g. the bound of Neyshabur et al. (2015) depends on the network depth exponentially.

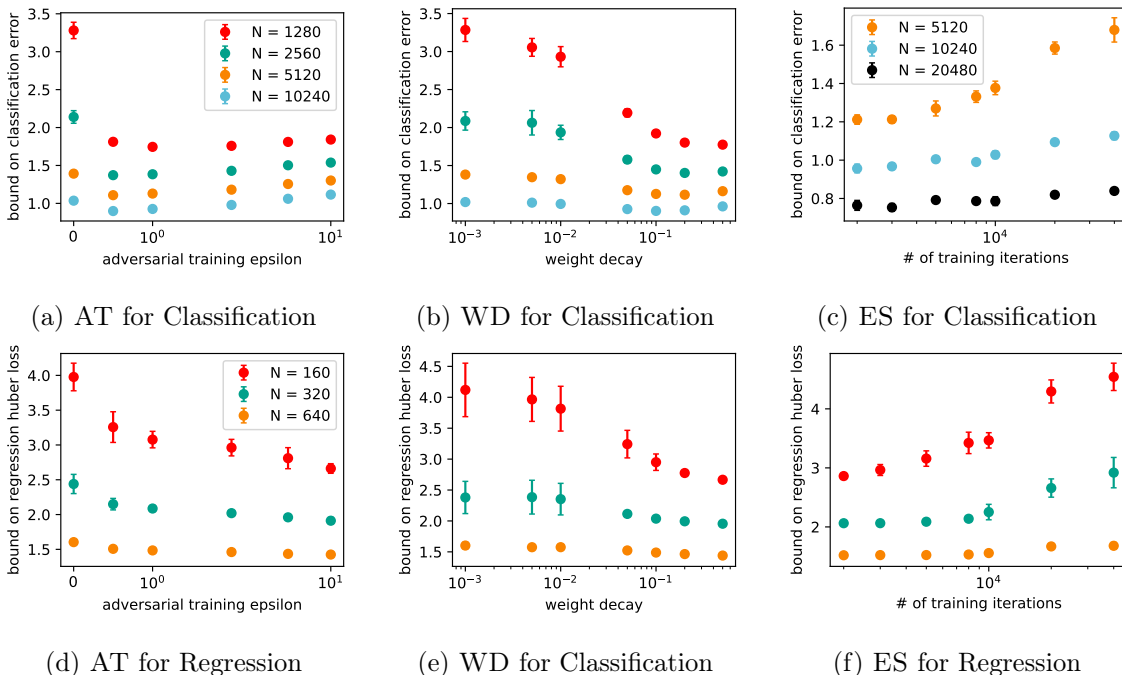


Figure 1: Effect of adversarial training (AT), weight decay (WD) and early stopping (ES). Generalization bounds of Theorem 4 and Corollary 6 suggest that these training techniques result in networks that generalize better. See Appendix D.3 for details of the plots.

$\mathcal{O}(\text{diam}(\mathcal{X}) \prod_{j=1}^l \|\mathbf{W}_j\|_F \sqrt{l/N})$. While these bounds have a polynomial dependency on the input dimension d , they quickly become vacuous for larger networks due to their polynomial dependence on network size. Therefore such bounds fail to capture or explain the benefit of over-parametrization (Bartlett and Long, 2021), in contrary to the observation that large neural networks tend to generalize better in practice (Zhang et al., 2017). A key issue with parametric analysis is that there are combinatorially many parameter configurations, or in cases even entire sub-spaces that correspond to the same neural network mapping. This artificially inflates the hypothesis space compared to the set of neural network functions that are actually considered by the learning algorithm.

Our bound is based on the geometry of the learned function rather than its parameters, and it does not suffer from the described issue. We observe that increasing the network size has negligible effects on the local regularity of the learned prediction function. Thus, the generalization bound of Theorem 4 hardly grows with the neural network size. To demonstrate this empirically, we train networks of increasingly larger width and depth and calculate the corresponding bounds. Fig. 2 shows that even increasing the number of neural network parameters by factor 1000 has only a minor effect on the value of the bound, in particular when the dataset size is large. Our geometry-based approach seems to capture the

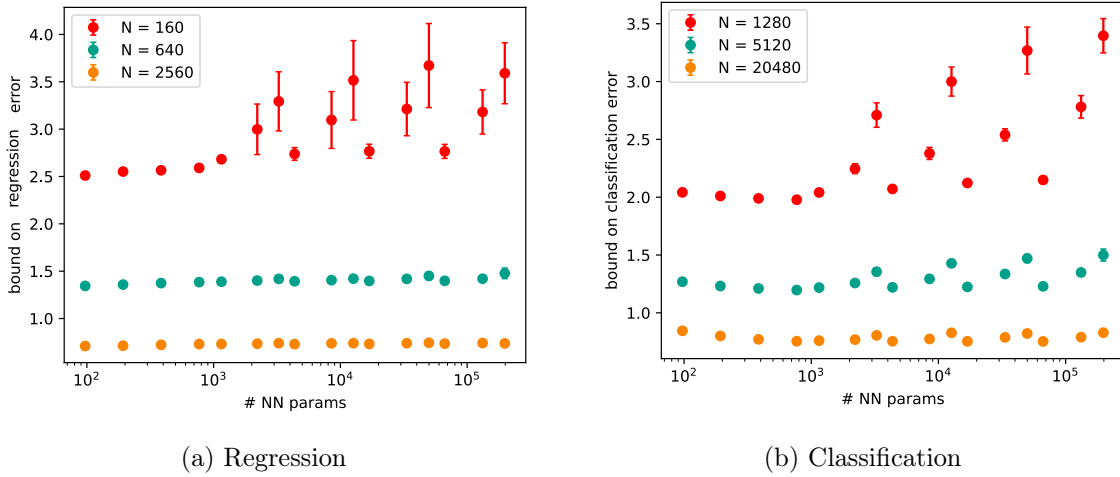


Figure 2: Generalization error bound vs. number of neural network parameters. The neural network size has only a minor effect on the bound values. See Appendix D.4 for details.

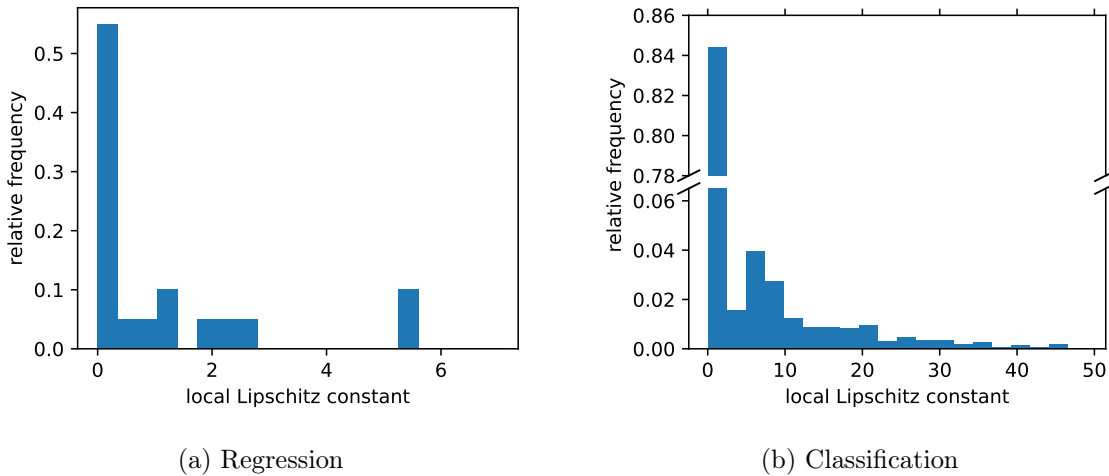


Figure 3: Frequency of local Lipschitz constant values per part $P \in \mathbf{P}$ for fully-connected ReLU nets trained with SGD. The local Lipschitz constant is small in the majority of parts P , contrary to the large global constant. For training details, see Appendix D.2.

empirically observed generalization behavior of over-parametrized neural networks better than previous bounds.

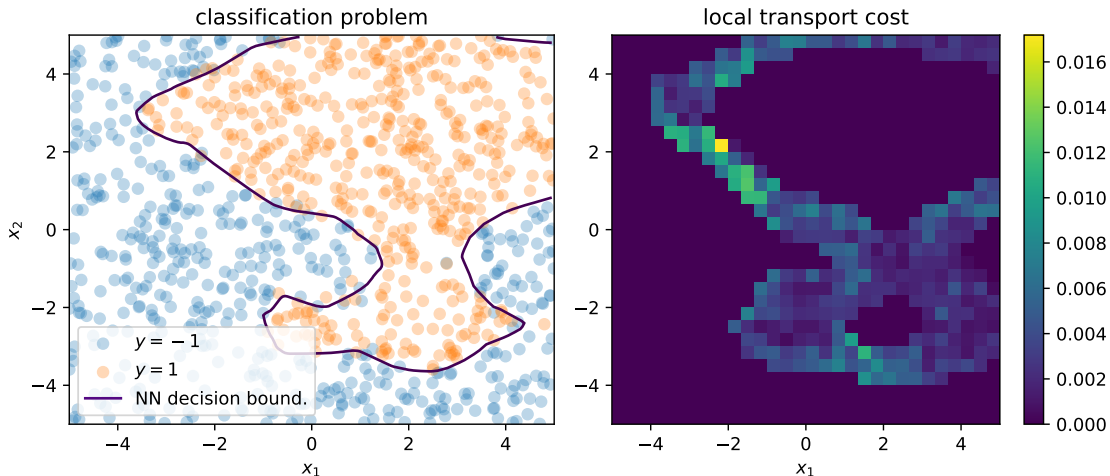
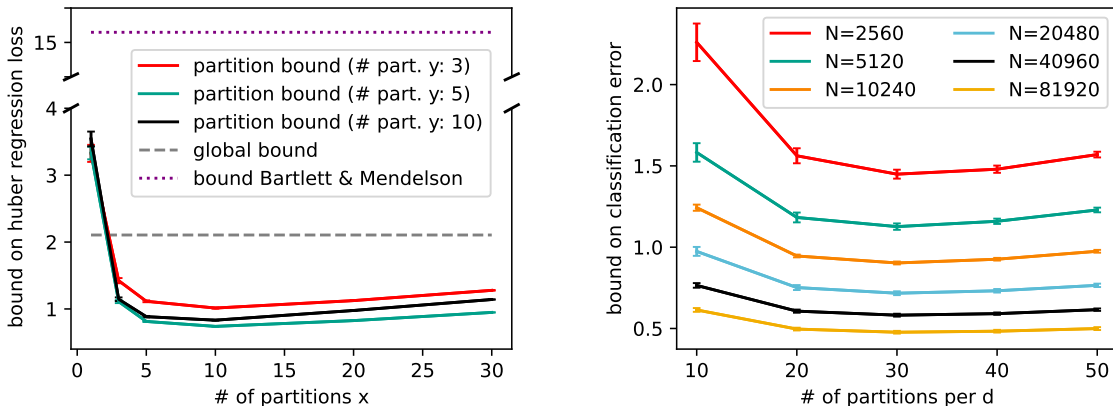


Figure 4: Local break down of $\text{cost}_{\text{transport}}$. Areas closer to the decision boundary of \hat{f}^N contribute more to the bound. Each cell P in the heatmap shows the local transport cost $\text{Lip}(\ell \circ \hat{f}^N|P)\text{diam}(P)\sqrt{N_P}/N$.

4.3 Localized vs. Global Analysis

In Theorem 4, we partition the domain into many subsets and locally bound the generalization risk when the domain is restricted to each subset. As a result, $\text{cost}_{\text{transport}}$, which is often the largest term of the bound and shrinks with N the slowest, is independent of the global Lipschitz constant and depends on \hat{f}^N only through $\text{Lip}(\hat{f}^N|P)$. When applied to neural networks, Theorem 4 tends to benefit from this localized analysis, as the local regularity of trained networks strongly varies across the domain. Figure 3 displays the distribution of local Lipschitz constants across parts $P \in \mathcal{P}$. We observe that $\text{Lip}(\hat{f}^N|P)$ is fairly low in the majority of parts, while there exist a few parts with much higher local Lipschitz constant, which contributes to a large overall global constant. Therefore, we expect our localized analysis to be more informative than a global argument which treats all parts uniformly and in turn, produces a bound depending on $\text{diam}(\mathcal{X})\text{Lip}(\hat{f}^N)$. Fig. 4 empirically supports this claim by visualizing the local generalization error restricted to each part P , which changes strongly across the neighborhoods. In particular, we observe that the local generalization error is small in areas away from the decision boundary of \hat{f}^N , and that it achieves its maximum in an area where the estimator misclassifies the training data.

As we partition the data domain into ever finer parts, two forces are at play: The diameter of each part P shrinks, and the local Lipschitz constant may get smaller, making $\text{cost}_{\text{transport}}$ and $\text{err}_{\text{transport}}$ in Theorem 4 shrink. At the same time, as parts become smaller, mismatches in probability mass of μ and μ^N become more pronounced so that we have to account for more transport of mass across the partitions, increasing the $\text{cost}_{\text{partition}}$ term. Hence, by making partitioning finer, we trade off $\text{cost}_{\text{transport}}$ and $\text{err}_{\text{transport}}$ against $\text{cost}_{\text{partition}}$. Figure 5 displays our bounds in response to an increasing partition size. We can empirically observe the trade-off as the bound values initially decrease and, as the partitioning becomes



(a) Regression. All curves correspond to $N = 2560$ samples. (b) Classification. Depending on N , the global bound is 8-20 times larger (see Table 1).

Figure 5: Generalization error bound values for a varying number of partitions. Left: Bound on the Huber regression loss for differing numbers of partitions in \mathcal{X} and \mathcal{Y} , together with the global Lipschitz bound in (2) the bound of Bartlett and Mendelson (2002). Right: Bound on the classification error for differing numbers of partitions per dimension of \mathcal{X} . For more details see Appendix D.5.

much fines, increase again. Since the marginal gains from a finer partitioning decrease, there is typically a sweet spot, i.e., a degree partitioning that leads to the tightest bounds.

# Train Data	Partition Bound (ours)	Global Bound (in 2)	Bound of B&M (2002)
2560	1.447 ± 0.028	18.407 ± 3.609	39.545 ± 2.951
5120	1.126 ± 0.019	10.234 ± 0.933	30.622 ± 1.096
10240	0.903 ± 0.009	7.023 ± 0.398	25.743 ± 0.606
20480	0.717 ± 0.012	5.232 ± 0.378	22.365 ± 0.665
40960	0.582 ± 0.008	3.813 ± 0.282	19.246 ± 0.600
81920	0.478 ± 0.009	2.856 ± 0.265	16.738 ± 0.658

Table 1: Comparison of bounds on the classification error. For the partition bound, we report to lowest bound values across a varying number of partitions. Our partition-based bounds are 5 to 10 times smaller than the global bound in (2) and at-least 20 times smaller than the uniform bound of Bartlett and Mendelson (2002).

Without partitioning, or equivalently by considering only one *global* partition $\mathbf{P}_{\text{global}} := \{\mathcal{X} \times \mathcal{Y}\}$, we obtain the *global* counterpart of Theorem 4:

$$\mathfrak{R}(\hat{f}^N; \mu) - \hat{\mathfrak{R}}(\hat{f}^N) \leq \text{cost}_{\text{transport}}(\mathbf{P}_{\text{global}}) + \text{err}_{\text{transport}}(\mathbf{P}_{\text{global}}). \quad (2)$$

We can also compare Theorem 4 with a classic uniform law on the generalization of Lipschitz estimators since our only assumption about the estimator is almost sure Lipschitz continuity. Let \mathcal{F}_L denote the class of L -Lipschitz functions mapping \mathcal{X} to \mathcal{Y} , and recall that \mathcal{D}^N is an i.i.d. random sample of size N drawn according to the probability distribution μ . Under Assumption 1 and 3, the Rademacher generalization bound (Theorem 9, Bartlett and Mendelson, 2002) implies that, with probability $1 - \delta$ there exists $C > 0$ for which every $f \in \mathcal{F}_L$ satisfies

$$\mathfrak{R}(f; \mu) - \hat{\mathfrak{R}}(f) \leq CL_\ell \left(\frac{(\text{diam}(\mathcal{X})L)^d d^2 D^2}{N} \right)^{1/(d+3)} + \|\ell\|_\infty \sqrt{\frac{8 \log 2/\delta}{N}} \quad (3)$$

where $D := \sup_{f \in \mathcal{F}_L} \|f\|_\infty$. In Appendix C, we formalize this statement, calculate C , and provide a proof for completeness. The first term on the right-hand-side of (3), which corresponds to the Rademacher complexity of \mathcal{F}_L , dominates this bound. It rapidly grows for large high-dimensional domains or for a large Lipschitz constant and converges at a $O(N^{-1/(d+3)})$ rate. Theorem 4 only marginally improves upon this rate since it converges with $O(N^{-1/(d+1)})$. However, the value of the constants is significantly smaller. The term $\text{cost}_{\text{transport}}$ has the slowest decay with N , and its constant is proportional to $\text{Lip}(\hat{f}^N | P) \text{diam}(P)$. Consequently, for a typical estimator \hat{f}^N , Theorem 4 yields a tighter bound than that of (2) and (3). This is further demonstrated in Fig. 5, which visualizes the bound values corresponding to Theorem 4 for a varying number of partitions and compares it to the global bound in (2), as well as to the global Rademacher bound of Eq. (3). For classification, we compare the bound values in Table 1. In particular, for the classification task, the global bound (8-20 times larger), as well as the Rademacher bound (> 20 times larger), are vacuous, while partitioning allows us to obtain non-vacuous guarantees.

Crucially, there exist estimators for which the generalization error of Theorem 4 is bounded and vanishes as $N \rightarrow \infty$, while the global inequality (2) diverges. In Proposition 7, we construct such an estimator. Perhaps surprisingly, we prove that the global bound diverges already for a shallow ReLU network with exactly one neuron defined over $\mathcal{X} = [0, 1]$, while a localized analysis with a partition with a size of order $|\mathbf{P}| = O(N^{0.6})$ gives a vanishing error bound. The proof is presented in Appendix A.4.

Proposition 7 (Partitioning can help) *Let $\mathcal{X} = [0, 1]$ and $\mathcal{Y} = [0, 1]$. For any probability measure $\mu \in \mathcal{P}(\mathcal{X} \times \mathcal{Y})$, there is an increasing sequence of Lipschitz constants $\{L_{f,N}\}_{N=1}^\infty$, a sequence of partitions $\{\mathbf{P}_N\}_{N=1}^\infty$, and a sequence of ReLU feedforward networks with one neuron $\{f_N\}_{N=1}^\infty$, such that the local bound of Theorem 4 for the tuple $(L_{f,N}, \mathbf{P}_N, f_N)$ converges:*

$$\lim_{N \rightarrow \infty} \text{cost}_{\text{transport}}(\mathbf{P}_N) + \text{err}_{\text{transport}}(\mathbf{P}_N) + \text{cost}_{\text{partition}}(\mathbf{P}_N) \rightarrow 0$$

while the global bound of Equation (2) diverges:

$$\lim_{N \rightarrow \infty} \text{cost}_{\text{transport}}(\mathbf{P}_{\text{global}}) + \text{err}_{\text{transport}}(\mathbf{P}_{\text{global}}) \rightarrow \infty.$$

where $\mathbf{P}_{\text{global}} = \{[0, 1] \times [0, 1]\}$, and the terms $\text{cost}_{\text{transport}}$, $\text{err}_{\text{transport}}$ and $\text{cost}_{\text{partition}}$ are defined as in Theorem 4.

5. Generalization under Distribution Shifts

A desirable characteristic of an estimator \hat{f}^N is robustness to changes in the data generating distribution μ between training and test time. A change in μ may be due to covariate shifts, adversarial attacks, or small changes in the data-generating process over time. In safety-critical applications such as perception systems in self-driving cars or models for medical diagnosis, it is crucial that we can certify the performance of \hat{f}^N under changes in the distribution. In this section, we employ our framework to obtain an instance-dependent generalization bound under distribution shift. In particular, we bound the risk calculated with respect to a shifted distribution μ^{adv} , by the *training* error of the estimator on a dataset of size N which is sampled from data generating distribution μ .

In addition to the transport of mass from μ^N to μ which is at the core of Theorem 4, our optimal transport-based approach allows us to seamlessly consider the additional change of measure from μ to μ^{adv} . In particular, we use the Wasserstein-1 distance $\mathcal{W}_1(\mu, \mu^{\text{adv}})$ to quantify the amount of distribution shift. This is consistent with previous work on robustness certificates which are often given for distributions within an ϵ -Wasserstein ball centered in the data generating distribution (Sinha et al., 2018; Lee and Raginsky, 2018; Blanchet and Murthy, 2019; Levine and Feizi, 2020; Gao and Kleywegt, 2022). The Wasserstein-1 distance is defined as the minimum ℓ_1 -cost of transporting probability mass from ν_1 to ν_2 , that

$$\mathcal{W}_1(\nu_1, \nu_2) := \inf_{\gamma \in \Gamma(\nu_1, \nu_2)} \int_{\mathcal{Z} \times \mathcal{Z}} \|x - x'\|_1 d(\gamma(x, x'))$$

where $\Gamma(\nu_1, \nu_2) \subset \mathcal{P}(\mathcal{Z} \times \mathcal{Z})$ denotes the set of all couplings between ν_1 and ν_2 , in other words, the set of joint distributions whose marginals are ν_1 and ν_2 .

Corollary 8 presents our instance-dependent generalization bound under distribution shift. For simplicity, we present this result in the global case of $\mathbf{P}_{\text{global}} = \{\mathcal{X} \times \mathcal{Y}\}$, i.e., without partitioning. However, the analysis can also be carried out locally with partitioning analogous to Theorem 4. The proof is given in Appendix A.5.

Corollary 8 (Locally Lipschitz estimators are robust to distribution shift) *Let $\mu^{\text{adv}} \in \mathcal{P}(\mathcal{X} \times \mathcal{Y})$. Under Assumptions 1, 2, and 3 with probability $\geq 1 - \delta$, we have*

$$\begin{aligned} \mathfrak{R}(\hat{f}^N; \mu^{\text{adv}}) - \hat{\mathfrak{R}}(\hat{f}^N) &\leq \text{cost}_{\text{transport}}(\mathbf{P}_{\text{global}}) + \text{err}_{\text{transport}}(\mathbf{P}_{\text{global}}) \\ &\quad + L_\ell \max\left(1, \text{Lip}(\hat{f}^N)\right) \mathcal{W}_1\left(\mu, \mu^{\text{adv}}\right) \end{aligned}$$

where \mathbf{P} denotes a data-independent partition on \mathcal{X} , and $(\text{cost}_{\text{transport}}, \text{err}_{\text{transport}})$ are identical to that of Theorem 4.

The bound of Corollary 8 has the same generalization terms as the global bound in Equation (2), plus an additional term which accounts for the potentially negative impact of the distribution shift by multiplying $\mathcal{W}_1(\mu, \mu^{\text{adv}})$ with the Lipschitz constants of the ℓ and \hat{f}^N . In particular, if $\mathcal{W}_1(\mu, \mu^{\text{adv}}) \leq \epsilon$, then the corollary implies that in the worst case, the estimator suffers from a $\epsilon L_\ell \text{Lip}(\hat{f}^N)$ increase in risk when evaluated on the perturbed distribution.

There are many connections between Corollary 8 and prior work on robust estimators. Gao (2022) verifies distributionally robust learnability of \mathcal{F}_L the class of Lipschitz functions

through a uniform bound. We expect this bound to be vacuous if calculated empirically, since it has large terms depending on the complexity of the function class, e.g., via Rademacher complexity or metric entropy. [Kuhn et al. \(2019\)](#) and [Cranko et al. \(2021\)](#) bound the difference between the finite-sample validation error of an estimator f , and $\mathfrak{R}(f; \mu^{\text{adv}})$. Both works use this robustness certificate to develop methods for distributionally robust optimization. Perhaps, closest to our result, is [Mohajerin Esfahani and Kuhn \(2018\)](#) who give a generalization bound for a data-dependent, robust estimator defined via

$$\hat{f}_\epsilon^N = \arg \min_{f \in \mathcal{F}_L} \max_{\mathcal{W}_1(\mu^N, \mu^{\text{adv}}) \leq \epsilon} \mathcal{R}(f; \mu^{\text{adv}}).$$

Taking a similar approach, [\(Staub and Jegelka, 2019\)](#) consider the data-dependent solution of an analogous minimax problem, where \mathcal{F}_L is replaced with a Reproducing Kernel Hilbert Space, and subsequently the Maximum Mean Discrepancy is used instead of the Wasserstein distance. In contrast to these works, [Corollary 8](#) holds for *any* data-dependent Lipschitz estimator. In practice, we have access to the training error $\hat{\mathfrak{R}}(\hat{f}^N)$ and aim to verify how this performance generalizes to unseen data generated from a shifted distribution ($\mathfrak{R}(f; \mu^{\text{adv}})$). Therefore, instance-dependent generalization bounds such as [Corollary 8](#) or [Mohajerin Esfahani and Kuhn \(2018\)](#) are of more practical relevance, compared to uniform ([Gao, 2022](#)) or deviation bounds ([Kuhn et al., 2019](#)). [Corollary 8](#) further suggests that (locally) Lipschitz estimators tend to be more robust towards distribution shifts, and contribute to prior results connecting distributional or adversarial robustness to Lipschitzness ([Cisse et al., 2017](#); [Finlay et al., 2018](#); [Sinha et al., 2018](#); [Anil et al., 2019](#)). [Corollary 8](#) does not depend on the number of model parameters. Hence, it gives a powerful guarantee when applied to over-parametrized neural networks, in particular when the data lies on a low-dimensional manifold. Therefore, it acts as an advocate for training methods that effectively regularize the Lipschitz constant of the network, (e.g., [Bartlett et al., 2017](#); [Cisse et al., 2017](#); [Anil et al., 2019](#); [Sagawa et al., 2020](#)).

6. Main Result

Our main result is a transport- and partitioning-based concentration inequality, which states that the empirical mean of a sample-dependent function concentrates around its expectation if the function satisfies some degree of regularity. Let $\mathcal{Z} \subset \mathbb{R}^{d_Z}$ denote the domain, and consider functions $g : \mathcal{Z} \rightarrow \mathbb{R}$. We work with two classes of regular functions, smooth and non-smooth. The \mathcal{C}^s -smooth class identifies functions that admit all partial derivatives up to order s , for an $s \in \mathbb{N}_+$. The smoothness of a \mathcal{C}^s -smooth function g , when restricted to $P \subset \mathcal{Z}$, is quantified by

$$\|g\|_{s:P} := \max_{|\beta| \leq s} \max_{z \in P} \left| \frac{\partial^{|\beta|} g(z)}{\partial \beta_1 \dots \partial \beta_{d_Z}} \right|,$$

where $\beta \in \mathbb{N}^{d_Z}$ is a multi-index and $|\beta| = \beta_1 + \dots + \beta_{d_Z}$. Further, when $P = \mathcal{Z}$, we simplify this notation to $\|g\|_s$. In machine learning applications, examples of smooth estimators include Gaussian processes with smooth kernels, physics-informed neural networks ([Raissi et al., 2019](#)), Neural-ODE solvers ([Chen et al., 2018](#)), invertible neural networks [Hyndman and Kratsios \(2021\)](#), and feedforward networks with smooth activation functions ([De Ryck et al., 2021](#)). For the non-smooth category, i.e., functions that are not differentiable, we use

the α -Hölder property as a geometric notion for quantifying regularity. More formally, for $0 < \alpha \leq 1$, a function g is α -Hölder if

$$\text{Lip}_\alpha(g|P) := \sup_{\substack{x_1, x_2 \in P \\ x_1 \neq x_2}} \frac{|g(x_1) - g(x_2)|}{\|x_1 - x_2\|_2^\alpha}.$$

is finite. For $\alpha = 1$ this recovers the case of Lipschitz functions, which are discussed in the previous sections. Further, when $P = \mathcal{Z}$, we simplify this notation to $\text{Lip}_\alpha(g)$ and, when $\alpha = 1$, we simplify the notation to $\text{Lip}(g|P)$. Setting $\alpha \in (0, 1)$ allows for rougher models which are typically used for making predictions from long time-series (Morrill et al., 2021) or rough paths of Neural-SDE models (Cuchiero et al., 2020). Further, we consider two classes of integral probability metrics to quantify how rapidly an empirical measure with i.i.d. samples from the data-generating distribution μ , concentrates around μ . Each metric evaluates the dissimilarity between any two probability measures as the worst-case average dissimilarity over all functions in a given class, e.g., Hölder or Smooth functions. This metric is equivalent to the Wasserstein-1 distance when defined over the class of Lipschitz (i.e., 1-Hölder) functions. We consider function classes of different regularity since, from the universal approximation standpoint, the performance of deep learning models is known to vary significantly across different smoothness classes Yarotsky and Zhevnerchuk (2020); Gühring and Raslan (2021). A discussion is provided in Section 6.1.

Definition 9 Let $\mathcal{Z} \subseteq \mathbb{R}^{d_z}$. For $\alpha \in (0, 1]$, we define the α -Hölder Wasserstein distance between two probability measures $\mu, \nu \in \mathcal{P}(\mathcal{Z})$ as

$$\mathcal{W}_\alpha(\mu, \nu) \stackrel{\text{def.}}{=} \sup_{g: \text{Lip}_\alpha(g) \leq 1} \int g(z) \mu(dz) - \int g(z) \nu(dz).$$

For $s \in [1, \infty)$, we define the s -smooth Wasserstein distance between two probability measures $\mu, \nu \in \mathcal{P}(\mathcal{Z})$ as

$$\mathcal{W}_{\mathcal{C}^s}(\mu, \nu) \stackrel{\text{def.}}{=} \sup_{g: \|g\|_s \leq 1} \int g(z) \mu(dz) - \int g(z) \nu(dz).$$

Theorem 10 formalizes our main result. In Section 6.2, we sketch the proof for Theorem 10 for the non-smooth case to highlight the main techniques. The complete proof can be found in Appendix A.6.

Theorem 10 Set $0 < \delta \leq 1$, and $N \in \mathbb{N}$. Let $\mathcal{Z} \subseteq \mathbb{R}^{d_z}$ be a compact set, $\mu \in \mathcal{P}(\mathcal{Z})$ be a probability measure, and \mathbf{P} a data-independent partitioning of any size $k \in \mathbb{N}$ on \mathcal{Z} . Suppose $g^N : \mathcal{Z} \mapsto \mathbb{R}$ is a real-valued random function that may depend on Z_1, \dots, Z_N , which are samples drawn independently from μ . For $\alpha \in (0, 1]$, define

$$\text{err}(\alpha) := \sqrt{\frac{\ln(4/\delta)}{N}} L \max_{P \in \mathbf{P}} \text{diam}(P)^\alpha + \frac{\|g^N\|_\infty}{\sqrt{N}} \max\{\sqrt{2 \ln(4/\delta)}, \sqrt{k}\}.$$

(i) **Non-Smooth:** Set $0 < \alpha \leq 1$, and let $\mathcal{F}_{L, \alpha} = \{g \in \mathcal{C}(\mathcal{Z}, \mathbb{R}) : \text{Lip}_\alpha(g) \leq L\}$. Suppose $g^N \in \mathcal{F}_{L, \alpha}$ almost surely. Then with probability greater than $1 - \delta$, we have

$$\mathbb{E}[g^N(Z)] - \frac{1}{N} \sum_{n=1}^N g^N(Z_n) \leq C_{d_z, \alpha} \sum_{P \in \mathbf{P}} \frac{N_P}{N} \text{rate}_{d_z, \alpha}(N_P) \text{diam}(P) \text{Lip}_\alpha(g^N|P) + \text{err}(\alpha)$$

where $\text{rate}_{d_{\mathcal{Z}},\alpha}$ and $C_{d_{\mathcal{Z}},\alpha}$ depend only on the Hölder coefficient α and on the dimension $d_{\mathcal{Z}}$. The explicit expressions are recorded in Table 2.

- (ii) **Smooth:** Set $s \geq 1$, and let $\mathcal{F}_{L,s} = \{g \in \mathcal{C}^s(\mathcal{Z}, \mathbb{R}) : \|g\|_s \leq L\}$. Suppose $g^N \in \mathcal{F}_{L,s}$ almost surely. Then there exists constant $C_{d_{\mathcal{Z}},s} > 0$ which with probability greater than $1 - \delta$ satisfies

$$\mathbb{E}[g^N(Z)] - \frac{1}{N} \sum_{n=1}^N g^N(Z_n) \leq C_{d_{\mathcal{Z}},s} \sum_{P \in \mathcal{P}} \frac{N_P}{N} \text{rate}_{d_{\mathcal{Z}},s}(N_P) \text{diam}(P) \|g^N\|_{s:P} + \text{err}(1)$$

where $\text{rate}_{d_{\mathcal{Z}},s}$ depends only on s and on $d_{\mathcal{Z}}$ and is recorded in Table 2.

Regularity	Dimension	Rate ($\text{rate}_{d_{\mathcal{Z}},\alpha}$, or $\text{rate}_{d_{\mathcal{Z}},s}$)	Constant ($C_{d_{\mathcal{Z}},\alpha}$ or $C_{d_{\mathcal{Z}},s}$)
α -Hölder	$d_{\mathcal{Z}} < 2\alpha$	$N_P^{-1/2}$	$C_{d_{\mathcal{Z}},\alpha} = \frac{d_{\mathcal{Z}}^{\alpha/2} 2^{d_{\mathcal{Z}}/2 - 2\alpha}}{1 - 2^{d_{\mathcal{Z}}/2 - \alpha}}$
	$d_{\mathcal{Z}} = 2\alpha$	$(\alpha 2^{\alpha+2} + \log_2(N_P)) N_P^{-1/2}$	$C_{d_{\mathcal{Z}},\alpha} = \frac{d_{\mathcal{Z}}^{\alpha/2}}{\alpha 2^{\alpha+1}}$
	$d_{\mathcal{Z}} > 2\alpha$	$N_P^{-\alpha/d_{\mathcal{Z}}}$	$C_{d_{\mathcal{Z}},\alpha} = 2 \left(\frac{d_{\mathcal{Z}} - \alpha}{2\alpha(1 - 2^{\alpha - d_{\mathcal{Z}}/2})} \right)^{2\alpha/d_{\mathcal{Z}}} \left(1 + \frac{\alpha}{2\alpha(\frac{d_{\mathcal{Z}}}{2} - \alpha)} \right) d_{\mathcal{Z}}^{\alpha/2}$
s -Smooth	$s > \frac{d_{\mathcal{Z}}}{2}$	$N_P^{-1/2}$	$\exists C_{d_{\mathcal{Z}},s} > 0$
	$s = \frac{d_{\mathcal{Z}}}{2}$	$(\log(N_P) + 1) N_P^{-1/2}$	$\exists C_{d_{\mathcal{Z}},s} > 0$
	$s < \frac{d_{\mathcal{Z}}}{2}$	$N_P^{-s/d_{\mathcal{Z}}}$	$\exists C_{d_{\mathcal{Z}},s} > 0$

Table 2: Rates and constants for Theorem 10

The generalization bounds of Sections 3, 4 and 5 all follow from Theorem 10 in the non-smooth case with $\alpha = 1$, so that the α -Hölder regularity coincides with Lipschitzness. Since we are concerned with the loss of a machine learning estimator, we use $g^N(Z) = g^N(X, Y) = \ell(\hat{f}^N(X), Y)$ to obtain the risk bounds.

6.1 Optimal Transport Interpretation

The risk bounds derived in Theorem 10 case (i) for α -Hölder functions, where $0 < \alpha \leq 1$ are closely connected to optimal transport theory. In particular, the Kantorovich-Rubinstein duality (e.g. see Dudley, 2002, Theorem 11.8.2) links the \mathcal{W}_1 distance from Definition 9, to the optimal transport problem as follows: for any $\mu, \nu \in \mathcal{P}(\mathcal{Z})$ we have

$$\mathcal{W}_1(\mu, \nu) = \inf_{\pi \in \Pi(\mu, \nu)} \mathbb{E}_{(X,Y) \sim \pi} [\|X - Y\|] \quad (4)$$

where $\Pi(\mu, \nu)$ is the set of couplings of μ and ν . A coupling $\pi \in \mathcal{P}(\mathcal{Z} \times \mathcal{Z})$ of μ and ν is a joint-probability measure whose marginals coincide with μ and ν . The right-hand side of (4) is the standard definition of the 1-Wasserstein distance (e.g., Villani (2009, Definition 6.1)).

A crucial property that allows us to similarly connect \mathcal{W}_α for $0 < \alpha \leq 1$ to optimal transport is the following: A function $f : \mathcal{Z} \rightarrow \mathbb{R}$ is α -Hölder when \mathcal{Z} is equipped with the Euclidean metric $\text{dist}(x, y) \stackrel{\text{def}}{=} \|x - y\|$, where $x, y \in \mathcal{Z}$, if and only if f is Lipschitz with respect to the *snowflaked Euclidean metric* on \mathcal{Z} defined for any $x, y \in \mathcal{Z}$ by

$$\text{dist}_\alpha(x, y) \stackrel{\text{def}}{=} \|x, y\|^\alpha$$

(see [Weaver, 2018](#), Proposition 2.52). Since $(\mathcal{Z}, \text{dist}_\alpha)$ is a metric space, the Kantorovich-Rubinstein duality for $(\mathcal{Z}, \text{dist}_\alpha)$ implies that

$$\mathcal{W}_\alpha(\mu, \nu) = \inf_{\pi \in \Pi(\mu, \nu)} \mathbb{E}_{(X, Y) \sim \pi} [\|X - Y\|^\alpha] \quad (5)$$

for any $\mu, \nu \in \mathcal{P}_\alpha(\mathcal{Z})$, where $\mathcal{P}_\alpha(\mathcal{Z})$ is the set of Borel probability measures on \mathcal{Z} for which $\mathbb{E}_{X \sim \nu} [\|X - z\|^\alpha] < \infty \forall z \in \mathcal{Z}$. Since \mathcal{Z} is compact, every Borel probability measure on \mathcal{Z} belongs to $\mathcal{P}_\alpha(\mathcal{Z})$ and, thus, (5) holds for any $\mu, \nu \in \mathcal{P}(\mathcal{Z})$. Hence, under mild regularity conditions, the α -Hölder Wasserstein distance from Definition 9 directly corresponds to the optimal transport (aka. Wasserstein) distance with transportation cost given by $\text{dist}_\alpha(X, Y)$.

6.2 Proof outline for Theorem 10

Since the function g^N is dependent on \mathcal{D}^N , we can not directly invoke concentration inequalities for bounded i.i.d. random variables such as, e.g., Hoeffding’s inequality. Potentially, g^N is over-fitted to the dataset \mathcal{D}^N such that the empirical risk is much smaller than the expected risk. Thus, our analysis has to consider relevant properties of g^N that capture how well we can expect g^N to generalize beyond the training data \mathcal{D}^N .

Crucially, given arbitrary but fixed measures $\nu_1 = \mu$ and $\nu_2 = \mu^N$, by Definition 9

$$\int_{z \in \mathcal{Z}} g(z) \nu_1(dz) \leq \text{Lip}_\alpha(g) \mathcal{W}_\alpha(\nu_1, \nu_2) + \int_{z \in \mathcal{Z}} g(z) \nu_2(dz), \quad (6)$$

for any g s.t. $\text{Lip}_\alpha(g) < \infty$, including any function that depends on μ^N . The Lipschitz constant $\text{Lip}(g^N)$ measures how regular g^N is and quantifies the maximum change of g when moving probability mass. Since the local regularity of g^N may vary strongly throughout the domain, we use a partitioning \mathbf{P} of the space \mathcal{Z} to obtain a *localized* variant of the change of measure inequality. Then we separately bound the sum of local transport costs as well as the potential error from partitioning with high probability. In the following, we elaborate on the three main steps of this proof.

Step 1: Local change of measure. To localize the change of measure inequality, we use a partitioning \mathbf{P} of the space \mathcal{Z} . In particular, we invoke (6) independently on each part $P \in \mathbf{P}$ by restricting the domain to P and using the lemma with the corresponding restricted measures $\nu_1 = \mu|_P$ and $\nu_2 = \mu^N|_P$. Taking a sum over all $P \in \mathbf{P}$ results in

$$\begin{aligned} \int_{z \in \mathcal{Z}} g^N(z) \mu(dz) &\leq \int_{z \in \mathcal{Z}} g^N(z) \mu^N(dz) + \overbrace{\sum_{P \in \mathbf{P}} \mu^N(P) \text{Lip}_\alpha(g^N|_P) \mathcal{W}_\alpha(\mu|_P, \mu^N|_P)}^{(I)} \\ &\quad + \underbrace{\sum_{P \in \mathbf{P}} \left(1 - \frac{\mu^N(P)}{\mu(P)}\right) \int_{z \in P} g^N(z) \mu(dz)}_{(II)}. \end{aligned}$$

Here, term (I) bounds the change of expectation from locally moving mass within each partition. Term (II) appears due to the potential mismatch in probability mass of $\mu(P)$ and $\mu^N(P)$ in the parts of the partitioning of \mathcal{Z} and thus accounts for probability mass that would have to be re-distributed across parts. We bound (I) and (II) separately.

Step 2: Bounding (I). To bound (I), the crucial element is to upper-bound the Wasserstein distance between μ and its empirical counterpart μ^N . Since μ^N is based on samples drawn independently from μ , $\mathcal{W}_\alpha(\mu, \mu^N)$ is bounded in expectation and concentrates as more samples are taken into account. More formally, we show that due to Kloeckner (2020, Theorem 2.1), for all $\epsilon > 0$, and $N \in \mathbb{N}$ it holds that

$$\mathbb{P}\left(\left|\mathcal{W}_\alpha(\mu, \mu^N) - \mathbb{E}\left[\mathcal{W}_\alpha(\mu, \mu^N)\right]\right| \geq \epsilon\right) \leq 2e^{-\frac{2N\epsilon^2}{\text{diam}(\mathcal{Z})^{2\alpha}}},$$

and

$$\mathbb{E}\left[\mathcal{W}_\alpha(\mu, \mu^N)\right] \leq C_{d_{\mathcal{Z}}, \alpha} \text{diam}(\mathcal{Z}) \text{rate}_{d_{\mathcal{Z}}, \alpha}(N)$$

where the rates and the constants are as in Table 2. We apply the above inequalities locally, by considering $\mathcal{W}_\alpha(\mu|_P, \mu^N|_P)$ and sum over all $P \in \mathbf{P}$. We show that this summation is a weighted sum of independent sub-Gaussian random variables, which we control via Lemma 22. This treatment results in an upper bound for (I).

Step 3: Bounding (II). We interpret this term as the penalty we face for partitioning. It is zero if the probability mass of the data generating and empirical measures match across the partitioning, i.e. $\mu(P) = \mu^N(P)$ for all $P \in \mathbf{P}$, or if the analysis is carried out globally, i.e. $\mathbf{P} = \{\mathcal{Z}\}$. Considering discretized measures $\tilde{\mu}, \tilde{\mu}^N \in \mathcal{P}(\mathbf{P})$, which satisfy $\tilde{\mu}(\{P\}) = \mu(P)$ and $\tilde{\mu}^N(\{P\}) = \mu^N(P)$, we use that

$$\sum_{P \in \mathbf{P}} \left(1 - \frac{\mu^N(P)}{\mu(P)}\right) \int_{z \in P} g^N(z) \mu(dz) \leq \|g\|_\infty \text{TV}(\tilde{\mu}, \tilde{\mu}^N)$$

where $\text{TV}(\tilde{\mu}, \tilde{\mu}^N)$ denotes the total variation distance between $\tilde{\mu}^N$ and $\tilde{\mu}$. Since the $\text{TV}(\tilde{\mu}, \tilde{\mu}^N)$ concentrates around zero and we can bound term II with high probability. Combining these three steps concludes the proof.

7. Conclusion

We presented novel instance-dependent generalization bounds for locally regular estimators. We empirically and theoretically demonstrated the benefits of an instance-dependent non-parametric bound and the effectiveness of a localized treatment of the risk. In particular, we showed that the instance-dependent bound remains relatively tight for over-parametrized models and captures a number of neural network generalization phenomena. In contrast, existing uniform or data-dependent parametric bounds tend to explode for large neural networks and fail to explain their good generalization behavior.

Key observations made in this work could be relevant for future work that aims to improve learning algorithms. For example, our result suggests that the *local* regularity of a model plays a crucial role in its generalization ability and robustness to distribution shifts. This might be of interest for developing robust and regularized training techniques. Finally, our optimal-transport-based approach constitutes a novel avenue towards theoretically analyzing generalization in machine learning. We introduce the necessary technical tools and proof methodology, hoping that future work can further explore this avenue and improve our results.

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Appendix A. Proofs

A.1 Proof of the Main Generalization Bound (Theorem 4)

Proof of Theorem 4: Let $\mathcal{Z} = (\mathcal{X}, \mathcal{Y})$, $d_{\mathcal{Z}} = d + 1$ and $g^N(x, y) = \ell(\hat{f}^N(x), y)$. Notice that $\nabla_y g^N = \nabla_{y_2} \ell \leq L_\ell$ and $\nabla_x g^N = \nabla_{y_1} \ell \cdot \nabla_x \hat{f}^N \leq L_\ell \cdot \text{Lip}(\hat{f}^N)$. So $\text{Lip}(g^N|P) \leq L_\ell \max\{1, \text{Lip}(\hat{f}^N|P_{\mathcal{X}})\} \leq L_\ell \max\{1, L_f\}$ for all $P \in \mathbf{P}$. Now we can apply Theorem 10 with $\alpha = 1$, $\mathcal{F}_{L,1} = \{g \in \mathcal{C}(\mathcal{Z}, \mathbb{R}) : \text{Lip}(g) \leq L = L_\ell \max\{1, L_f\}\}$. Then for all partitions \mathbf{P} with $|\mathbf{P}| \leq k$, for all $0 < \delta \leq 1$ and $N \in \mathbb{N}$, there exists an explicit constant $C_{d_{\mathcal{Z}},1} > 0$ s.t. with probability greater than $1 - \delta$

$$\begin{aligned}
 & \mathfrak{R}(\hat{f}^N; \mu) - \hat{\mathfrak{R}}(\hat{f}^N) \\
 &= \mathbb{E}[g^N(Z)] - \frac{1}{N} \sum_{n=1}^N g^N(Z_n) \\
 &\leq C_{d_{\mathcal{Z}},\alpha} \sum_{P \in \mathbf{P}} \frac{N_P}{N} \text{rate}_{d_{\mathcal{Z}},\alpha}(N_P) \text{diam}(P) \text{Lip}(g^N|P) + \epsilon \\
 &\leq C_{d_{\mathcal{Z}},\alpha} \sum_{P \in \mathbf{P}} \frac{N_P}{N} \text{rate}_{d_{\mathcal{Z}},\alpha}(N_P) \text{diam}(P) L_\ell \max\{1, \text{Lip}(\hat{f}^N|P_{\mathcal{X}})\} + \epsilon \\
 &= \frac{C_{d+1,1} L_\ell}{N} \sum_{P \in \mathbf{P}} N_P^{\frac{d}{d+1}} \max\{1, \text{Lip}(\hat{f}^N|P_{\mathcal{X}})\} \text{diam}(P) + \epsilon \\
 &= \text{cost}_{\text{transport}} + \epsilon,
 \end{aligned} \tag{A.1}$$

where

$$\begin{aligned}
 \epsilon &:= \sqrt{\frac{\ln(4/\delta)}{N}} L_\ell \max\{1, L_f\} \max_{P \in \mathbf{P}} \text{diam}(P) + \frac{\|\ell\|_\infty}{\sqrt{N}} \max\{\sqrt{2 \ln(4/\delta)}, \sqrt{k}\} \\
 &= \text{err}_{\text{transport}} + \text{cost}_{\text{partition}}.
 \end{aligned} \tag{A.2}$$

■

A.2 Proof of Generalization Bound on Manifold Domain (Proposition 5)

We start by proving the manifold extension of our main result in Theorem 10. Then present the proof of Theorem 5 as a corollary of this theorem.

Theorem 11 (Concentration of measure on a compact manifold) *Set $0 < \delta \leq 1$, and $N \in \mathbb{N}$. Let \mathcal{Z} be a $d_{\mathcal{Z}}$ -dimensional compact class C^1 Riemannian manifold. Let μ be a Borel probability measure on \mathcal{Z} , and \mathbf{P} a partition of size k on \mathcal{Z} . Suppose g^N is a real-valued random function on \mathcal{Z} depending on Z_1, \dots, Z_N . Let $\mathcal{F}_L = \{g \in \mathcal{C}(\mathcal{Z}, \mathbb{R}) : \text{Lip}(g) \leq L\}$. Suppose $g^N \in \mathcal{F}_L$ almost surely. Then with probability greater than $1 - \delta$*

$$\mathbb{E}[g^N(Z)] - \frac{1}{N} \sum_{n=1}^N g^N(Z_n) \leq C_{\mathcal{Z}} \sum_{P \in \mathbf{P}} \frac{N_P^{1-1/d_{\mathcal{Z}}}}{N} \text{diam}(P) \text{Lip}(g^N|P) + \text{err} \tag{A.3}$$

where $C_Z > 0$ is a constant depending on d_Z and

$$\text{err} := \sqrt{\frac{\ln(4/\delta)}{N}} L \max_{P \in \mathbf{P}} \text{diam}(P) + \frac{\|g^N\|_\infty}{\sqrt{N}} \max\{\sqrt{2 \ln(4/\delta)}, \sqrt{k}\}.$$

Proof of Theorem 11. All the steps are similar to the proof of Theorem 10, and we only invoke a different Wasserstein concentration lemma. Deploying Lemma 18 we find that there exists $C_Z > 0$ for which

$$\sum_{P \in \mathbf{P}} \mu^N(P) \text{Lip}(g^N|P) \mathbb{E}[\mathcal{W}_1(\mu|_P, \mu^N|_P)|N_P] \leq C_Z \sum_{P \in \mathbf{P}} \mu^N(P) \text{Lip}(g^N|P) \text{diam}(P) N_P^{-1/d_Z}$$

and that

$$\mathbb{P}\left(\left|\mathcal{W}_1(\mu|_P, \mu^N|_P) - \mathbb{E}[\mathcal{W}_1(\mu|_P, \mu^N|_P)|N_P]\right| \geq \epsilon \mid N_P\right) \leq 2e^{-\frac{2N_P \epsilon^2}{\text{diam}(P)^2}}.$$

By defining X_P similar to proof of Theorem 10 and invoking Lemma 22, we get that there exists $C_Z > 0$ such that the following holds with probability $1 - \delta_1$

$$\begin{aligned} \mathcal{B}^N &\leq C_Z \sum_{P \in \mathbf{P}} \mu^N(P) \text{Lip}(g^N|P) \text{diam}(P) N_P^{-1/d_Z} \\ &\quad + L \max_{P \in \mathbf{P}} \text{diam}(P) \left(\frac{\ln(2/\delta_1)}{N}\right)^{1/2}. \end{aligned}$$

Terms \mathcal{E}^N and \mathcal{R}^N are identical to proof of Theorem 10. Therefore, plugging in everything we get,

$$\mathbb{E}[g^N(Z)] - \frac{1}{N} \sum_{n=1}^N g^N(Z_n) \leq C_Z \sum_{P \in \mathbf{P}} \frac{N_P^{1-1/d_Z}}{N} \text{diam}(P) \text{Lip}(g^N|P) + \text{err}.$$

■

Proof of Proposition 5 The proof is nearly identical to Theorem 4, however, here we invoke Theorem 11 instead of Theorem 10. For completeness we repeat some of the steps. Let \mathcal{Z} be the manifold which denotes the support of μ , then $d_Z = \tilde{d}$. Let $g^N(x, y) = \ell(\hat{f}^N(x), y)$. We then apply Theorem 11 with $\mathcal{F}_L := \{g : \mathcal{Z} \rightarrow \mathbb{R} : \text{Lip}(g) \leq L = L_\ell \max\{1, L_{\hat{f}}\}\}$. Then for all partition \mathbf{P} with $|\mathbf{P}| \leq k$, for all $0 < \delta \leq 1$, $N \in \mathbb{N}$, there exists an $C_Z > 0$ s.t. with probability greater than $1 - \delta$

$$\begin{aligned} \mathfrak{R}(\hat{f}^N; \mu) - \hat{\mathfrak{R}}(\hat{f}^N) &= \mathbb{E}[g^N(Z)] - \frac{1}{N} \sum_{n=1}^N g^N(Z_n) \\ &\leq C_Z \sum_{P \in \mathbf{P}} \frac{N_P^{1-1/d_Z}}{N} \text{diam}(P) \text{Lip}(g^N|P) + \text{err} \\ &= \frac{C(\tilde{d})L_\ell}{N} \sum_{P \in \mathbf{P}} N_P^{1-1/\tilde{d}} \max\{1, \text{Lip}(\hat{f}^N|P_{\mathcal{X}})\} \text{diam}(P) \\ &\quad + \text{err}_{\text{transport}}(\mathbf{P}) + \text{cost}_{\text{partition}}(\mathbf{P}) \end{aligned}$$

where the last line is obtained by using the definition of $\text{err}_{\text{transport}}$ and $\text{cost}_{\text{partition}}$ (as defined in Theorem 4). In addition, to transparently show the dependency of $C_{\mathcal{Z}}$ on \tilde{d} , we have renamed the constant. \blacksquare

A.3 Proof of the Classification Error Bound (Corollary 6)

Proof of Corollary 6 Let $g^N((x, y)) := \ell_\gamma(\hat{f}^N(x), y)$, where $y \in \{-1, 1\}$ and $x \in \mathcal{X} \subset \mathbb{R}^d$. Consider a partition $\mathbf{P} \in \mathcal{X}$ of size k . For all $P \in \mathbf{P}$, we may define two sets

$$P_{(-)} := \{(x, -1), \forall x \in P\} \quad P_{(+)} := \{(x, +1), \forall x \in P\}.$$

Note that $\text{diam}(P_{(-)}) = \text{diam}(P_{(+)}) = \text{diam}(P)$. We construct the \mathbf{P}_\pm partition on $\mathcal{X} \times \mathcal{Y}$ as

$$\mathbf{P}_\pm = \{P_{(-)} \mid P \in \mathbf{P}\} \cup \{P_{(+)} \mid P \in \mathbf{P}\}.$$

Note that $|\mathbf{P}_\pm| = 2|\mathbf{P}|$. For some $P_{(+)} \in \mathbf{P}_\pm$, we calculate $\text{Lip}(g|P_{(+)})$,

$$\text{Lip}(g|P_{(+)}) = \text{Lip}\left(\ell_\gamma\left(\hat{f}^N(\cdot), +1\right) \Big| P\right) \leq \frac{1}{\gamma} \text{Lip}(\hat{f}^N|P)$$

and similarly for any $P_{(-)} \in \mathbf{P}_\pm$. Now invoking Theorem 10, for g and \mathbf{P}_\pm we get, with probability greater than $1 - \delta$

$$\begin{aligned} \mathfrak{R}_\gamma(\hat{f}^N; \mu) - \hat{\mathfrak{R}}_\gamma(\hat{f}^N) &\leq C_{d,1} \sum_{P_\pm \in \mathbf{P}_\pm} \frac{N_{P_\pm}^{1-1/d}}{N} \text{diam}(P_\pm) \text{Lip}(g^N|P_\pm) + \text{err} \\ &= C_{d,1} \sum_{P \in \mathbf{P}} \left[\frac{N_{P_{(+)}}^{1-1/d}}{N} \text{diam}(P_{(+)}) \text{Lip}(g^N|P_{(+)}) \right. \\ &\quad \left. + \frac{N_{P_{(-)}}^{1-1/d}}{N} \text{diam}(P_{(-)}) \text{Lip}(g^N|P_{(-)}) \right] + \text{err} \\ &\leq C_{d,1} \sum_{P \in \mathbf{P}} \frac{N_{P_{(+)}}^{1-1/d} + N_{P_{(-)}}^{1-1/d}}{N} \text{diam}(P) \frac{\text{Lip}(\hat{f}^N|P)}{\gamma} + \text{err} \\ &\leq 2^{1/d} C_{d,1} \sum_{P \in \mathbf{P}} \frac{N_P^{1-1/d}}{N} \text{diam}(P) \frac{\text{Lip}(\hat{f}^N|P)}{\gamma} + \text{err} \end{aligned}$$

where $N_P = N_{P_{(+)}} + N_{P_{(-)}}$ and

$$\begin{aligned} \text{err} &= \sqrt{\frac{\ln(4/\delta)}{N}} \text{Lip}(g^N) \max_{P_\pm \in \mathbf{P}_\pm} \text{diam}(P_\pm) + \frac{1}{\sqrt{N}} \max\{\sqrt{2 \ln(4/\delta)}, \sqrt{|\mathbf{P}_\pm|}\} \\ &= \sqrt{\frac{\ln(4/\delta)}{N}} \frac{\text{Lip}(\hat{f}^N)}{\gamma} \max_{P \in \mathbf{P}} \text{diam}(P) + \frac{1}{\sqrt{N}} \max\{\sqrt{2 \ln(4/\delta)}, \sqrt{2k}\}. \end{aligned}$$

The ramp loss acts as a Lipschitz proxy of the zero-one loss and allows us to analyze the classification error defined via

$$\mathfrak{R}_{01}(\hat{f}^N; \mu) := \mathbb{E}_{(X,Y) \sim \mu} \ell_{01}(\hat{f}^N(X), Y) = \mathbb{P}(\hat{f}^N(X) \neq Y).$$

Here $\ell_{01}(y_1, y_2) := \mathbf{1}_{[y_1 y_2 \leq 0]}$ is the zero-one loss which is not Lipschitz itself. Finally, we note that $\ell_\gamma \geq \ell_{01}$, and therefore

$$\mathfrak{R}_\gamma(\hat{f}^N; \mu) \geq \mathfrak{R}_{01}(\hat{f}^N; \mu) = \mathbb{P}(\hat{f}^N(X) \neq Y)$$

which implies,

$$\mathbb{P}(\hat{f}^N(X) \neq Y) \leq \frac{1}{N} \sum_{i=1}^N \ell_\gamma(\hat{f}^N(X_i), Y_i) + 2C_{d,1} \sum_{P \in \mathcal{P}} \frac{N_P^{1-1/d}}{N} \text{diam}(P) \frac{\text{Lip}(\hat{f}^N|_P)}{\gamma} + \text{err}$$

with probability greater than $1 - \delta$. ■

A.4 Proof of Proposition 7 on Partitioning

Proof of Proposition 7: Let σ be ReLU activation function and, for every $N \in \mathbb{N}$, consider the feedforward neural network f_N with one layer and one neuron defined by

$$f_N(x) = \frac{\sqrt{N}}{\log_2(\log_2(N))} \cdot \sigma(x - 1 + \frac{1}{N}).$$

We directly compute the following “global quantities” associated to f_N :

$$\|f_N\|_\infty \leq 1 \text{ and } \text{Lip}(f_N) = \frac{\sqrt{N}}{\log_2(\log_2(N))}.$$

Set $L_{f,N} \stackrel{\text{def.}}{=} \sqrt{N}$ and define each \mathcal{F}_N to be the set of Lipschitz functions from $[0, 1]$ to

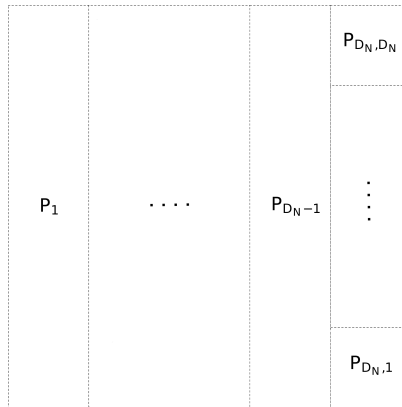


Figure 6: Partition \mathbf{P}_N used in proof of Proposition 7

itself with Lipschitz constant at-most $L_{f,N}$, as in Assumption 2. We build a partition \mathbf{P}_N of $[0, 1] \times [0, 1]$ as follows. Let $\Delta_N = N^{-0.6}$ and $D_N = \lceil \frac{1}{\Delta_N} \rceil = \mathcal{O}(N^{0.6})$.

$$\begin{aligned} \mathbf{P}_N &= \{B_1 \times \mathcal{Y}, \dots, B_{D_N-1} \times \mathcal{Y}, B_{D_N} \times B_1, \dots, B_{D_N} \times B_{D_N}\} \\ &= \{P_1, \dots, P_n, \dots, P_{D_N-1}, P_{D_N,1}, \dots, P_{D_N,n}, \dots, P_{D_N,D_N}\} \end{aligned}$$

where the sets B_1, \dots, B_{D_N} subdivide $[0, 1]$ into D_N intervals as defined by

$$B_n := \begin{cases} [\frac{n-1}{D_N}, \frac{n}{D_N}) & : n = 1, \dots, D_N - 1 \\ [\frac{D_N-1}{D_N}, 1] & : n = D_N \end{cases}$$

Fig. 6 illustrates \mathbf{P}_N . This partition implies the following estimates on local Lipschitz constant of f_N

$$\text{Lip}(f_N|_{P_n}) = 0 \quad \text{and} \quad \text{Lip}(f_N|_{P_{D_N,n}}) = \frac{\sqrt{N}}{\log_2(\log_2(N))}, \quad \forall n \in \mathbb{N}_{<D_N}.$$

Furthermore, we compute the following partition related quantities

$$\max_{P \in \mathbf{P}_N} \text{diam}(P) \leq \sqrt{2} \quad \text{and} \quad |\mathbf{P}_N| = 2D_N - 1 = \mathcal{O}(N^{0.6}).$$

From the above quantities, we compute the ‘‘localized bound’’ of Theorem 4, by calculating the terms $\text{cost}_{\text{partition}}$, $\text{cost}_{\text{transport}}$, $\text{err}_{\text{transport}}$.

$$\begin{aligned} \text{cost}_{\text{transport}} &= \frac{C_{d+1,1}L_\ell}{N} \sum_{P \in \mathbf{P}} N_P^{\frac{d}{d+1}} \max\{1, \text{Lip}(f_N|_P)\} \text{diam}(P) \\ &= C_{2,1} \sum_{P \in \mathbf{P}_N} \frac{(8 + \log_2(N_P))N_P^{1/2}}{N} \text{diam}(P)L_\ell \max\{1, \text{Lip}(f_N|_P)\} \\ &= C_{2,1}L_\ell \sum_{n=1}^N \frac{(8 + \log_2(N_{P_{D_N,n}}))N_{P_{D_N,n}}^{1/2}}{N} \frac{\sqrt{N}}{\log_2(\log_2(N))} \frac{\sqrt{2}}{D_N} \\ &\leq \sqrt{2}C_{2,1}L_\ell \sum_{j=1}^N \frac{(8 + \log_2(1))}{N} \frac{\sqrt{N}}{\log_2(\log_2(N))} \frac{1}{N^{0.6}} \quad (\text{by Jensen's inequality}) \\ &= \frac{8\sqrt{2}C_{2,1}L_\ell}{\log_2(\log_2(N))N^{0.1}}. \end{aligned}$$

For the other two terms,

$$\begin{aligned} \text{err}_{\text{transport}} &\leq \sqrt{\frac{\ln(4/\delta)}{N}} L_\ell \max\{1, L_f\} \max_{P \in \mathbf{P}} \sqrt{\text{diam}(P)^2 + 4B_y^2} \\ &\leq \sqrt{2}L_\ell \frac{\sqrt{N}}{\log_2(\log_2(N))} \frac{\sqrt{\ln(4/\delta)}}{N^{1/2}} \\ \text{cost}_{\text{partition}} &= \frac{B_\ell}{\sqrt{N}} \max\left\{\sqrt{2\ln(4/\delta)}, \sqrt{k_N}\right\} \\ &\leq \frac{\|\ell\|_\infty}{N^{1/2}} \max\left\{\sqrt{2\ln(4/\delta)}, \sqrt{\mathcal{O}(N^{0.6})}\right\} \end{aligned}$$

Therefore, in the $N \rightarrow \infty$ limit,

$$\lim_{N \rightarrow \infty} \text{cost}_{\text{transport}} + \text{err}_{\text{transport}} + \text{cost}_{\text{partition}} = 0$$

and the “local bound” of Theorem 4 converges. In contrast, upon inspecting the “global bound” of Equation (2), we note that it is bounded below by the following quantity

$$\begin{aligned} \text{cost}_{\text{transport}} &= C_{2,1} \frac{(8 + \log_2(N))}{N^{1/2}} \text{diam}(\mathcal{X} \times \mathcal{Y}) L_\ell \max\{1, \text{Lip}(f_N)\} \\ &= \sqrt{2} C_{2,1} L_\ell \frac{(8 + \log_2(N))}{N^{1/2}} \sqrt{N} \\ &= \sqrt{2} C_{2,1} L_\ell (8 + \log_2(N)). \end{aligned} \tag{A.4}$$

We conclude that the “global bound” diverges as N approaches infinity, since the quantity in (A.4) does; i.e. $\lim_{N \rightarrow \infty} \sqrt{2} C_{2,1} L_\ell (8 + \log_2(N)) = \infty$. \blacksquare

A.5 Proof of Corollary 8 on Robustness to Distribution Shifts

Proof of Corollary 8. By Kantorovich Duality (Villani, 2009, Theorem 5.10), we have

$$\mathfrak{R}(\hat{f}^N; \mu^{\text{adv}}) \leq \mathfrak{R}(\hat{f}^N; \mu) + L_\ell \max\left(1, \text{Lip}(\hat{f}^N)\right) \mathcal{W}\left(\mu, \mu^{\text{adv}}\right). \tag{A.5}$$

By Theorem 4, we have for all $0 < \delta \leq 1$ with probability greater than $1 - \delta$,

$$\mathfrak{R}(\hat{f}^N; \mu^{\text{adv}}) - \hat{\mathfrak{R}}(\hat{f}^N) \leq \text{cost}_{\text{transport}}(\mathbf{P}_{\text{global}}) + \text{err}_{\text{transport}}(\mathbf{P}_{\text{global}}). \tag{A.6}$$

Combining (A.5) and (A.6), we complete the proof of Corollary 8. \blacksquare

A.6 Proof of Theorem 10

Proof of Theorem 10. We first consider the non-smooth case. That is, we consider the case where g^N is almost surely α -Hölder with $0 < \alpha \leq 1$.

Step 1 (Change of measure). By Lemma 13, we deduce that

$$\begin{aligned} \int_{z \in \mathcal{Z}} g^N(z) \mu(dz) &\leq \int_{z \in \mathcal{Z}} g^N(z) \nu(dz) + \sum_{P \in \mathcal{P}} \nu(P) \text{Lip}_\alpha(g^N|P) \mathcal{W}_\alpha(\mu|_P, \nu|_P) \\ &\quad + \sum_{P \in \mathcal{P}} \left(1 - \frac{\nu(P)}{\mu(P)}\right) \int_{z \in P} g^N(z) \mu(dz). \end{aligned} \tag{A.7}$$

Setting $\nu = \mu^N$, we obtain an estimation of the expectation of g^N under μ since

$$\begin{aligned} \int_{z \in \mathcal{Z}} g^N(z) \mu(dz) &\leq \int_{z \in \mathcal{Z}} g^N(z) \mu^N(dz) + \sum_{P \in \mathcal{P}} \mu^N(P) \text{Lip}_\alpha(g^N|P) \mathcal{W}_\alpha(\mu|_P, \mu^N|_P) \\ &\quad + \sum_{P \in \mathcal{P}} \left(1 - \frac{\mu^N(P)}{\mu(P)}\right) \int_{z \in P} g^N(z) \mu(dz). \end{aligned}$$

We simplify notations by defining for each $P \in \mathbf{P}$, the following three abbreviations:

$$\begin{aligned}\mathcal{D}_P &\stackrel{\text{def.}}{=} W_\alpha(\mu|_P, \mu^N|_P), \\ \mathcal{B}^N &\stackrel{\text{def.}}{=} \sum_{P \in \mathbf{P}} \mu^N(P) \text{Lip}_\alpha(g^N|P) \mathcal{W}_\alpha(\mu|_P, \mu^N|_P), \\ \mathcal{R}^N &\stackrel{\text{def.}}{=} \sum_{P \in \mathbf{P}} \left(1 - \frac{\mu^N(P)}{\mu(P)}\right) \int_{z \in P} g^N(z) \mu(dz).\end{aligned}$$

With these notational short-hands, we concisely rewrite (A.7) as

$$\int_{z \in \mathcal{Z}} g^N(z) \mu(dz) \leq \int_{z \in \mathcal{Z}} g^N(z) \mu^N(dz) + \mathcal{B}^N + \mathcal{R}^N. \quad (\text{A.8})$$

In order to control our upper-bound in (A.8), we must control the terms \mathcal{B}^N and \mathcal{R}^N ; which we now do.

Step 2 (Integral probability metric concentration). First we control the term \mathcal{B}^N . Notice that

$$\begin{aligned}\mathcal{B}^N &= \sum_{P \in \mathbf{P}} \mu^N(P) \text{Lip}_\alpha(g^N|P) \mathcal{W}_\alpha(\mu|_P, \mu^N|_P) \\ &= \sum_{P \in \mathbf{P}} \mu^N(P) \text{Lip}_\alpha(g^N|P) \mathbb{E}[\mathcal{W}_\alpha(\mu|_P, \mu^N|_P) | N_P] \\ &\quad + \sum_{P \in \mathbf{P}} \mu^N(P) \text{Lip}_\alpha(g^N|P) \left(\mathcal{W}_\alpha(\mu|_P, \mu^N|_P) - \mathbb{E}[\mathcal{W}_\alpha(\mu|_P, \mu^N|_P) | N_P] \right) \\ &\leq \sum_{P \in \mathbf{P}} \mu^N(P) \text{Lip}_\alpha(g^N|P) \mathbb{E}[\mathcal{W}_\alpha(\mu|_P, \mu^N|_P) | N_P] \\ &\quad + L \sum_{P \in \mathbf{P}} \mu^N(P) \left| \mathcal{W}_\alpha(\mu|_P, \mu^N|_P) - \mathbb{E}[\mathcal{W}_\alpha(\mu|_P, \mu^N|_P) | N_P] \right|.\end{aligned}$$

Deploying Lemma 16, we find that

$$\begin{aligned}&\sum_{P \in \mathbf{P}} \mu^N(P) \text{Lip}_\alpha(g^N|P) \mathbb{E}[\mathcal{W}_\alpha(\mu|_P, \mu^N|_P) | N_P] \\ &\leq \sum_{P \in \mathbf{P}} \mu^N(P) \text{Lip}_\alpha(g^N|P) C_{d_{\mathcal{Z}}, \alpha} \text{diam}(P) \text{rate}_{d_{\mathcal{Z}}, \alpha}(N_P).\end{aligned} \quad (\text{A.9})$$

and that

$$\mathbb{P}\left(\left| \mathcal{W}_\alpha(\mu|_P, \mu^N|_P) - \mathbb{E}[\mathcal{W}_\alpha(\mu|_P, \mu^N|_P) | N_P] \right| \geq \epsilon \mid N_P\right) \leq 2e^{-\frac{2N_P \epsilon^2}{\text{diam}(P)^{2\alpha}}}.$$

Synchronizing our notation with that of Lemma 22 we set $C_P = 2$, $\sigma_P^2 = \text{diam}(P)^{2\alpha}/4N_P$, $\alpha_P = L\mu^N(P)$ and

$$X_P = \left| \mathcal{W}_\alpha(\mu|_P, \mu^N|_P) - \mathbb{E}[\mathcal{W}_\alpha(\mu|_P, \mu^N|_P) | N_P] \right|.$$

for $P \in \mathbf{P}$. Apply Lemma 22 while conditioning on N_P we have that for every $\epsilon > 0$ and each $N \in \mathbb{N}$

$$\mathbb{P}\left[\left|\sum_{P \in \mathbf{P}} \alpha_P X_P\right| \geq \epsilon \mid N_P\right] \leq 2e^{-\frac{\epsilon^2}{8\tilde{\sigma}^2}},$$

where $\tilde{\sigma}^2$ is can be bounded above as follows

$$\begin{aligned} \tilde{\sigma}^2 &= \sum_{P \in \mathbf{P}} C_P^2 \alpha_P^2 \sigma_P^2 \\ &= \sum_{P \in \mathbf{P}} 4\mu^N(P)^2 L^2 \frac{\text{diam}(P)^{2\alpha}}{4N_P} \\ &= \sum_{P \in \mathbf{P}} \frac{N_P}{N^2} L^2 \text{diam}(P)^{2\alpha} \\ &\leq \frac{L^2}{N} \max_{P \in \mathbf{P}} \text{diam}(P)^{2\alpha}. \end{aligned}$$

Therefore, we deduce the following concentration inequality

$$\begin{aligned} \mathbb{P}\left(\left|\sum_{P \in \mathbf{P}} \alpha_P X_P\right| \geq \epsilon\right) &= \mathbb{E}\left[\mathbb{P}\left(\left|\sum_{P \in \mathbf{P}} \alpha_P X_P\right| \geq \epsilon \mid N_P\right)\right] \\ &\leq \mathbb{E}\left[2 \exp\left\{-\frac{N\epsilon^2}{L^2 \max_{P \in \mathbf{P}} \text{diam}(P)^{2\alpha}}\right\}\right] \\ &= 2 \exp\left\{-\frac{N\epsilon^2}{L^2 \max_{P \in \mathbf{P}} \text{diam}(P)^{2\alpha}}\right\} \end{aligned}$$

Fix our ‘‘threshold probability’’ $0 < \delta_1 \leq 1$. With probability $1 - \delta_1$ it holds that

$$L \sum_{P \in \mathbf{P}} \mu^N(P) \left| \mathcal{W}_\alpha(\mu|_P, \mu^N|_P) - \mathbb{E}[\mathcal{W}_\alpha(\mu|_P, \mu^N|_P) \mid N_P] \right| \leq L \max_{P \in \mathbf{P}} \text{diam}(P)^\alpha \left(\frac{\ln(2/\delta_1)}{N}\right)^{1/2}.$$

Combining it with (A.9), we conclude that the following holds with probability $1 - \delta_1$

$$\begin{aligned} \mathcal{B}^N &\leq \sum_{P \in \mathbf{P}} \mu^N(P) \text{Lip}_\alpha(g^N|_P) C_{d_Z, \alpha} \text{diam}(P) \text{rate}_{d_Z, \alpha}(N_P) \\ &\quad + L \max_{P \in \mathbf{P}} \text{diam}(P)^\alpha \left(\frac{\ln(2/\delta_1)}{N}\right)^{1/2}. \end{aligned} \tag{A.10}$$

Step 3: (Global concentration). It remains to estimate the term \mathcal{R}^N . Let $\delta_2 > 0$, by Lemma 24, we know that the following holds with probability $1 - \delta_2$

$$\mathcal{R}_N \leq \frac{\|g^N\|_\infty}{N^{1/2}} \max\{\sqrt{2 \ln(2/\delta_2)}, \sqrt{k}\}. \tag{A.11}$$

Combining (A.8), (A.10) and (A.11), we have with probability greater than $(1 - \delta_1)(1 - \delta_2)$

$$\mathbb{E}[g^N(Z)] - \frac{1}{N} \sum_{n=1}^N g^N(Z_n) \leq \sum_{P \in \mathbf{P}} \mu^N(P) \text{Lip}_\alpha(g^N|_P) C_{d_Z, \alpha} \text{diam}(P) \text{rate}_{d_Z, \alpha}(N_P) + \epsilon,$$

where the term ϵ is given by

$$\epsilon \stackrel{\text{def.}}{=} L \max_{P \in \mathbf{P}} \text{diam}(P)^\alpha \left(\frac{\ln(2/\delta_1)}{N} \right)^{1/2} + \frac{\|g^N\|_\infty}{N^{1/2}} \max\{\sqrt{2 \ln(2/\delta_2)}, \sqrt{k}\}.$$

Let $\delta \in (0, 1]$. Set $\delta_1 \stackrel{\text{def.}}{=} \delta_2 \stackrel{\text{def.}}{=} \delta/2$. We now have with probability greater than $1 - \delta$ that

$$\epsilon \stackrel{\text{def.}}{=} L \max_{P \in \mathbf{P}} \text{diam}(P) \left(\frac{\ln(4/\delta)}{N} \right)^{1/2} + \frac{\|g^N\|_\infty}{N^{1/2}} \max\{\sqrt{2 \ln(4/\delta)}, \sqrt{k}\}.$$

We now turn our attention to the proof of the smooth case; which is similar modulo some a few changes at key points in its proof.

Step 1 (Change of measure). By Applying Lemma 15 we have

$$\begin{aligned} \int_{z \in \mathcal{Z}} g^N(z) \mu(dz) &\leq \int_{z \in \mathcal{Z}} g^N(z) \nu(dz) + \sum_{P \in \mathbf{P}} \nu(P) \|g^N\|_{s:P} \mathcal{W}_{\mathcal{C}^s}(\mu|_P, \nu|_P) \\ &\quad + \sum_{P \in \mathbf{P}} \left(1 - \frac{\nu(P)}{\mu(P)}\right) \int_{z \in P} g^N(z) \mu(dz). \end{aligned}$$

Setting $\nu \stackrel{\text{def.}}{=} \mu^N$ in the above equation, we estimate the mean of g^N with respect to μ

$$\begin{aligned} \int_{z \in \mathcal{Z}} g^N(z) \mu(dz) &\leq \int_{z \in \mathcal{Z}} g^N(z) \mu^N(dz) + \sum_{P \in \mathbf{P}} \mu^N(P) \|g^N\|_{s:P} \mathcal{W}_{\mathcal{C}^s}(\mu|_P, \mu^N|_P) \\ &\quad + \sum_{P \in \mathbf{P}} \left(1 - \frac{\mu^N(P)}{\mu(P)}\right) \int_{z \in P} g^N(z) \mu(dz). \end{aligned} \tag{A.12}$$

As before, we simplify our notation. For each for all $P \in \mathbf{P}$ we abbreviate

$$\begin{aligned} \mathcal{D}_P &\stackrel{\text{def.}}{=} \mathcal{W}_{\mathcal{C}^s}(\mu|_P, \mu^N|_P) \\ \mathcal{B}^N &\stackrel{\text{def.}}{=} \sum_{P \in \mathbf{P}} \mu^N(P) \|g^N\|_{s:P} \mathcal{W}_{\mathcal{C}^s}(\mu|_P, \mu^N|_P) \\ \mathcal{R}^N &\stackrel{\text{def.}}{=} \sum_{P \in \mathbf{P}} \left(1 - \frac{\mu^N(P)}{\mu(P)}\right) \int_{z \in P} g^N(z) \mu(dz). \end{aligned}$$

Therefore, (A.12) can be succinctly written as

$$\int_{z \in \mathcal{Z}} g^N(z) \mu(dz) \leq \int_{z \in \mathcal{Z}} g^N(z) \mu^N(dz) + \mathcal{B}^N + \mathcal{R}^N. \tag{A.13}$$

As in the non-smooth case, we need only bound the terms \mathcal{B}^N and \mathcal{R}^N in order to control the left-hand side of (A.13).

Step 2 (Integral probability metric concentration). We again first control the term \mathcal{B}^N .

Observe that

$$\begin{aligned}
 \mathcal{B}^N &= \sum_{P \in \mathcal{P}} \mu^N(P) \|g^N\|_{s:P} \mathcal{W}_{\mathcal{C}^s}(\mu|_P, \mu^N|_P) \\
 &= \sum_{P \in \mathcal{P}} \mu^N(P) \|g^N\|_{s:P} \mathbb{E}[\mathcal{W}_{\mathcal{C}^s}(\mu|_P, \mu^N|_P) | N_P] \\
 &\quad + \sum_{P \in \mathcal{P}} \mu^N(P) \|g^N\|_{s:P} \left(\mathcal{W}_{\mathcal{C}^s}(\mu|_P, \mu^N|_P) - \mathbb{E}[\mathcal{W}_{\mathcal{C}^s}(\mu|_P, \mu^N|_P) | N_P] \right) \\
 &\leq \sum_{P \in \mathcal{P}} \mu^N(P) \|g^N\|_{s:P} \mathbb{E}[\mathcal{W}_{\mathcal{C}^s}(\mu|_P, \mu^N|_P) | N_P] \\
 &\quad + L \sum_{P \in \mathcal{P}} \mu^N(P) \left| \mathcal{W}_{\mathcal{C}^s}(\mu|_P, \mu^N|_P) - \mathbb{E}[\mathcal{W}_{\mathcal{C}^s}(\mu|_P, \mu^N|_P) | N_P] \right|
 \end{aligned}$$

Applying Lemma 17 we have both that

$$\begin{aligned}
 &\sum_{P \in \mathcal{P}} \mu^N(P) \|g^N\|_{s:P} \mathbb{E}[\mathcal{W}_{\mathcal{C}^s}(\mu|_P, \mu^N|_P) | N_P] \\
 &\leq \sum_{P \in \mathcal{P}} \mu^N(P) \|g^N\|_{s:P} C_{d_{\mathcal{Z}},s} \text{diam}(P) \text{rate}_{d_{\mathcal{Z}},s}(N_P).
 \end{aligned} \tag{A.14}$$

and that

$$\mathbb{P}\left(\left| \mathcal{W}_{\mathcal{C}^s}(\mu|_P, \mu^N|_P) - \mathbb{E}[\mathcal{W}_{\mathcal{C}^s}(\mu|_P, \mu^N|_P) | N_P] \right| \geq \epsilon \mid N_P\right) \leq 2e^{-\frac{2N_P \epsilon^2}{\text{diam}(P)^2}}.$$

Synchronizing notation with Lemma 22 we denote $C_P = 1$, $\sigma_P^2 = \text{diam}(P)^2/4N_P$ and $\alpha_P = L\mu^N(P)$,

$$X_P = \left| \mathcal{W}_{\mathcal{C}^s}(\mu|_P, \mu^N|_P) - \mathbb{E}[\mathcal{W}_{\mathcal{C}^s}(\mu|_P, \mu^N|_P) | N_P] \right|$$

for $P \in \mathcal{P}$. Applying Lemma 22 while conditioning on N_P , we have for all $\epsilon > 0$ and $N \in \mathbb{N}$,

$$\mathbb{P}\left[\left| \sum_{P \in \mathcal{P}} \alpha_P X_P \right| \geq \epsilon \mid N_P\right] \leq 2e^{-\frac{\epsilon^2}{8\tilde{\sigma}^2}},$$

where, similarly to the smooth case, $\tilde{\sigma}^2$ is bounded above by

$$\begin{aligned}
 \tilde{\sigma}^2 &= \sum_{P \in \mathcal{P}} C_P^2 \alpha_P^2 \sigma_P^2 \\
 &= \sum_{P \in \mathcal{P}} 4\mu^N(P)^2 L^2 \frac{\text{diam}(P)^2}{4N_P} \\
 &= \sum_{P \in \mathcal{P}} \frac{N_P}{N^2} L^2 \text{diam}(P)^2 \\
 &\leq \frac{L^2}{N} \max_{P \in \mathcal{P}} \text{diam}(P)^2.
 \end{aligned}$$

Therefore, we arrive at the fact that

$$\begin{aligned} \mathbb{P}\left(\left|\sum_{P \in \mathcal{P}} \alpha_P X_P\right| \geq \epsilon\right) &= \mathbb{E}\left[\mathbb{P}\left(\left|\sum_{P \in \mathcal{P}} \alpha_P X_P\right| \geq \epsilon \mid N_P\right)\right] \\ &\leq \mathbb{E}\left[2 \exp\left\{-\frac{N\epsilon^2}{L^2 \max_{P \in \mathcal{P}} \text{diam}(P)^2}\right\}\right] \\ &= 2 \exp\left\{-\frac{N\epsilon^2}{L^2 \max_{P \in \mathcal{P}} \text{diam}(P)^2}\right\}. \end{aligned}$$

Let $0 < \delta_1 \leq 1$ be our ‘‘threshold probability’’. Thus, with probability $1 - \delta_1$ it holds that

$$L \sum_{P \in \mathcal{P}} \mu^N(P) \left| \mathcal{W}_{\mathcal{C}^s}(\mu|_P, \mu^N|_P) - \mathbb{E}[\mathcal{W}_{\mathcal{C}^s}(\mu|_P, \mu^N|_P) \mid N_P] \right| \leq L \max_{P \in \mathcal{P}} \text{diam}(P) \left(\frac{\ln(2/\delta_1)}{N}\right)^{1/2}.$$

Combine this with (A.14) we conclude that with probability $1 - \delta_1$

$$\mathcal{B}^N \leq \sum_{P \in \mathcal{P}} \mu^N(P) \|g^N\|_{s:P} C_{d_{\mathcal{Z}},s} \text{diam}(P) \text{rate}_{d_{\mathcal{Z}},s}(N_P) + L \max_{P \in \mathcal{P}} \text{diam}(P) \left(\frac{\ln(2/\delta_1)}{N}\right)^{1/2}. \quad (\text{A.15})$$

Step 3: (Global concentration). As with the smooth case, it only now remains to control the term \mathcal{R}^N . Set $0 < \delta_2 \leq 1$. Then, by Lemma 24, we have that the following holds with probability at-least $1 - \delta_2$

$$\mathcal{R}_N \leq \frac{\|g^N\|_\infty}{N^{1/2}} \max\{\sqrt{2 \ln(2/\delta_2)}, \sqrt{k}\}. \quad (\text{A.16})$$

Combining (A.13), (A.15) and (A.16), we have with probability greater than $(1 - \delta_1)(1 - \delta_2)$

$$\mathbb{E}[g^N(Z)] - \frac{1}{N} \sum_{n=1}^N g^N(Z_n) \leq \sum_{P \in \mathcal{P}} \mu^N(P) \|g^N\|_{s:P} C_{d_{\mathcal{Z}},s} \text{diam}(P) \text{rate}_{d_{\mathcal{Z}},s}(N_P) + \epsilon,$$

where the ‘‘error term’’ ϵ is given by

$$\epsilon \stackrel{\text{def.}}{=} L \max_{P \in \mathcal{P}} \text{diam}(P) \left(\frac{\ln(2/\delta_1)}{N}\right)^{1/2} + \frac{\|g^N\|_\infty}{N^{1/2}} \max\{\sqrt{2 \ln(2/\delta_2)}, \sqrt{k}\}.$$

Fix $\delta \in (0, 1]$. Set $\delta_1 \stackrel{\text{def.}}{=} \delta_2 \stackrel{\text{def.}}{=} \delta/2$. Thus, from our analysis we arrive at the conclusion that

$$\epsilon = L \max_{P \in \mathcal{P}} \text{diam}(P) \left(\frac{\ln(4/\delta)}{N}\right)^{1/2} + \frac{\|g^N\|_\infty}{N^{1/2}} \max\{\sqrt{2 \ln(4/\delta)}, \sqrt{k}\},$$

holds with probability at-least $1 - \delta$. This concludes our proof. \blacksquare

Appendix B. Helper Lemmas

B.1 Change of Measure Helper Inequalities

Lemma 12 (Change of Measure of Hölder Function) *Let \mathcal{Z} be a complete and separable metric space and $\mu, \nu \in \mathcal{P}_1(\mathcal{Z})$. Let $g : \mathcal{Z} \rightarrow \mathbb{R}$ be an α -Hölder function with $\alpha \in (0, 1]$. Then*

$$\int_{z \in \mathcal{Z}} g(z) \mu(dz) \leq \text{Lip}_\alpha(g) \mathcal{W}_\alpha(\mu, \nu) + \int_{z \in \mathcal{Z}} g(z) \nu(dz).$$

Proof of Lemma 12. (Villani, 2009, Theorem 5.10) By the definition of $\mathcal{W}_\alpha(\mu, \nu)$ as an integral probability distance,

$$\mathcal{W}_\alpha(\mu, \nu) = \sup_{f \in \mathcal{C}(\mathcal{Z}), \text{Lip}_\alpha(f) \leq 1} \int f(z) \mu(dz) - \int f(z) \nu(dz).$$

If g is constant, then Lemma 12 holds trivially. Otherwise, $0 < \text{Lip}_\alpha(g) < \infty$ and we let $\tilde{g} \stackrel{\text{def.}}{=} \text{Lip}_\alpha(g)^{-1} g$. Then $\text{Lip}_\alpha(\tilde{g}) \leq 1$ and we have

$$\begin{aligned} \text{Lip}_\alpha(g)^{-1} \int g \mu(dz) &= \int \tilde{g}(z) \mu(dz) \\ &\leq \mathcal{W}_\alpha(\mu, \nu) + \int \tilde{g}(z) \nu(dz) \\ &= \mathcal{W}_\alpha(\mu, \nu) + \text{Lip}_\alpha(g)^{-1} \int g(z) \nu(dz). \end{aligned}$$

Multiplying across by $\text{Lip}_\alpha(g) > 0$ yields the desired result. ■

Lemma 13 (Local Change of Measure of Hölder Function) *Let \mathcal{Z} a subset of \mathbb{R}^{d_Z} and $\mu, \nu \in \mathcal{P}_1(\mathcal{Z})$. Let $g : \mathcal{Z} \rightarrow \mathbb{R}$ be locally α -Hölder with $\alpha \in (0, 1]$. Then*

$$\begin{aligned} \int_{z \in \mathcal{Z}} g(z) \mu(dz) &\leq \int_{z \in \mathcal{Z}} g(z) \nu(dz) + \sum_{P \in \mathbf{P}} \nu(P) \text{Lip}_\alpha(g|P) \mathcal{W}_\alpha(\mu|P, \nu|P) \\ &\quad + \sum_{P \in \mathbf{P}} \left(1 - \frac{\nu(P)}{\mu(P)}\right) \int_{z \in P} g(z) \mu(dz). \end{aligned}$$

Proof of Lemma 13. Let for all $P \in \mathbf{P}$ that

$$\mu_P(\cdot) = \mu(\cdot \cap P), \quad \nu_P(\cdot) = \nu(\cdot \cap P), \quad \tilde{\mu}_P = \frac{\nu(P)}{\mu(P)} \mu_P.$$

Then we apply Lemma 12 to $\tilde{\mu}_P$ and ν_P for all $P \in \mathbf{P}$ and have

$$\begin{aligned} \int_{z \in \mathcal{Z}} -g(z) \nu(dz) &= \sum_{P \in \mathbf{P}} \int_{z \in \mathcal{Z}} -g(z) \nu_P(dz) \\ &\leq \sum_{P \in \mathbf{P}} \left(\text{Lip}_\alpha(g|P) \mathcal{W}_\alpha(\nu_P, \tilde{\mu}_P) + \int_{z \in \mathcal{Z}} -g(z) \tilde{\mu}_P(dz) \right). \end{aligned}$$

By adding $\int_{z \in \mathcal{Z}} g(z) \mu(dz)$ on both sides and rearranging terms we have that

$$\begin{aligned}
 \int_{z \in \mathcal{Z}} g(z) \mu(dz) &\leq \int_{z \in \mathcal{Z}} g(z) \nu(dz) + \sum_{P \in \mathcal{P}} \text{Lip}_\alpha(g|P) \mathcal{W}_\alpha(\nu_P, \tilde{\mu}_P) \\
 &\quad + \int_{z \in \mathcal{Z}} g(z) \mu_P(dz) - \int_{z \in \mathcal{Z}} g(z) \frac{\nu(P)}{\mu(P)} \mu_P(dz) \\
 &\leq \int_{z \in \mathcal{Z}} g(z) \nu(dz) + \sum_{P \in \mathcal{P}} \text{Lip}_\alpha(g|P) \nu(P) \mathcal{W}_\alpha\left(\frac{\nu_P}{\nu(P)}, \frac{\mu_P}{\mu(P)}\right) \\
 &\quad + \int_{z \in \mathcal{Z}} g(z) \mu_P(dz) - \int_{z \in \mathcal{Z}} g(z) \frac{\nu(P)}{\mu(P)} \mu_P(dz) \\
 &\leq \int_{z \in \mathcal{Z}} g(z) \nu(dz) + \sum_{P \in \mathcal{P}} \text{Lip}_\alpha(g|P) \nu(P) \mathcal{W}_\alpha(\nu|_P, \mu|_P) \\
 &\quad + \left(1 - \frac{\nu(P)}{\mu(P)}\right) \int_{z \in \mathcal{Z}} g(z) \mu_P(dz) \\
 &\leq \int_{z \in \mathcal{Z}} g(z) \mu^N(dz) + \sum_{P \in \mathcal{P}} \text{Lip}_\alpha(g|P) \mu^N(P) \mathcal{W}_\alpha\left((\mu|_P)^{N_P}, \mu|_P\right) \\
 &\quad + \left(1 - \frac{\mu^N(P)}{\mu(P)}\right) \int_{z \in \mathcal{Z}} g(z) \mu_P(dz).
 \end{aligned}$$

■

Lemma 14 (Change of Measure of Smooth Function) *Let $\mathcal{Z} = \mathbb{R}^{d_{\mathcal{Z}}}$, $d_{\mathcal{Z}} \in \mathbb{N}$ and $\mu, \nu \in \mathcal{P}_1(\mathcal{Z})$. Let $g \in \mathcal{C}^s(\mathcal{Z})$ with $s \in \mathbb{N}$. Then*

$$\int_{z \in \mathcal{Z}} g(z) \mu(dz) \leq \|g\|_s \mathcal{W}_{\mathcal{C}^s}(\mu, \nu) + \int_{z \in \mathcal{Z}} g(z) \nu(dz).$$

Proof of Lemma 14. The proof is similar with the proof of Lemma 12. Recall the definition of $\mathcal{W}_{\mathcal{C}^s}$ that

$$\mathcal{W}_{\mathcal{C}^s}(\mu, \nu) \stackrel{\text{def.}}{=} \sup_{f \in \mathcal{C}(\mathcal{Z}), \|f\|_s \leq 1} \int f(z) \mu(dz) - \int f(z) \nu(dz).$$

If g is constant, then Lemma 14 holds trivially. Otherwise, $0 < \|g\|_s < \infty$ and we Let $\tilde{g} \stackrel{\text{def.}}{=} \|g\|_s^{-1} g$. Then $\|\tilde{g}\|_s \leq 1$ and we have

$$\begin{aligned}
 \|g\|_s^{-1} \int g \mu(dz) &= \int \tilde{g}(z) \mu(dz) \\
 &\leq \mathcal{W}_\alpha(\mu, \nu) + \int \tilde{g}(z) \nu(dz) \\
 &= \mathcal{W}_\alpha(\mu, \nu) + \|g\|_s^{-1} \int g(z) \nu(dz).
 \end{aligned}$$

Multiplying across by $\|g\|_s > 0$ yields the desired result. ■

Lemma 15 (Local Change of Measure of Smooth Function) *Let $\mathcal{Z} = \mathbb{R}^{dz}$, $dz \in \mathbf{N}$ and $\mu, \nu \in \mathcal{P}_1(\mathcal{Z})$. Let $g \in \mathcal{C}^s(\mathcal{Z})$ with $s \in \mathbf{N}$. Then*

$$\begin{aligned} \int_{z \in \mathcal{Z}} g(z) \mu(dz) &\leq \int_{z \in \mathcal{Z}} g(z) \nu(dz) + \sum_{P \in \mathbf{P}} \mu(P) \|g\|_{s:P} \mathcal{W}_{\mathcal{C}^s}(\mu|_P, \nu|_P) \\ &\quad + \sum_{P \in \mathbf{P}} \left(1 - \frac{\nu(P)}{\mu(P)}\right) \int_{z \in P} g(z) \mu(dz). \end{aligned}$$

Proof of Lemma 15. The proof is similar with the proof of Lemma 13. Let for all $P \in \mathbf{P}$ that

$$\mu_P(\cdot) = \mu(\cdot \cap P), \quad \nu_P(\cdot) = \nu(\cdot \cap P), \quad \tilde{\mu}_P = \frac{\nu(P)}{\mu(P)} \mu_P.$$

Then we apply Lemma 12 to $\tilde{\mu}_P$ and ν_P for all $P \in \mathbf{P}$ and have

$$\begin{aligned} \int_{z \in \mathcal{Z}} -g(z) \nu(dz) &= \sum_{P \in \mathbf{P}} \int_{z \in \mathcal{Z}} -g(z) \nu_P(dz) \\ &\leq \sum_{P \in \mathbf{P}} \left(\|g\|_{s:P} \mathcal{W}_{\mathcal{C}^s}(\nu_P, \tilde{\mu}_P) + \int_{z \in \mathcal{Z}} -g(z) \tilde{\mu}_P(dz) \right). \end{aligned}$$

By adding $\int_{z \in \mathcal{Z}} g(z) \mu(dz)$ on both sides and rearranging terms we have that

$$\begin{aligned} \int_{z \in \mathcal{Z}} g(z) \mu(dz) &\leq \int_{z \in \mathcal{Z}} g(z) \nu(dz) + \sum_{P \in \mathbf{P}} \|g\|_{s:P} \mathcal{W}_{\mathcal{C}^s}(\nu_P, \tilde{\mu}_P) \\ &\quad + \int_{z \in \mathcal{Z}} g(z) \mu_P(dz) - \int_{z \in \mathcal{Z}} g(z) \frac{\nu(P)}{\mu(P)} \mu_P(dz) \\ &\leq \int_{z \in \mathcal{Z}} g(z) \nu(dz) + \sum_{P \in \mathbf{P}} \nu(P) \|g\|_{s:P} \mathcal{W}_{\mathcal{C}^s}\left(\frac{\nu_P}{\nu(P)}, \frac{\mu_P}{\mu(P)}\right) \\ &\quad + \int_{z \in \mathcal{Z}} g(z) \mu_P(dz) - \int_{z \in \mathcal{Z}} g(z) \frac{\nu(P)}{\mu(P)} \mu_P(dz) \\ &\leq \int_{z \in \mathcal{Z}} g(z) \nu(dz) + \sum_{P \in \mathbf{P}} \nu(P) \|g\|_{s:P} \mathcal{W}_{\mathcal{C}^s}(\nu|_P, \mu|_P) \\ &\quad + \left(1 - \frac{\nu(P)}{\mu(P)}\right) \int_{z \in \mathcal{Z}} g(z) \mu_P(dz) \\ &\leq \int_{z \in \mathcal{Z}} g(z) \mu^N(dz) + \sum_{P \in \mathbf{P}} \mu^N(P) \|g\|_{s:P} \mathcal{W}_{\mathcal{C}^s}\left((\mu|_P)^{N_P}, \mu|_P\right) \\ &\quad + \left(1 - \frac{\mu^N(P)}{\mu(P)}\right) \int_{z \in \mathcal{Z}} g(z) \mu_P(dz). \end{aligned}$$

■

B.2 Helper Wasserstein Concentration Inequalities

Lemma 16 (Concentration of Hölder Wasserstein Metric) *Let \mathcal{Z} a compact subset of $\mathbb{R}^{d_{\mathcal{Z}}}$ and $\mu \in \mathcal{P}_1(\mathcal{Z})$. Then for all $\epsilon > 0$, $N \in \mathbb{N}$*

$$\mathbb{P}\left(\left|\mathcal{W}_\alpha(\mu, \mu^N) - \mathbb{E}\left[\mathcal{W}_\alpha(\mu, \mu^N)\right]\right| \geq \epsilon\right) \leq 2e^{-\frac{2N\epsilon^2}{\text{diam}(\mathcal{Z})^{2\alpha}}}$$

where $C_{d_{\mathcal{Z}}, \alpha}$ is given in Table 2 and

$$\mathbb{E}\left[\mathcal{W}_\alpha(\mu, \mu^N)\right] \leq C_{d_{\mathcal{Z}}, \alpha} \text{diam}(\mathcal{Z}) \text{rate}_{d_{\mathcal{Z}}, \alpha}(N)$$

with $\text{rate}_{d_{\mathcal{Z}}, \alpha}(N)$ also given in Table 2.

Proof of Lemma 16. In the proof of Lemma 16, we consider two different norms on the cube $[0, 1]^{d_{\mathcal{Z}}}$ in order to apply (Kloeckner, 2020, Theorem 2.1). The first is the *Euclidean norm* $\|u\|_2^2 := \sum_{i=1}^{d_{\mathcal{Z}}} u_i^2$ and the second is the ∞ -norm defined by $\|u\|_\infty := \max_{i=1, \dots, d_{\mathcal{Z}}} |u_i|$. When needed from the context, we emphasize implicitly used when defining the Wasserstein distance by $\mathcal{W}_{\alpha:2}$ and $\mathcal{W}_{\alpha:\infty}$ for the Euclidean and ∞ norms, respectively. By (Kloeckner, 2020, Theorem 2.1), we have for $\mathcal{Z} = [0, 1]^{d_{\mathcal{Z}}}$ and $N \in \mathbb{N}$

$$\mathbb{E}\left[\mathcal{W}_{\alpha:\infty}(\mu, \mu^N)\right] \leq d_{\mathcal{Z}}^{-\alpha/2} C_{d_{\mathcal{Z}}, \alpha} \text{rate}_{d_{\mathcal{Z}}, \alpha}(N).$$

By a simple fact that $\mathcal{W}_{\alpha:2} \leq d_{\mathcal{Z}}^{\alpha/2} \mathcal{W}_{\alpha:\infty}$, we have

$$\mathbb{E}\left[\mathcal{W}_\alpha(\mu, \mu^N)\right] = \mathbb{E}\left[\mathcal{W}_{\alpha:2}(\mu, \mu^N)\right] \leq C_{d_{\mathcal{Z}}, \alpha} \text{rate}_{d_{\mathcal{Z}}, \alpha}(N).$$

We scale $[0, 1]^{d_{\mathcal{Z}}}$ with $\text{diam}(\mathcal{Z})$ to conclude that for general $\mathcal{Z} \subset \mathbb{R}^{d_{\mathcal{Z}}}$

$$\mathbb{E}\left[\mathcal{W}_\alpha(\mu, \mu^N)\right] \leq C_{d_{\mathcal{Z}}, \alpha} \text{diam}(\mathcal{Z}) \text{rate}_{d_{\mathcal{Z}}, \alpha}(N).$$

Now we define $f: \mathcal{Z}^N \rightarrow \mathbb{R}$ s.t.

$$f_N(z_1, \dots, z_N) \stackrel{\text{def.}}{=} \mathcal{W}_\alpha\left(\frac{1}{N} \sum_{n=1}^N \delta_{z_n}, \mu\right).$$

For every $i = 1, \dots, N$ and every $(z_1, \dots, z_N), (z'_1, \dots, z'_N) \in \mathcal{Z}^N$ that differs only in the i -th coordinate, we have

$$|f(z_1, \dots, z_N) - f(z'_1, \dots, z'_N)| \leq \mathcal{W}_\alpha\left(\frac{1}{N} \sum_{n=1}^N \delta_{z_n}, \frac{1}{N} \sum_{n=1}^N \delta_{z'_n}\right) \leq \frac{\text{diam}(\mathcal{Z})^\alpha}{N}.$$

Therefore, with $c = \frac{\text{diam}(\mathcal{Z})^\alpha}{N}$, f has (c, \dots, c) -bounded differences property i.e. Lipschitz w.r.t. Hamming distance. Applying Lemma 23 (the McDiarmid's inequality) with f proves that for all $\epsilon > 0$

$$\mathbb{P}\left(\left|\mathcal{W}_\alpha(\mu, \mu^N) - \mathbb{E}\left[\mathcal{W}_\alpha(\mu, \mu^N)\right]\right| \geq \epsilon\right) \leq 2e^{-\frac{2N\epsilon^2}{\text{diam}(\mathcal{Z})^{2\alpha}}}.$$

■

Lemma 17 (Concentration of Smooth Wasserstein Metric) *Let $\mathcal{Z} = \mathbb{R}^{d_{\mathcal{Z}}}$, $d_{\mathcal{Z}} \in \mathbb{N}$ and $\mu, \nu \in \mathcal{P}_1(\mathcal{Z})$. Let $g \in \mathcal{C}^s(\mathcal{Z})$ with $s \in \mathbb{N}$. Then there exist constant $C_{d_{\mathcal{Z}},s} > 0$ s.t. for all $\epsilon > 0$, $N \in \mathbb{N}$*

$$\mathbb{P}\left(\left|\mathcal{W}_{\mathcal{C}^s}(\mu, \mu^N) - \mathbb{E}\left[\mathcal{W}_{\mathcal{C}^s}(\mu, \mu^N)\right]\right| \geq \epsilon\right) \leq 2e^{-\frac{2N\epsilon^2}{\text{diam}(\mathcal{Z})^2}},$$

and

$$\mathbb{E}\left[\mathcal{W}_{\mathcal{C}^s}(\mu, \mu^N)\right] \leq C_{d_{\mathcal{Z}},s} \text{diam}(\mathcal{Z}) \text{rate}_{d_{\mathcal{Z}},s}(N).$$

Proof of Lemma 17. The proof is similar to the proof of Lemma 16. By (Kloeckner, 2020, Theorem 1.4), we have for $\mathcal{Z} = [0, 1]^{d_{\mathcal{Z}}}$ and $N \in \mathbb{N}$

$$\mathbb{E}\left[\mathcal{W}_{\mathcal{C}^s}(\mu, \mu^N)\right] \leq C_{d_{\mathcal{Z}},s} \text{rate}_{d_{\mathcal{Z}},s}(N).$$

Next, we scale $[0, 1]^{d_{\mathcal{Z}}}$ with $\text{diam}(\mathcal{Z})$ to conclude that for general $\mathcal{Z} \subset \mathbb{R}^{d_{\mathcal{Z}}}$

$$\mathbb{E}\left[\mathcal{W}_{\mathcal{C}^s}(\mu, \mu^N)\right] \leq C_{d_{\mathcal{Z}},s} \text{diam}(\mathcal{Z}) \text{rate}_{d_{\mathcal{Z}},s}(N).$$

Now we define $f: \mathcal{Z}^N \rightarrow \mathbb{R}$ s.t.

$$f_N(z_1, \dots, z_N) \stackrel{\text{def.}}{=} \mathcal{W}_{\mathcal{C}^s}\left(\frac{1}{N} \sum_{n=1}^N \delta_{z_n}, \mu\right).$$

For every $i = 1, \dots, N$ and every $(z_1, \dots, z_N), (z'_1, \dots, z'_N) \in \mathcal{Z}^N$ that differs only in the i -th coordinate, we have

$$|f(z_1, \dots, z_N) - f(z'_1, \dots, z'_N)| \leq \mathcal{W}_{\mathcal{C}^s}\left(\frac{1}{N} \sum_{n=1}^N \delta_{z_n}, \frac{1}{N} \sum_{n=1}^N \delta_{z'_n}\right) \leq \frac{\text{diam}(\mathcal{Z})}{N}.$$

Therefore, with $c = \frac{\text{diam}(\mathcal{Z})}{N}$, f has (c, \dots, c) -bounded differences property i.e. Lipschitz w.r.t. Hamming distance. Applying Lemma 23 (the McDiarmid's inequality) with f proves that for all $\epsilon > 0$

$$\mathbb{P}\left(\left|\mathcal{W}_{\mathcal{C}^s}(\mu, \mu^N) - \mathbb{E}\left[\mathcal{W}_{\mathcal{C}^s}(\mu, \mu^N)\right]\right| \geq \epsilon\right) \leq 2e^{-\frac{2N\epsilon^2}{\text{diam}(\mathcal{Z})^2}}.$$

■

Lemma 18 (Concentration of Wasserstein Metric on a Manifold) *Let \mathcal{Z} be a $d_{\mathcal{Z}}$ -dimensional compact class C^1 Riemannian manifold. Let μ be a Borel probability measure on \mathcal{Z} , and let μ^N denote the corresponding empirical distribution based on a sample of size N . Then exist for every $\epsilon > 0$ and $N \in \mathbb{N}$,*

$$\mathbb{P}\left(\left|\mathcal{W}_1(\mu^N, \mu) - \mathbb{E}\left[\mathcal{W}_1(\mu^N, \mu)\right]\right| \geq \epsilon\right) \leq 2e^{-\frac{2N\epsilon^2}{\text{diam}(\mathcal{Z})^2}}, \quad (\text{B.1})$$

and there exists constant $C_{\mathcal{Z}} > 0$ such that

$$\mathbb{E}\left[\mathcal{W}_1(\mu^N, \mu)\right] \leq C_{\mathcal{Z}} \cdot \text{diam}(\mathcal{Z}) N^{-1/d_{\mathcal{Z}}}. \quad (\text{B.2})$$

Proof of Lemma 18. We recall that a $d_{\mathcal{Z}}$ -dimensional class C^1 -Riemannian manifold is $d_{\mathcal{Z}}$ -dimensional topological manifold which is locally C^1 -diffeomorphic to an open subset of $\mathbb{R}^{d_{\mathcal{Z}}}$. We first show that \mathcal{Z} has Assouad dimension $d_{\mathcal{Z}}$ see (Robinson, 2011, Definitions 9.1 and 9.5). Then, we deduce the desired concentration inequality for metric spaces of Assouad dimension $d_{\mathcal{Z}}$. The for compact Riemannian \mathcal{Z} then follows.

Step 1 - Computing \mathcal{Z} 's Metric (Assouad) Dimension Since \mathcal{Z} is a $d_{\mathcal{Z}}$ -dimensional manifold then, there exists smooth charts $\{(U_k, \phi_k)\}_{k=1}^K$ where $\mathcal{Z} = \cup_{k \leq K} U_k$, $K \in \mathbb{N} \cup \{\infty\}$, and for $k = 1, \dots, K$, $\phi_k : U_k \rightarrow B_{\mathbb{R}^{d_{\mathcal{Z}}}}(0, 1)$ is a (class C^1) diffeomorphism and each U_k is an open and bounded subset of \mathcal{Z} and such that

$$\mathcal{Z} = \bigcup_{k=1}^K \phi_k^{-1} \left[B_{\mathbb{R}^{d_{\mathcal{Z}}}}(0, 1/2) \right].$$

Since \mathcal{Z} is compact and $\{U_k\}_{k \leq K}$ is an open cover thereof then, we may without loss of generality assume that K is finite.

Applying (Robinson, 2011, Lemma 9.6 (iii)) we deduce that both $B_{\mathbb{R}^{d_{\mathcal{Z}}}}(0, 1/2)$ and $B_{\mathbb{R}^{d_{\mathcal{Z}}}}(0, 1)$ have Assouad dimension $d_{\mathcal{Z}}$. By (Robinson, 2011, Lemma 9.6 (i)) we deduce that the closed Euclidean ball $\overline{B_{\mathbb{R}^{d_{\mathcal{Z}}}}(0, 1/2)} = \{u \in \mathbb{R}^{d_{\mathcal{Z}}} : \|u\| \leq 1/2\}$ must have Assouad dimension $d_{\mathcal{Z}}$. Since each ϕ_k is a diffeomorphism onto its image then $\phi_k^{-1} : B_{\mathbb{R}^{d_{\mathcal{Z}}}}(0, 1) \rightarrow U_k$ and ϕ_k are both locally Lipschitz. Thus, each ϕ_k is bi-Lipschitz when restricted to the compact set $\overline{B_{\mathbb{R}^{d_{\mathcal{Z}}}}(0, 1/2)}$. Consequentially, (Robinson, 2011, Lemma 9.6 (v)) implies that each $\phi_k^{-1}[\overline{B_{\mathbb{R}^{d_{\mathcal{Z}}}}(0, 1/2)}]$ has Assouad dimension $d_{\mathcal{Z}}$. Since K is finite and U_1, \dots, U_K all have Assouad dimension $d_{\mathcal{Z}}$ and since $\{\phi_k^{-1}[\overline{B_{\mathbb{R}^{d_{\mathcal{Z}}}}(0, 1/2)}]\}_{k=1}^K$ is a cover of \mathcal{Z} (since $\{\phi_k^{-1}[\overline{B_{\mathbb{R}^{d_{\mathcal{Z}}}}(0, 1/2)}]\}_{k=1}^K$ is a cover of \mathcal{Z}) then (Robinson, 2011, Lemma 9.6 (ii)) implies that \mathcal{Z} has Assouad dimension $d_{\mathcal{Z}}$.

Step 2 - The Concentration Inequality The assumption that \mathcal{Z} has finite Assouad dimension $d_{\mathcal{Z}}$ is equivalent to the existence of a constant $K_{\mathcal{Z}}$ satisfying: for every $r > 0$

$$\mathcal{N}_{\mathcal{Z}}^{cov}(r) \leq K_{\mathcal{Z}} \left(\frac{\text{diam}(\mathcal{Z})}{r} \right)^{d_{\mathcal{Z}}} \tag{B.3}$$

Therefore, \mathcal{Z} satisfies the Assumption made in (Boissard and Le Gouic, 2014, Equation (2)); hence, we may apply (Boissard and Le Gouic, 2014, Corollary 1.2) to conclude that:

$$\mathbb{E} [\mathcal{W}_1(\mu^N, \mu)] \leq c \cdot K_{\mathcal{Z}}^{1/d_{\mathcal{Z}}} \left(\frac{2}{d_{\mathcal{Z}} - 2} \right)^{2/d_{\mathcal{Z}}} \text{diam}(\mathcal{Z}) N^{-1/d_{\mathcal{Z}}}; \tag{B.4}$$

for some constant $0 \leq c \leq \frac{2^6}{3}$. Let $C_{\mathcal{Z}} \stackrel{\text{def.}}{=} c \cdot K_{\mathcal{Z}}$ and we prove (B.2). Next, since $\text{diam}(\mathcal{Z}) < \infty$ and μ is a Borel measure on the polish space \mathcal{Z} then, (Weed and Bach, 2019, Proposition 20) applies; hence for every $\epsilon > 0$ we have the estimate

$$\mathbb{P} \left(\left| \mathcal{W}_1(\mu^N, \mu) - \mathbb{E} [\mathcal{W}_1(\mu^N, \mu)] \right| \geq \epsilon \right) \leq 2e^{\frac{-2N\epsilon^2}{\text{diam}(\mathcal{Z})^2}}.$$

■

Remark 19 (Acceleration of Rates in Lemma 18 Under Additional Regularity)

If \mathcal{Z} is a submanifold of Euclidean space with finite-reach³ and if μ has a density with respect to the volume measure on \mathcal{Z} then a variant of Lemma 18 with a faster concentration rate can be derived using the results of Block et al. (2021) instead of Boissard and Le Gouic (2014).

B.3 Helper Sub-Gaussian Concentration Inequalities

Definition 20 (Sub-Gaussian distribution) A centered random variable X is called sub-Gaussian if there exists $C > 0$ and $\sigma > 0$ s.t. for all $x > 0$ that

$$\mathbb{P}[|X| \geq x] \leq Ce^{-\frac{x^2}{2\sigma^2}},$$

denoted by $X \sim \text{subG}(C, \sigma^2)$.

Lemma 21 Let $C, \sigma > 0$, $\tilde{\sigma} = \sigma \max\{C, 1\}$, and X be a centered random variable. Then each statement bellow implies the next:

1. $X \sim \text{subG}(C, \sigma^2)$.
2. $\mathbb{P}[|X| \geq x] \leq Ce^{-\frac{x^2}{2\sigma^2}}$ for all $x > 0$.
3. $\mathbb{E}[|X|^k] \leq (2\tilde{\sigma}^2)^{\frac{k}{2}} \left(\frac{k}{2}\right) \Gamma\left(\frac{k}{2}\right)$ for all $k \in \mathbb{N}_{\geq 2}$.
4. $\mathbb{E}[\exp(tX)] \leq e^{4\tilde{\sigma}^2 t^2}$ for all $t \in \mathbb{R}$.
5. $X \sim \text{subG}(2, 4\tilde{\sigma}^2)$.

Proof of Lemma 21. (i) \Rightarrow (ii) by definition. (ii) \Rightarrow (iii) is true by the following estimate:

$$\begin{aligned} \mathbb{E}[|X|^k] &= \int_0^\infty \mathbb{P}[|X|^k \geq t] dt = \int_0^\infty \mathbb{P}[|X| \geq t^{\frac{1}{k}}] dt \\ &\leq \int_0^\infty Ce^{-\frac{t^{2/k}}{2\sigma^2}} dt \\ &\leq \frac{Ck(2\sigma^2)^{\frac{k}{2}}}{2} \int_0^\infty e^{-u} u^{\frac{k}{2}-1} du \\ &\leq C \left(\frac{k}{2}\right) (2\sigma^2)^{\frac{k}{2}} \Gamma\left(\frac{k}{2}\right) \\ &\leq \max\{1, C^2\} (2\sigma^2)^{\frac{k}{2}} \left(\frac{k}{2}\right) \Gamma\left(\frac{k}{2}\right) \\ &\leq (2 \max\{1, C^2\} \sigma^2)^{\frac{k}{2}} \left(\frac{k}{2}\right) \Gamma\left(\frac{k}{2}\right) \\ &\leq (2\tilde{\sigma}^2)^{\frac{k}{2}} \left(\frac{k}{2}\right) \Gamma\left(\frac{k}{2}\right). \end{aligned}$$

3. The *reach* of a submanifold \mathcal{Z} of a Euclidean space is the largest radius $r \geq 0$ for which each point in the Euclidean space whose Euclidean distance from \mathcal{Z} is at-most r has a unique projection onto \mathcal{Z} ; see Genovese et al. (2012) for further details.

(iii) \Rightarrow (iv) is true because for all $t \in \mathbb{R}$

$$\begin{aligned}
 \mathbb{E}[e^{tX}] &\leq 1 + \sum_{k=2}^{\infty} \frac{t^k \mathbb{E}[|X|^k]}{k!} \\
 &\leq 1 + \sum_{k=1}^{\infty} \frac{(2\tilde{\sigma}^2 t^2)^k 2k \Gamma(k)}{(2k)!} + \sum_{k=1}^{\infty} \frac{(2\tilde{\sigma}^2 t^2)^{k+1/2} (2k+1) \Gamma(k+1/2)}{(2k+1)!} \\
 &\leq 1 + (2 + \sqrt{2\tilde{\sigma}^2 t^2}) \sum_{k=1}^{\infty} \frac{(2\tilde{\sigma}^2 t^2)^k k!}{(2k)!} \\
 &\leq 1 + \left(1 + \sqrt{\frac{\tilde{\sigma}^2 t^2}{2}}\right) \sum_{k=1}^{\infty} \frac{(2\tilde{\sigma}^2 t^2)^k}{k!} \\
 &\leq e^{2\tilde{\sigma}^2 t^2} + \sqrt{\frac{\tilde{\sigma}^2 t^2}{2}} (e^{2\tilde{\sigma}^2 t^2} - 1) \leq e^{4\tilde{\sigma}^2 t^2}.
 \end{aligned}$$

(iv) \Rightarrow (v): for all $x > 0$ and $t > 0$

$$\mathbb{P}(X > x) = \mathbb{P}(e^{tX} > e^{tx}) \leq \frac{\mathbb{E}[e^{tX}]}{e^{tx}} \leq e^{4\tilde{\sigma}^2 t^2 - tx}.$$

Therefore we have that

$$\mathbb{P}(X \geq x) \leq e^{-\frac{x^2}{8\tilde{\sigma}^2}}, \text{ and } \mathbb{P}(X \leq -x) \leq e^{-\frac{x^2}{8\tilde{\sigma}^2}}.$$

Therefore we conclude that

$$\mathbb{P}(|X| \geq x) \leq 2e^{-\frac{x^2}{8\tilde{\sigma}^2}},$$

that is $X \sim \text{subG}(2, 4\tilde{\sigma}^2)$. ■

Lemma 22 *Let X_1, \dots, X_n be independent with $X_i \sim \text{subG}(C, \sigma_i^2)$, and let $\alpha_i \geq 0$, $\tilde{\sigma}_i = \sigma_i \max\{C, 1\}$, for all $i = 1, \dots, n$. Then we have*

$$\sum_{i=1}^n \alpha_i X_i \sim \text{subG}\left(2, 4 \sum_{i=1}^n \alpha_i^2 \tilde{\sigma}_i^2\right),$$

that is, for all $x > 0$,

$$\mathbb{P}\left[\left|\sum_{i=1}^n \alpha_i X_i\right| \geq x\right] \leq 2e^{-\frac{x^2}{8\tilde{\sigma}^2}},$$

where $\tilde{\sigma}^2 = \sum_{i=1}^n \alpha_i^2 \tilde{\sigma}_i^2$.

Proof of Lemma 22. By Lemma 21, for all $i = 1, \dots, n$ and $t \in \mathbb{R}$ we have that

$$\mathbb{E}[\exp(tX_i)] \leq e^{4\tilde{\sigma}_i^2 t^2}.$$

Then, by independence, we obtain

$$\mathbb{E}[\exp(t \sum_{i=1}^n \alpha_i X_i)] = \prod_{i=1}^n \mathbb{E}[\exp(t \alpha_i X_i)] \leq \exp\left(4 \sum_{i=1}^n \alpha_i^2 \tilde{\sigma}_i^2 t^2\right).$$

Then, by Lemma 21 we conclude that

$$\sum_{i=1}^n \alpha_i X_i \sim \text{subG}(2, 4 \sum_{i=1}^n \alpha_i^2 \tilde{\sigma}_i^2).$$

■

Lemma 23 (McDiarmid's inequality) *Let X_1, \dots, X_N be independent random variables, where X_i has range \mathcal{X}_i . Let $f: \mathcal{X}_1 \times \dots \times \mathcal{X}_N \rightarrow \mathbb{R}$ be any function with the (c_1, \dots, c_N) -bounded differences property: for every $i = 1, \dots, N$ and every $(x_1, \dots, x_N), (x'_1, \dots, x'_N) \in \mathcal{X}_1 \times \dots \times \mathcal{X}_N$ that differs only in the i -th coordinate, we have*

$$|f(x_1, \dots, x_N) - f(x'_1, \dots, x'_N)| \leq c_i.$$

Then for all $t > 0$ we have that

$$\mathbb{P}\left[f(X_1, \dots, X_N) - \mathbb{E}[f(X_1, \dots, X_N)] \geq t\right] \leq e^{\frac{-2t^2}{\sum_{i=1}^N c_i^2}},$$

and

$$\mathbb{P}\left[|f(X_1, \dots, X_N) - \mathbb{E}[f(X_1, \dots, X_N)]| \geq t\right] \leq 2e^{\frac{-2t^2}{\sum_{i=1}^N c_i^2}}.$$

Proof of Lemma 23. See [McDiarmid et al. \(1989\)](#). ■

Lemma 24 (Concentration of Global Mismatch) *Let \mathcal{Z} a compact metric space and $\mu \in \mathcal{P}_1(\mathcal{Z})$. Let $g^N: \mathcal{Z} \rightarrow \mathbb{R}$ continuous and \mathbf{P} a finite partition of \mathcal{Z} . Then for all $\delta > 0$, with probability $1 - \delta$*

$$\sum_{P \in \mathbf{P}} \left(1 - \frac{\mu^N(P)}{\mu(P)}\right) \int_{z \in P} g(z) \mu(dz) \leq \frac{\|g^N\|_\infty}{N^{1/2}} \max\{\sqrt{2 \ln(2/\delta)}, \sqrt{k}\}.$$

Proof of Lemma 24. We notice that

$$\begin{aligned} \mathcal{R}^N &\stackrel{\text{def.}}{=} \sum_{P \in \mathbf{P}} \left(1 - \frac{\mu^N(P)}{\mu(P)}\right) \int_{z \in P} g^N(z) \mu(dz) \\ &\leq \|g^N\|_\infty \sum_{P \in \mathbf{P}} |\mu^N(P) - \mu(P)|. \end{aligned}$$

Let $\tilde{\mu}$ be a discrete distribution on \mathbf{P} s.t. $\tilde{\mu}(P) = \mu(P)$ and ν^N the empirical measure of ν with N samples. Then we have

$$\sum_{P \in \mathbf{P}} |\mu^N(P) - \mu(P)| = \text{TV}(\nu, \nu^N),$$

where $\text{TV}(\cdot, \cdot)$ denote the total variation distance. Therefore, by the empirical estimation under total variation distance (Canonne, 2020, Theorem 1), for all $\epsilon > 0$, $N \geq \max\{\frac{|\mathbf{P}|}{\epsilon^2}, \frac{2}{\epsilon^2} \log(2/\delta)\}$

$$\text{TV}(\nu, \nu^N) \leq \epsilon.$$

Thus, we have with probability $1 - \delta$

$$\mathcal{R}_N \leq \frac{\|g^N\|_\infty}{N^{1/2}} \max\{\sqrt{2 \ln(2/\delta)}, \sqrt{k}\}.$$

■

Appendix C. Uniform Rademacher Generalization Bound

In this section we present the Rademacher generalization bound of Equation (3) with more rigor. Consider \mathcal{F}_L the class of Lipschitz functions mapping \mathcal{X} to \mathcal{Y} , with Lipschitz constant of at most L , and let $\tilde{\mathcal{F}}_L = \{\ell \circ f : f \in \mathcal{F}_L\}$. Under assumptions 1 and 3, for any random sample \mathcal{D}^N of size N , and $0 < \delta < 1$ Bartlett and Mendelson (Theorem 8, 2002) states that with probability greater than $1 - \delta$

$$\sup_{f \in \tilde{\mathcal{F}}_L} \left\{ \mathfrak{R}(f; \mu) - \hat{\mathfrak{R}}(f) \right\} \leq 2\hat{\mathcal{R}}_N(\tilde{\mathcal{F}}_L) + \|\ell\|_\infty \sqrt{\frac{8 \log 2/\delta}{N}} \quad (\text{C.1})$$

where $\hat{\mathcal{R}}_N(\tilde{\mathcal{F}}_L)$ is the empirical Rademacher complexity of $\tilde{\mathcal{F}}_L$ which is defined via

$$\hat{\mathcal{R}}_N(\mathcal{H}) = \mathbb{E}_\epsilon \sup_{h \in \mathcal{H}} \frac{1}{N} \sum_{i=1}^N \epsilon_i h(X_i, Y_i)$$

with $\epsilon = (\epsilon_1, \dots, \epsilon_N)$ being an i.i.d. vector of Rademacher random variables. By contraction of Rademacher complexity (Theorem 12, Bartlett and Mendelson, 2002), since the Loss is L_ℓ -Lipschitz we get

$$\hat{\mathcal{R}}_N(\tilde{\mathcal{F}}_L) \leq 2L_\ell \hat{\mathcal{R}}_N(\mathcal{F}_L). \quad (\text{C.2})$$

In the next lemma, we bound the Rademacher complexity.

Lemma 25 *The Rademacher complexity of the class of L -lipschitz functions, defined on a d -dimensional domain is bounded as*

$$\hat{\mathcal{R}}_N(\mathcal{F}_L) \leq \left(\frac{8(d+1)^2 D^2 (16BL)^d}{N} \right)^{1/(d+3)} + 4\sqrt{2}D \left(\frac{1}{N} \frac{(16BL)^d}{(8(d+1)D)^{d+1}} \right)^{1/(d+3)}$$

where $D := \sup_{f \in \mathcal{F}_L} \|f\|_\infty$ and $\text{diam}(\mathcal{X}) \leq B$.

By this lemma, and due to Equations (C.1) and (C.2),

$$\begin{aligned} \mathfrak{R}(f; \mu) - \hat{\mathfrak{R}}(f) &\leq 4L_\ell \left(\frac{8(d+1)^2 D^2 (16BL)^d}{N} \right)^{\frac{1}{(d+3)}} + 16L_\ell \sqrt{2} D \left(\frac{1}{N} \frac{(16BL)^d}{(8(d+1)D)^{d+1}} \right)^{\frac{1}{(d+3)}} \\ &\quad + \|\ell\|_\infty \sqrt{\frac{8 \log 2/\delta}{N}} \end{aligned} \quad (\text{C.3})$$

Therefore, there exists C for which with probability greater than $1 - \delta$,

$$\mathfrak{R}(f; \mu) - \hat{\mathfrak{R}}(f) \leq CL_\ell \left(\frac{(dD)^2 (BL)^d}{N} \right)^{1/(d+3)} + \|\ell\|_\infty \sqrt{\frac{8 \log 2/\delta}{N}}$$

implying that the generalization error vanishes at a $\mathcal{O}(N^{-1/(d+3)})$ rate. This concludes the derivation of Equation (3).

Proof of Lemma 25. We start by bounding the Metric Entropy of the function class, and then applying a discretization bound. Without a loss of generality, we may assume that $\mathbf{0}$ is included in $\mathcal{X} \times \mathcal{Y}$. Since \mathcal{X} and \mathcal{Y} are compact, then there exists B such that $\mathcal{X} \subset [0, B]^d$. By [Gottlieb et al. \(2017, Lemma 6\)](#), the metric entropy of \mathcal{F}_L is bounded as

$$\log \mathcal{N}(\delta, \mathcal{F}_L, \|\cdot\|_\infty) \leq \left(\frac{16BL}{\delta} \right)^d \ln(8/\delta) \leq \left(\frac{16BL}{\delta} \right)^d \frac{8}{\delta}.$$

Moreover, the 1-step discretization bound ([Wainwright, 2019, Proposition 5.17](#)) implies that

$$\mathcal{R}_N(\mathcal{F}_L) \leq \frac{1}{\sqrt{N}} \inf_{\delta > 0} \left(\delta \sqrt{N} + 2 \sqrt{D^2 \log \mathcal{N}(\delta, \mathcal{F}_L, \|\cdot\|_\infty)} \right)$$

where D^2 is used to upper bound the N -norm $\sup_{f \in \mathcal{F}_L} \sum_{i=1}^N f^2(X_i)/N$. By solving for δ and plugging in the optimal value we get that there exists constants C_1 and C_2 for which

$$\mathcal{R}_N(\mathcal{F}_L) \leq \left(\frac{8(d+1)^2 D^2 (16BL)^d}{N} \right)^{1/(d+3)} + 4\sqrt{2} D \left(\frac{1}{N} \frac{(16BL)^d}{(8(d+1)D)^{d+1}} \right)^{1/(d+3)}.$$

■

Appendix D. Experiment Details

We include the details of the experiments in Section 4. For visualizing the bounds in all experiments, we have used $\text{Lip}(\ell \circ \hat{f}^N | P)$ since splitting the constant as $L_\ell \text{Lip}(\hat{f}^N | P)$ may loosen the bound, in particular for the classification experiments. All experiments are repeated for multiple random seeds, in each run the following are randomized: the learning problem (i.e. the training and test sets), the network initialization, the training algorithm.

D.1 Task Descriptions

We generate random datasets for two toy regression and classification tasks.

Regression problem. For our empirical evaluations of neural network regression, we use the simplistic problem of regressing on noisy observation of a modified logistic function. Formally, our target function $f^* : \mathcal{X} \mapsto \mathcal{Y}$ with $\mathcal{X} = [-5, 5] \subset \mathbb{R}$ and $\mathcal{Y} = [-1, 2] \subset \mathbb{R}$ is defined as

$$f^*(x) = \frac{1}{1 + \exp(5(x + 2))}. \quad (\text{D.1})$$

The inputs $x \in \mathcal{X}$ are sampled i.i.d from a uniform distribution $\mathcal{U}(-5, 5)$ and the corresponding regression labels follow as $y = f^*(x) + \epsilon$, where $\epsilon \sim \mathcal{N}(0, 0.1^2)$ is i.i.d. Gaussian observation noise. Figure 7a illustrates f^* together with 20 noisy observations. The loss function we use for regression is the Huber loss,

$$\ell(y, \hat{f}(x)) := \begin{cases} (y - \hat{f}(x))^2 & \text{if } |y - \hat{f}(x)| < 1 \\ |y - \hat{f}(x)| & \text{otherwise} \end{cases} \quad (\text{D.2})$$

which behaves like the mean squared error (MSE) for small and like the mean absolute error for large regression residuals. Training with the Huber loss is equivalent to training with the MSE plus gradient clipping and thus a common choice of practitioners to prevent large regression residuals from destabilizing the neural network training.

Classification problem. We also consider a binary classification with $\mathcal{X} = [-5, 5]^2 \subset \mathbb{R}^2$ and $\mathcal{Y} = \{-1, 1\}$. The input features $x \in \mathcal{X}$ are sampled i.i.d. from a uniform distribution over \mathcal{X} . The labels are sampled i.i.d from the Bernoulli distribution $\mathcal{B}(\sigma(f^{\text{logit}}(x_1, x_2)))$ where $\sigma(z) = 1/(1 + \exp(-z))$ is the logistic function and

$$f^{\text{logit}}(x_1, x_2) = 10\sqrt{(x_1 - 2)^2 + (x_2 - 2)^2} - \frac{1}{4}\sin(2x_1) + \frac{3}{2}\cos(x_2). \quad (\text{D.3})$$

This binary classification problem is illustrated in Figure 7b. During training, we use the negative cross-entropy error,

$$\ell(x_1, x_2, y) = (1 - y)f^{\text{logit}}(x_1, x_2) - \log(\sigma(f^{\text{logit}}(x_1, x_2))) \quad (\text{D.4})$$

which is commonly used for training neural network classifiers. We visualize the bound of Corollary 6. To calculate the bound we consider the ramp lost ℓ_γ (see Section 4) with $\gamma = 5$.

D.2 Details on the Neural Network Training

In our empirical evaluations in Section 4, we use fully-connected neural networks with leaky ReLU activation functions,

$$\rho(z) = \begin{cases} z & z \geq 0, \\ z/10 & \text{otherwise.} \end{cases}$$

We train the neural network by stochastic gradient descent with the AdamW (Loshchilov and Hutter, 2019) optimizer which combines the adaptive learning rate method Adam with weight decay. Unless stated otherwise, we set the weight decay parameter to 0 (i.e., no weight decay), use an initial learning rate of 0.05 and decay the learning rate every 1000 iterations by 0.85. By default, we train for 20000 iterations with a mini-batch size of 8 in case of regression and 16 in case of classification. In the experiments where we do not vary the neural network size, we use $l = 3$ hidden layers with $w = 64$ neurons each.

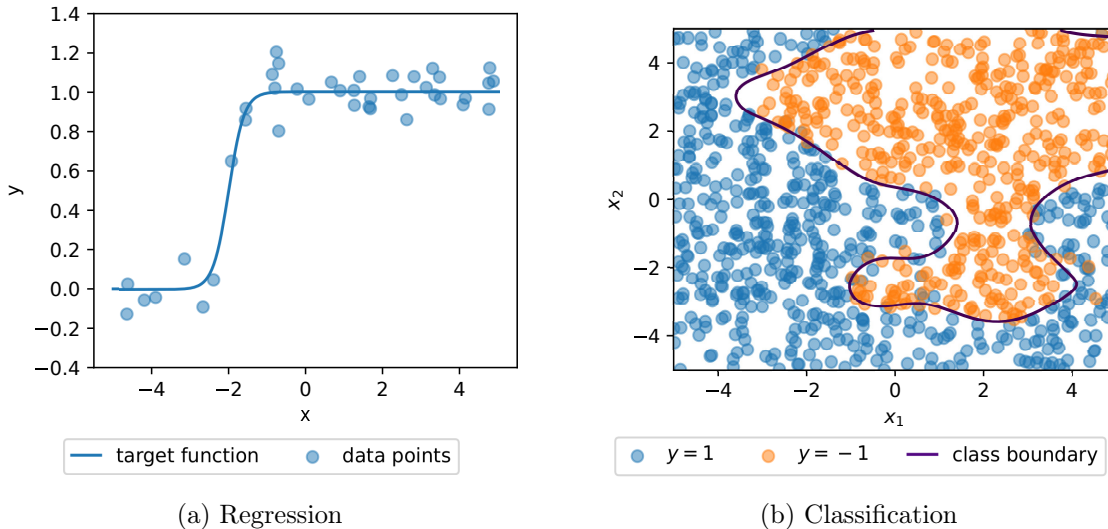


Figure 7: One instance of the random datasets generated for neural network experiments

D.3 Details of the Training Techniques Experiment (Fig. 1)

The experiment is repeated for 10 random seeds and the error bounds show the standard error. Across the regression bounds, we use a partition of size 25, i.e. a square mesh with 5 intervals along the x and y dimension. As for the classification plots, we use a partition of size 1800. We construct it as the union of two 30×30 meshes in \mathcal{X} , one located at $y = 1$ and the other $y = -1$. For a more formal definition, see Appendix A.3. We visualize the plots for different sizes of training set N , the legend in Fig. 1 shows the values.

For adversarial training, we use perturbed samples $x^{\text{adv}} = x + \epsilon \nabla_x l(x, y)$ during stochastic gradient descent. How strongly the adversary perturbs the training inputs is controlled by ϵ , i.e., the higher ϵ , the higher the regularization effect. The x -axis of Fig. 1a and Fig. 1d corresponds to this parameter. For training with weight-decay we effectively use the loss function $\ell_{\text{new}}(\mathbf{W}) = \ell_{\text{original}}(\mathbf{W}) + \lambda \|\mathbf{W}\|_F^2$ where \mathbf{W} denotes the network weights and λ is the weight-decay constant which is down on the x -axis of Fig. 1b and Fig. 1e. Lastly, Fig. 1c and Fig. 1f show the effect of early stopping, where the x -axis corresponds to the number of gradient descent iterations.

D.4 Details of the Network Size Experiment (Fig. 2)

For this experiment, we pick the depth of the network as $l \in \{1, 2, 3, 4\}$ and the width of the network as $w \in \{32, 64, 128, 256\}$. The experiment is repeated for 10 random seeds and the error bounds show the standard error. For the regression plot, we use a partition of size 50, i.e. a 10×5 mesh with 10 intervals along the \mathcal{X} and 5 along the \mathcal{Y} dimension. As for the classification plots, we use a partition of size 5000. A partition is constructed as union of two 50×50 meshes in \mathcal{X} , one located at $y = 1$ and the other $y = -1$. We visualize the plots for different sizes of training set N , the legend in Fig. 1 shows the values.

D.5 Details of the Partitioning Experiment (Fig. 5)

The experiment is repeated for 10 random seeds and the error bounds show the standard error. For the regression bounds, we consider partitions that divide the space into a uniform $M_{\mathcal{X}} \times M_{\mathcal{Y}}$ mesh. The legend of Fig. 5a shows $M_{\mathcal{Y}}$, i.e. the number of parts made in \mathcal{Y} , and the horizontal axis shows $M_{\mathcal{X}}$ the number of parts along \mathcal{X} . The regression curves are all for a dataset size of $N = 2560$.

For the classification plot, we consider partitions of size $2M^2$, where M is shown on the horizontal axis of the plot. A partition is constructed as union of two $M \times M$ meshes in \mathcal{X} , one located at $y = 1$ and the other $y = -1$. For a more formal definition, see Appendix A.3. We visualize the plots for different sizes of training set N , the legend in Fig. 5b shows the values.

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