Chapter 4 Expectation of Discrete R.V.s

Expectation

<u>Defn</u>: The **expectation of a discrete r.v.** X, written E[X], is the sum of the possible values of X, each weighted by its probability:

$$\boldsymbol{E}[X] = \sum_{x} x \cdot \boldsymbol{P}\{X = x\}$$

 $\boldsymbol{E}[X]$ also represents the **mean of the distribution** from which X is drawn.

Average Cost of Lunch

| MON | TUES | WED | THUR | FRI | SAT | SUN |
|-----|------|------|------|------|-----|-----|
| \$7 | \$7 | \$12 | \$12 | \$12 | \$0 | \$9 |

Average Cost =
$$\frac{7 + 7 + 12 + 12 + 12 + 0 + 9}{7}$$

$$E[Cost] = 7 \cdot \frac{2}{7} + 12 \cdot \frac{3}{7} + 9 \cdot \frac{1}{7} + 0 \cdot \frac{1}{7}$$

Expectation of Bernoulli(p)

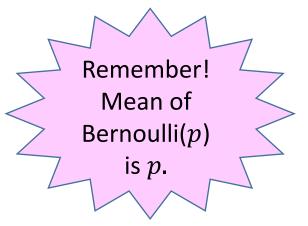
X =value of the coin flip



Probability *p* of heads

Q: What is E[X]?

$$\mathbf{E}[X] = 1 \cdot p + 0 \cdot (1 - p) = p$$



Expected Time Until Disk Fails

Disk has probability $\frac{1}{3}$ of failing each year.



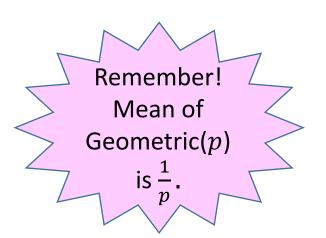
Q: On average, how many years will it be until the disk fails?

$$X \sim Geometric(p) \text{ where } p = \frac{1}{3}$$

$$E[X] = \sum_{n=1}^{\infty} n (1-p)^{n-1} p$$

$$= p \sum_{n=1}^{\infty} n q^{n-1}$$

$$= p \cdot (1+2q+3q^2+4q^3+\cdots) = p \cdot \frac{1}{(1-q)^2} = \frac{1}{p}$$



Expectation of Poisson(λ)

 $X \sim Poisson(\lambda)$

Q: What is E[X]?

$$p_X(i) = \frac{e^{-\lambda}\lambda^i}{i!}, i = 0, 1, 2, ...$$

$$E[X] = \sum_{i=0}^{\infty} i \cdot \frac{e^{-\lambda} \lambda^{i}}{i!} = \sum_{i=1}^{\infty} i \cdot \frac{e^{-\lambda} \lambda^{i}}{i!}$$

$$= \lambda e^{-\lambda} \sum_{i=1}^{\infty} \frac{\lambda^{i-1}}{(i-1)!}$$
Remember me from Chpt 1?
$$= \lambda e^{-\lambda} \left(\sum_{k=0}^{\infty} \frac{\lambda^{k}}{k!} \right)^{*} = \lambda e^{-\lambda} \cdot e^{\lambda} = \lambda$$



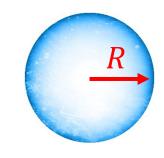
Expectation of a Function of a R.V.

<u>Defn</u>: A **expectation of a function** $g(\cdot)$ **of a discrete r.v.** X is defined as follows:

$$\mathbf{E}[g(X)] = \sum_{x} g(x) \cdot p_{X}(x)$$

Consider a sphere, whose radius is a random variable R:

$$R = \begin{cases} 1 & \text{w.p. } \frac{1}{3} \\ 2 & \text{w.p. } \frac{1}{3} \\ 3 & \text{w.p. } \frac{1}{3} \end{cases}$$

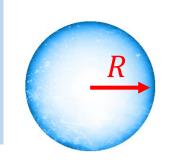


Q: What is the expected volume of the sphere?

Expectation of a Function of a R.V.

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$$\boldsymbol{E}[g(X)] = \sum_{x} g(x) \cdot p_{X}(x)$$



$$E[Volume] = E\left[\frac{4}{3}\pi R^{3}\right]$$

$$= \frac{4}{3}\pi \cdot 1^{3} \cdot \frac{1}{3} + \frac{4}{3}\pi \cdot 2^{3} \cdot \frac{1}{3} + \frac{4}{3}\pi \cdot 3^{3} \cdot \frac{1}{3}$$

$$= 16\pi$$

$$R = \begin{cases} 1 & \text{w.p. } \frac{1}{3} \\ 2 & \text{w.p. } \frac{1}{3} \\ 3 & \text{w.p. } \frac{1}{3} \end{cases}$$

Q: Is
$$E[R^3] = (E[R]^3)$$
?

Expectation of a Product

<u>Defn</u>: Let *X* and *Y* be r.v.s. The **expectation of the product** *XY* is defined as follows:

$$E[XY] = \sum_{x} \sum_{y} xy \cdot p_{X,Y}(x,y)$$

$$\mathbf{E}[g(X)f(Y)] = \sum_{x} \sum_{y} g(x)f(y) \cdot p_{X,Y}(x,y)$$

where $p_{X,Y}(x,y) = P\{X = x \& Y = y\}.$

Expectation of Product under Independence

Theorem 4.8: (Expectation of a product) If $X \perp Y$, then $E[XY] = E[X] \cdot E[Y]$.

Proof:

$$E[XY] = \sum_{x} \sum_{y} xy \cdot P\{X = x, Y = y\}$$

$$= \sum_{x} \sum_{y} xy \cdot P\{X = x\} \cdot P\{Y = y\}$$

$$= \sum_{x} x \cdot P\{X = x\} \sum_{y} y \cdot P\{Y = y\}$$

$$= E[X] \cdot E[Y]$$

Via the same proof: If $X \perp Y$, then $E[g(X)f(Y)] = E[g(X)] \cdot E[f(Y)]$.

Alternative Definition of Expectation

Theorem 4.9: (Alternative Definition of Expectation) Let X be a non-negative,

discrete, integer-valued random variable. Then

$$E[X] = \sum_{x=0}^{\infty} P\{X > x\}.$$
 This alternative formulation can be very useful.

Proof: See exercise in textbook. Hint: Rewrite the inside probability as a sum.

Expectation Property

Theorem: For random variable X and real number a,

$$\mathbf{E}[aX] = a\mathbf{E}[X]$$

$$\mathbf{E}[aX] = \sum_{x} a x p_X(x) = a \sum_{x} x p_X(x) = a\mathbf{E}[X]$$

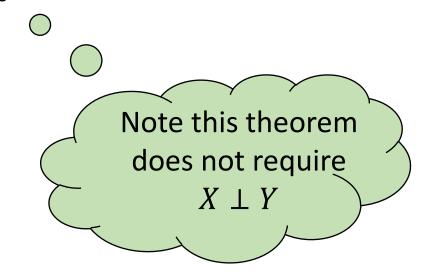
Linearity of Expectation

The following theorem greatly simplifies the computation of an expectation by breaking up the random variable into smaller pieces.

Theorem 4.10: [Linearity of Expectation] For random variables X and Y,

$$E[X+Y] = E[X] + E[Y]$$

Proof: First try proving this yourself. It's similar to the $\boldsymbol{E}[XY]$ derivation, but you aren't allowed to split the $p_{X,Y}(x,y)$...



Linearity of Expectation

Theorem 4.10: [Linearity of Expectation] For random variables X and Y,

$$\mathbf{E}[X+Y] = \mathbf{E}[X] + \mathbf{E}[Y]$$

Proof:

$$E[X + Y] = \sum_{y} \sum_{x} (x + y) \cdot p_X(x, y)$$

$$= \sum_{y} \sum_{x} x \cdot p_X(x, y) + \sum_{y} \sum_{x} y \cdot p_X(x, y)$$

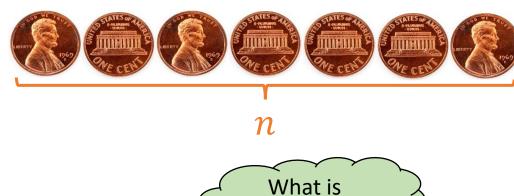
$$= \sum_{x} \sum_{y} x \cdot p_X(x, y) + \sum_{y} \sum_{x} y \cdot p_X(x, y)$$

$$= \sum_{x} \sum_{y} p_X(x, y) + \sum_{y} \sum_{x} p_X(x, y) = \sum_{x} x p_X(x) + \sum_{y} y p_Y(y) = E[X] + E[Y]$$

Expectation of Binomial(n, p)

Experiment: Flip a coin, with probability p of Heads, n times

Random Variable X = number of heads



Key Observation:

$$X = X_1 + X_2 + \cdots + X_n$$
, where $X_i \sim \text{Bernoulli}(p)$

Applying Linearity of Expectation:

$$\begin{aligned} \textbf{\textit{E}}[\textbf{\textit{X}}] &= \textbf{\textit{E}}[\textbf{\textit{X}}_1] + \textbf{\textit{E}}[\textbf{\textit{X}}_2] + \cdots + \textbf{\textit{E}}[\textbf{\textit{X}}_n] \\ &= p + p + \cdots + p = np \circ \circ \circ \end{aligned}$$
 Should make intuitive sense

Remember!

Mean of
Binomial(n, p)is np.

Expectation of Binomial(n, p)

Experiment: Flip a coin, with probability p of Heads, n times

Random Variable X = number of heads



Key Observation:

$$X = X_1 + X_2 + \cdots + X_n$$
, where $X_i \sim \text{Bernoulli}(p)$

Applying Linearity of Expectation:

<u>Defn</u>: The X_i here are called **indicator r.v.s**, because they take on values of 1 or 0.

At a party, n people put their drink on a table. Later that night, no one can remember which cup is theirs, so they simply each grab any cup at random.

Let X = number of people who get back their own cup.

Q: What is E[X]? Is it increasing with n?



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Idea:
$$X = X_1 + X_2 + \cdots + X_n$$

Q: What do the X_i represent?



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Q: What is E[X]?

Idea:
$$X = X_1 + X_2 + \dots + X_n$$

$$X_i = \begin{cases} 1 & \text{w.p. } 1/n \\ 0 & \text{o.w.} \end{cases}$$



 $X_i = 1 \Leftrightarrow \text{person } i \text{ got back their own cup}$

Q: Are the X_i independent Bernoulli distributions? If so, is X Binomially distributed?

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 $X_i = 1 \Leftrightarrow \text{person } i \text{ got back their own cup}$

A: The X_i 's are NOT independent. Nevertheless, Linearity of Expectation applies:

$$\boldsymbol{E}[X] = \boldsymbol{E}[X_1] + \boldsymbol{E}[X_2] + \dots + \boldsymbol{E}[X_n] = n\boldsymbol{E}[X_i] = n \cdot \frac{1}{n} = 1.^{\circ} \circ \circ \underbrace{\text{regardless}}_{\text{of } n}$$

There are n coupons we're trying to collect. Each draw we get a random coupon (sampling with replacement).











Let D = number of draws until we've collected all the coupons.

Q: What is E[D]?

There are n coupons we're trying to collect. Each draw we get a random coupon (sampling with replacement).











Let D = number of draws until we've collected all the coupons.

Q: What is E[D]?

Idea:
$$D = D_1 + D_2 + D_3 + \cdots + D_n$$

Q: But what do the D_i represent? \circ

What's wrong with letting D_i represent number of draws to get coupon i?

There are n coupons we're trying to collect. Each draw we get a random coupon (sampling with replacement).











Let D = number of draws until we've collected all the coupons.

Q: What is E[D]?

Idea: $D = D_1 + D_2 + D_3 + \cdots + D_n$

draws to get 1st distinct coupon

Additional draws needed to reach 2 distinct coupons

Additional draws needed to reach 3 distinct coupons

 $D_i = \text{number of draws needed to get } i \text{th distinct coupon,}$ given already have i-1 distinct coupons

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Additional draws needed to reach 3 distinct coupons

 $D_i = \text{number of draws needed to get } i \text{th distinct coupon,}$ given already have i-1 distinct coupons

$$D_1 \sim Geometric(1)$$

$$D_2 \sim Geometric\left(\frac{n-1}{n}\right)$$

$$D_3 \sim Geometric\left(\frac{n-2}{n}\right)$$

$$D_n \sim Geometric\left(\frac{1}{n}\right)$$

There are n coupons we're trying to collect. Each draw we get a random coupon (sampling with replacement).











Let D = number of draws until we've collected all the coupons.

Q: What is E[D]?

Idea:
$$D = D_1 + D_2 + D_3 + \dots + D_n$$

$$E[D] = E[D_1] + E[D_2] + E[D_3] + \dots + E[D_n]$$

$$= 1 + \frac{n}{n-1} + \frac{n}{n-2} + \dots + n$$

$$= n \cdot \left(\frac{1}{n} + \frac{1}{n-1} + \frac{1}{n-2} + \dots + 1\right)$$

$$D_1 \sim Geometric(1)$$

$$D_2 \sim Geometric\left(\frac{n-1}{n}\right)$$

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$$D_n \sim Geometric\left(\frac{1}{n}\right)$$
 25

There are n coupons we're trying to collect. Each draw we get a random coupon (sampling with replacement).











Let D = number of draws until we've collected all the coupons.

Q: What is E[D]?

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$$D = D_1 + D_2 + D_3 + \cdots + D_n$$

$$E[D] = E[D_1] + E[D_2] + E[D_3] + \cdots + E[D_n]$$

$$= 1 + \frac{n}{n-1} + \frac{n}{n-2} + \dots + n$$

$$= n \cdot \left(\frac{1}{n} + \frac{1}{n-1} + \frac{1}{n-2} + \dots + 1\right) = n \cdot H_n \approx n \ln(n)$$

Conditional p.m.f.

We often want the expected value of a r.v. X conditioned on some event, A, e.g.

E[Price of hotel room | Month is March]

To define E[X|A] we will need to define a conditional p.m.f., $p_{X|A}(x)$.

<u>Defn 4.14</u>: Let X be a discrete r.v. with p.m.f. $p_X(x)$.

Let A be an event s.t. $P{A} > 0$.

Then $p_{X|A}(x)$ is the **conditional p.m.f. of** X **given event** A **where:**

$$p_{X|A}(x) = P\{X = x \mid A\} = \frac{P\{(X = x) \cap A\}}{P\{A\}}$$

Conditioning on an Event

Let r.v. *X* denote the size of a job:

$$X = \begin{cases} 1 & \text{w.p. } 0.1\\ 2 & \text{w.p. } 0.2\\ 3 & \text{w.p. } 0.3\\ 4 & \text{w.p. } 0.2\\ 5 & \text{w.p. } 0.2 \end{cases}$$

Let A denote the event that the job is "small," meaning its size is ≤ 3 .

Q: What is $p_{X|A}(1)$? How does this compare with $p_X(1)$?

Conditioning on an Event

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Let A denote the event that the job is "small," meaning its size is ≤ 3 .

Q: What is $p_{X|A}(1)$? How does this compare with $p_X(1)$?

A:
$$p_{X|A}(1) = P\{X = 1 \mid A\} = \frac{P\{X = 1 \& A\}}{P\{A\}} = \frac{P\{X = 1\}}{P\{A\}} = \frac{\frac{1}{10}}{\frac{6}{10}} = \frac{1}{6}$$

Conditioning on an Event

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Let A denote the event that the job is "small," meaning its size is ≤ 3 .

Q: What is $p_{X|A}(x)$ if $x \notin A$? **Answer:** 0

Lemma 4.16: A conditional p.m.f. is a p.m.f., i.e.,

$$\sum_{x} p_{X|A}(x) = \sum_{x \in A} p_{X|A}(x) = 1$$

Example: Conditioning on an Event

Table shows $p_{X,Y}(x,y)$

| | X = 0 | X = 1 | X = 2 |
|-------|-------|-------|-------|
| Y = 0 | 1/6 | 1/8 | 0 |
| Y = 1 | 1/8 | 1/6 | 1/8 |
| Y = 2 | 0 | 1/6 | 1/8 |

Q: What is $p_{X|Y=2}(1)$?

A:
$$p_{X|Y=2}(1) = P\{X = 1 \mid Y = 2\} = \frac{P\{X = 1 \& Y = 2\}}{P\{Y = 2\}} = \frac{\frac{1}{6}}{\frac{1}{6} + \frac{1}{8}} = \frac{4}{7}$$

Conditional Expectation

The *conditional* expectation, E[X|A], is based on the *conditional* p.m.f., $p_{X|A}(x)$.

Defn: Let *X* be a discrete r.v.

The **conditional expectation of** *X* **given event** *A* is defined as:

$$E[X|A] = \sum_{x} x \cdot p_{X|A}(x) = \sum_{x} x \cdot \frac{P\{(X=x) \cap A\}}{P\{A\}}$$

Let r.v. *X* denote the size of a job:

$$X = \begin{cases} 1 & \text{w.p. } 0.1 \\ 2 & \text{w.p. } 0.2 \\ 3 & \text{w.p. } 0.3 \\ 4 & \text{w.p. } 0.2 \\ 5 & \text{w.p. } 0.2 \end{cases}$$

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$$X = \begin{cases} 1 & \text{w.p. } 0.1\\ 2 & \text{w.p. } 0.2\\ 3 & \text{w.p. } 0.3\\ 4 & \text{w.p. } 0.2\\ 5 & \text{w.p. } 0.2 \end{cases}$$

Let A denote the event that the job is "small," meaning its size is ≤ 3 .

Q: What is E[X|A]?

A:
$$E[X|A] = 1 \cdot p_{X|A}(1) + 2 \cdot p_{X|A}(2) + 3 \cdot p_{X|A}(3)$$

= $1 \cdot \frac{1}{6} + 2 \cdot \frac{2}{6} + 3 \cdot \frac{3}{6} = \frac{14}{6}$

Table shows $p_{X,Y}(x,y)$

| | X = 0 | X = 1 | X = 2 |
|-------|-------|-------|-------|
| Y = 0 | 1/6 | 1/8 | 0 |
| Y = 1 | 1/8 | 1/6 | 1/8 |
| Y = 2 | 0 | 1/6 | 1/8 |

Q: What E[X | Y = 2]?

Table shows $p_{X,Y}(x,y)$

| | X = 0 | X = 1 | X = 2 |
|-------|-------|-------|-------|
| Y = 0 | 1/6 | 1/8 | 0 |
| Y = 1 | 1/8 | 1/6 | 1/8 |
| Y = 2 | 0 | 1/6 | 1/8 |

Q: What E[X | Y = 2]?

A:
$$E[X \mid Y = 2] = 0 \cdot p_{X|Y=2}(0) + 1 \cdot p_{X|Y=2}(1) + 2 \cdot p_{X|Y=2}(2)$$

= $1 \cdot \frac{4}{7} + 2 \cdot \frac{3}{7} = \frac{10}{7}$

Computing Expectations via Conditioning

Theorem 4.22: Let X be a discrete r.v.

Let events F_1, F_2, \dots, F_n partition the space Ω . Then

$$\boldsymbol{E}[X] = \sum_{i=1}^{n} \boldsymbol{E}[X | F_i] \cdot \boldsymbol{P}\{F_i\}$$

For a discrete r.v. Y:

$$E[X] = \sum_{y} E[X | Y = y] \cdot P\{Y = y\}$$

Expected Value of Geometric, Revisited

 $X \sim Geometric(p)$. Derive E[X] by conditioning.



Q: What should we condition on?

Expected Value of Geometric, Revisited

 $X \sim Geometric(p)$. Derive E[X] by conditioning.



Q: What should we condition on?

A: Condition on the value of the first flip,

$$E[X] = E[X | Y = 1] \cdot P\{Y = 1\} + E[X]$$
 What is this? $P\{Y = 0\}$

$$= E[X | Y = 1] \cdot p + E[X | Y = 0] \quad (1 - p)$$

Expected Value of Geometric, Revisited

 $X \sim Geometric(p)$. Derive E[X] by conditioning.



Q: What should we condition on?

A: Condition on the value of the first flip, *Y*.

$$E[X] = E[X | Y = 1] \cdot P\{Y = 1\} + E[X | Y = 0] \cdot P\{Y = 0\}$$

$$= E[X | Y = 1] \cdot p + E[X | Y = 0] \cdot (1 - p)$$

$$= 1 \cdot p + (1 + E[X]) \cdot (1 - p)$$

$$\Rightarrow E[X] = \frac{1}{p}$$

Simpson's Paradox

Consider two treatments for kidney stones: Treatment A and Treatment B

- Treatment A is more effective on small kidney stones
- Treatment A is also more effective on large kidney stones

But if we ignore the type of stones, **Treatment B** is more effective!

Simpson's Paradox

Q: How is this possible?

| | Treatment A | Treatment B |
|---------------|---------------|---------------|
| small stones | 90% effective | 80% effective |
| large stones | 60% effective | 50% effective |
| aggregate mix | 63% effective | 77% effective |

Simpson's Paradox

| | Treatment A | Treatment B |
|---------------|--|--|
| small stones | 90% effective (successful on 90 out of 100) | 80% effective (successful on 800 out of 1000) |
| large stones | 60% effective (successful on 600 out of 1000) | 50% effective (successful on 50 out of 100) |
| aggregate mix | 63% effective (successful on 690 out of 1100) | 77% effective (successful on 850 out of 1100) |

| | Treatment A | Treatment B |
|---------------|--|--|
| small stones | 90% effective (successful on 90 out of 100) | 80% effective (successful on 800 out of 1000) |
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Is treatment B better?

- □ No! Treatment A is better on both small stones and on large stones. It is the better treatment!
- ☐ But because A is better, it is given more "hard cases" the large stone cases and hence has lower average scores.